- **Phenological sensitivity to climate across taxa and trophic levels**
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Among-species differences in phenological responses to climate change can desynchronise ecological interactions, threatening ecosystem function. To assess these threats, we must quantify the relative impact of climate change on species at different trophic levels. Here, we apply a novel Climate Sensitivity Profile approach to 10,003 terrestrial and aquatic phenological data sets, spatially-matched to temperature and precipitation data, quantifying variation in climate sensitivity. The direction, magnitude and timing of climate sensitivity varied markedly among organisms within taxonomic and trophic groups. Despite this, we detected systematic variation in the direction and magnitude of phenological climate sensitivity. Secondary consumers showed consistently lower climate sensitivity than other groups. Based upon mid-century climate change projections, we estimate that the timing of phenological events could change more for primary consumers than for other trophic levels (6.2 versus 2.5 - 2.9 days earlier on average), with substantial taxonomic variation (1.1 - 14.8 days earlier on average).

18 Numerous long-term ecological changes have been attributed to climate change¹. Shifts in the seasonal timing of recurring biological events such as reproduction and migration (i.e. 20 phenological changes) are especially well documented^{2,3}. Long-term ecosystem studies^{4–7} and 21 global meta-analyses^{2,3,8} have demonstrated that many spring and summer phenological events now occur earlier in the year. Substantial among-species variation in responses has fuelled concerns that key seasonal species interactions may desynchronise over time, with 24 potentially severe consequences for wild populations and, hence, for ecosystem functioning⁹.

Identifying systematic taxonomic and trait-based differences in phenological climate sensitivity (i.e. change in seasonal timing per unit change in climatic conditions) would have significant socio-ecological implications. This would afford some predictability to future ecological outcomes and would identify species that represent effective sentinels of climate impact, facilitating the development of indicators and estimates of vulnerability for 31 conservation and national adaptation programmes ^{$10-12$}. Unfortunately, such generalisations are currently elusive.

Analytical approach and data sets

Among-species differences in phenological change may arise from two aspects of climate sensitivity. Firstly, variation may reflect differences in physiological and behavioural responses, microclimate use, and the importance of non-climate related cues, such as 38 photoperiod¹³ or resource availability¹⁴. Therefore, even when species have the same seasonal period (window) for which they are most sensitive to climate change, they show different phenological responses to a given climatic change. Secondly, co-occurring species may vary in their seasonal periods of climate sensitivity, each typified by different levels of directional 42 climate change^{15–17}. We conceptualise these two aspects of phenological responses as species- (or population-) specific Climate Sensitivity Profiles (CSPs, Fig. 1). The CSP approach differs fundamentally from attempts to identify single "critical" seasonal periods 45 within which climatic change most strongly affects seasonal events¹⁷, by quantifying the full range of phenological responses to seasonal climatic change. We ask "*How sensitive are phenological events to temperature and precipitation change at different times of year?*". By applying this approach to a large, taxonomically-diverse national-scale data set, we discern

coherent patterns within a multitude of idiosyncratic biological climate-responses. We assess whether systematic differences in climate sensitivity underpin differences in phenological 51 change among taxonomic and trophic groups in the $UK⁸$.

We elected against using published climate responses that may be biased in favour of species 54 showing an effect. Instead, we analysed 10,003 long-term (\geq 20 year) phenological time series for 812 marine, freshwater and terrestrial taxa over the period 1960-2012. Our data set aggregates many of the UK's foremost long-term biological monitoring schemes (Supplementary Table 1), including phenological information on amphibians (spawning), birds (egg laying, migration), planktonic crustaceans (population peaks), fish (spawning, migration), insects (flight periods), mammals (birth dates), phytoplankton (population peaks) and plants (flowering, fruiting, leafing). These taxa represent three broad trophic levels: primary producers (phytoplankton, plants), primary consumers (granivorous birds, herbivorous insects, mammals, planktonic crustaceans) and secondary consumers (predatory amphibians, birds, fish, insects, mammals, planktonic crustaceans). We spatially-matched all 64 10,003 phenological time series with local temperature and precipitation data from a $5\times5km$ resolution gridded data set, before statistically modelling the relationship between seasonal timing and climatic variables. Between 1960 and 2012 mean UK air temperatures increased in all months, and mean precipitation increased in most months (Fig. 2a).

Spatial variability in climatic change (Fig. 2b,c), necessitates local matching of phenological and climatic datasets rather than the use of regionally-averaged climate data (e.g. Central England Temperatures) or large-scale climatic indicators (e.g. North Atlantic Oscillation). We did not make the restrictive assumption that biological events would be related to annual

mean climatic conditions, or to conditions within periods based upon calendar months. Our CSP approach identified seasonal periods within which climatic change has its most positive and negative correlations with phenology (hereafter referred to as upper and lower limits of climate sensitivity, respectively). We could identify, for each phenological series, up to two seasonal periods within which climatic variation had a significant correlation with seasonal timing. The method was flexible enough to 1) allow situations in which climatic variation within only a single period had a significant correlation, and 2) identify seasonal windows ranging from a few days to a whole year in length. Our analysis captured the idiosyncrasies of phenological responses, allowed their categorisation into generic types of climate response, and is consistent with current biological understanding of climate-phenology 83 relationships $15,16$.

Climate response-types in the UK

CSPs fall into three categories. The qualitative type of climate-phenology correlation (positive or negative) may remain consistent, irrespective of when in the year climatic change occurs. In this case only the magnitude of the phenological response differs with the time of year at which climatic variables change. The climate-phenology correlation may be consistently negative (CSP type I, Fig. 1, red curve) or positive (CSP type III, Fig. 1, blue curve). Alternatively, opposing correlations between seasonal climatic change and the timing of biological events may exist i.e. the direction and magnitude of the phenological response varies (CSP type II, Fig. 1, orange curve). We determined CSPs for responses to temperature 94 (CSP_{temp}) and precipitation (CSP_{precup}).

Focusing on temperature, CSP type II was most common (Extended Data Table 1, 69.7 % of phenological series): seasonal events were advanced by (i.e. negatively correlated with) warming during one period of the year, and delayed by (i.e. positively correlated with) warming in another period. After multiple testing correction, 44.8% of the observed phenological advances (but only 1.0% of delays) with warming were statistically significant (P<0.05). CSP type I was the next most common response-type: warming in different seasonal windows was consistently correlated with earlier seasonal events (i.e. negative correlations, 24.7% of series). In this case the lower and upper limits of CSPs represent the "strongest" and "weakest" phenological advances with warming, respectively, and 58.1% of the "strongest" responses were statistically significant (P<0.05, correcting for multiple testing).

Phenological events most commonly demonstrated opposing (Fig. 1, CSP type II, 53.0% of series) or consistently positive (Fig. 1, CSP type III, 28.0% of phenological series) correlations with increasing seasonal precipitation. Though delayed phenological events may commonly be associated with higher precipitation (81.0% of events show this type of response), few of these associations were significant (Extended Data Table 1).

Climate sensitivity at the UK-scale

We matched each phenological series with four climate variables: mean temperature during 116 the seasonal windows at the upper and lower limits of CSP_{temp} , and similarly-averaged 117 precipitation data for the seasonal windows at the upper and lower limits of CSP_{precip} . We then combined all 10,003 phenological series and their matched climate data, and modelled

the relationships between seasonal timing and climate variables using linear mixed effects (LME) models. Initially we fitted a "global" model to quantify upper and lower limits of temperature and precipitation sensitivity, averaged across all phenological events. Marine plankton data were excluded at this stage, due to a lack of precipitation data.

124 Most phenological events occurred earlier with seasonal warming (average rate -2.6 days °C⁻ $\frac{1}{1}$, Fig. 3a, Extended Data Table 2). Variation in the strength of this correlation was similar among sites and species (random-effects variances in site and species level seasonal timing – temperature slopes were 2.1 and 1.9, respectively). Some phenological events occurred later 128 with seasonal warming (Fig. 3a) though, in other cases, the upper limit of CSP_{temp} was in fact a "weak" advance with warming. The upper limit of temperature sensitivity was more variable among species than sites (random effects variances in species and site level seasonal timing – temperature slopes were 2.3 and 0.4, respectively). Averaged across species and populations, temperature responses were most consistent with CSP type II.

Most phenological events showed opposing responses to increasing seasonal precipitation (Fig. 1, CSP type II). The tendency for delays with rising precipitation was greatest: the 136 average upper limit of CSP_{precip} exceeded the lower limit (1.4 days mm⁻¹ and -0.4 days mm⁻¹, 137 respectively, Fig. 3b, Extended Data Table 2). The upper limit of CSP_{precin} was more variable among species than sites (species and site level random-effects variances in the seasonal timing – precipitation slopes were 1.9 and 1.2, respectively). The fitted climate-phenology model was better supported by the data than a year-only model with the same random effects structure (delta-AIC 293,516). This indicates the presence of real associations between climate and seasonality, rather than purely spurious correlations due to shared temporal

trends. Average sensitivity to temperature was very similar in the model that included marine plankton data, but excluded precipitation effects (Supplementary Discussion, Extended Data Fig. 1).

Taxonomic and trophic group sensitivity

We tested the hypothesis that the limits of seasonal climate sensitivity differ coherently among taxonomic groups by including a fixed-effect interaction between taxonomic group 150 and each climatic variable (Fig. 4, Extended Data Table 2). The lower limit of CSP_{temp} was negative for all groups ("earliness" with warming), the strongest responses being found for plants, freshwater phytoplankton, insects and amphibians (4.3, 4.1, 3.7 and 3.4 days earlier \degree C⁻¹, respectively). Upper limits of CSP_{temp} indicated that freshwater phytoplankton and mammals experienced the greatest phenological delays with seasonal warming (2.9 and 2.0 155 days later ${}^{\circ}C^{-1}$, respectively) but that plants showed little evidence of such delays. The strongest phenological delays with rising seasonal precipitation were found for freshwater 157 phytoplankton and insects $(2.5 \text{ and } 2.2 \text{ days later mm}^{-1})$, respectively), while freshwater phytoplankton also exhibited the strongest phenological advances with rising precipitation 159 during other seasonal windows $(1.1 \text{ days earlier mm}^{-1})$. Average temperature and precipitation responses were consistent with a CSP type II in most cases. There was considerable within-group variability in sensitivity.

We examined trophic-level differences in climate sensitivity by including this in interaction 164 with each climate variable in the global model. The lower limit of CSP_{temp} showed greater systematic variation among trophic levels than the upper limit (Fig. 3c,e). The tendency

Estimating future change

Overall, "net", phenological responses to climatic change combine potentially-opposing responses to conditions in different seasonal periods. We estimated "net" responses by the

2050s by applying our fitted models to UKCP09 probabilistic projections (bias-corrected relative to a 1961-90 baseline) of temperature and precipitation change under low, medium and high emissions scenarios. Rather than predicting the *absolute* timing of future phenological events, we contrasted possible *changes* in seasonal timing among organism groups based upon established climate scenarios and contemporary patterns of climate sensitivity. Estimated average phenological changes for primary producers and secondary consumers were less than those for primary consumers (Fig. 5a). This occurred because, averaged across species, the opposing climate responses of primary producers and secondary consumers are more similar in magnitude than are those for primary consumers (Fig. 3), effectively "cancelling each other out". Our models suggest greater average advances for crustacea, fish and insects than for other groups, such as freshwater phytoplankton, birds and mammals (Fig. 5b). However, response-variation is high for crustacea (Fig. 5b).

Discussion

In the UK, phenological climate sensitivity varies greatly, suggesting effects of locally-204 varying non-climatic drivers such as population structure¹⁸, resource availability¹⁹ and 205 adaptation²⁰. This is relevant to the use of phenological change as a tangible climate change 206 indicator^{1,21}. Mediators of phenological climate sensitivity are only known locally for some 207 of the groups in our data set e.g. nutrient availability (freshwater phytoplankton)²². However, for others, the climate sensitivity of different biological traits is known to be mediated by 209 alternative drivers^{23,24}. High climate-response variability necessitates wide site and species coverage in long-term monitoring schemes aiming to develop robust aggregate indicators of 211 change²¹. Since climatic conditions are more spatially-variable across broader geographic domains, site-level replication of phenological monitoring is particularly important when

interpreting phenology at continental to global scales. In the UK, average responses for fish and insects appear to provide climate-indicator potential. These groups show consistently strong phenological advances with seasonal warming, and only weak opposing responses, resulting in relatively large (net) changes in seasonal timing. Interpretation of phenological changes for other groups is more complex. For example, freshwater phytoplankton show strong evidence of opposing phenological responses to climatic variation at different times of year and these are near-equivalent in magnitude, such that estimated net changes are negligible. This highlights that long-term observations represent the net effect of potentially-221 opposing biological responses²⁵. To fully capitalise on the indicator potential of phenological change, we must advance mechanistic understanding of responses to potentially opposing climate and non-climate drivers.

Despite this variability, we identified coherent patterns in climate sensitivity among the idiosyncratic responses of many wild plant and animal populations. For the first time we show that, on average, trophic levels differ in the magnitude of seasonal climate sensitivity, but not the time-of-year within which climatic change has its most pronounced effects. This may be a key mechanism underpinning observations of trophic level differences in 230 . phenological change in the $UK⁸$. Lower trophic levels demonstrated more pronounced variation in their sensitivity to changing temperature and precipitation at different times of year, and stronger phenological responses to climatic change during defined (taxon- and population-specific) seasonal periods.

In response to climatic changes projected for the 2050s, relative changes in seasonal timing are likely to be greatest for primary consumers, particularly in the terrestrial environment.

The difference in magnitude between opposing climate responses is greatest for primary consumers, resulting in greater "net" change. Our approach makes the simplifying assumption that climatic change has the overriding influence upon seasonality. Nevertheless, this suggests that systematic differences in climate sensitivity could result in widespread phenological desynchronisation. However, factors that shape phenological climate-responses introduce uncertainty into projections of future phenological change. These results should catalyse research to improve predictive capacity in the face of multiple environmental and demographic drivers that not only mediate rates of change, but might also confer resilience to desynchronisation e.g. population density-dependence²⁶, compensatory range shifts²⁷, and the 246 formation of novel inter-specific interactions^{28,29}. These findings also underscore the importance of developing our capacity to manage ecosystems within a "safe operating space" 248 with respect to the likely impacts of projected climate change³⁰.

Supplementary Information is linked to the online version of the paper at

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investment in scientific endeavour. Whilst we are committed to sharing these data for

scientific research, users are requested to collaborate before publication of these data to

ensure accurate biological interpretation. We thank four referees for their comments.

Author Contributions

SJT and SW conceived and co-ordinated the study, and led writing of the manuscript. PAH developed the analysis routine and wrote statistical code to be applied to all data sets. DH extracted all climatic and sea surface temperature data. IDJ and EBM calculated water temperatures for lakes and streams, respectively. SJT, JRB, MSB, SB, PH, TTH, DJ, DIL, EBM and DM led analysis of specific data sets using code from PAH. SA, PJB, TMB, LC, THC-B, CD, ME, JME, SJGH, RH, JWP-H, LEBK, JMP, THS, PMT, IW and IJW derived phenological data for analysis, advised on interpretation, and assisted in assigning species traits. All co-authors commented on the manuscript.

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405 METHODS
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Data sets

We integrated data from many major UK biological monitoring schemes (Supplementary

Table 1), resulting in 10,003 long-term (at least 20-years between 1960 and 2012)

phenological series for 812 marine, freshwater and terrestrial taxa. The amassed data sets

included records for plants, phytoplankton, zooplankton, insects, amphibians, fish, mammals

and birds (379,081 individual phenological observations). For each study we used a single

population-level phenological measure per year (Supplementary Table 1). Since the sampling

resolution for the marine plankton data was monthly, prior to analysis we re-scaled these data

into units of days. Trophic level, taxonomic Class and environmental affinity were assigned

to each taxon, to permit analyses of correlations between these attributes and climate

sensitivity.

Daily air temperature and precipitation data were extracted from the Met Office National 419 Climate Information Centre (NCIC) 5km-resolution gridded data set³¹ for the spatial locations of all biological monitoring sites across the UK land surface. If available, recorded water temperatures from the same site were used in place of air temperatures, for phenological time series representing obligate aquatic taxa (freshwater plankton and fish). 423 Water temperatures were interpolated onto a daily time-step prior to analysis³². If these data were not available, daily water temperature data were estimated from air temperatures using a

fitted empirical site-specific relationship between air and water temperature. For the sea trout 426 (*Salmo trutta*) data, an existing linear relationship³³ was used, while for the Atlantic salmon 427 (*Salmo salar*) data, a non-linear relationship³⁴ was calculated for a nearby river, the Tarland Burn, and applied to air temperatures from the sampling site. For the marine plankton, mean monthly sea surface temperatures were extracted from the Met Office Hadley Centre Sea Ice 430 and Sea Surface Temperature (HadISST) data set³⁵ for each of the Standard Areas³⁶ in which phenological data were available. Precipitation data were not available for marine Standard Areas.

Statistics

Our analysis was conducted in two distinct phases (Supplementary Notes). Firstly, the CSP for each phenological series was estimated using generalized linear models to quantify associations between the timing of seasonal events and mean temperature and precipitation (within defined seasonal time windows) at the same location. Secondly, the phenological time series were aggregated and a single linear mixed effects (LME) model was run, capturing upper and lower limits of climate sensitivity across many species. CSPs for precipitation were not estimated for marine plankton data (see above), and so the second-phase LME models were run twice: once to examine correlations with temperature and precipitation for all but the marine plankton phenological series (9,800 series), and once to examine only correlations with temperature for the whole data set (10,003 series).

Phase 1: Estimating Climate Sensitivity Profiles (CSPs) for each time series

We used consistent methods to "screen" all phenological events with respect to their climate sensitivity, finding periods of the year in which temperature and precipitation have their most positive and negative correlations with seasonal timing (the upper and lower limits of climate sensitivity). This approach was flexible enough to detect when these limits represented opposing correlations between temperature or precipitation and seasonality, depending upon the seasonal timing of climatic change e.g. spring warming may advance budburst, but winter 453 warming may delay it³⁷ (Fig. 1, CSP type II). It could also detect when the direction of the correlation between climatic variables and seasonal timing was consistent irrespective of the seasonal timing of climatic change, with only the magnitude of the correlation varying between the limits of the CSP (Fig. 1, CSP types I and III).

For each phenological time series, we calculated the day of year by which 95% of the recorded seasonal events had occurred (*doy*95). Inter-annual variations in seasonal timing were statistically modelled as a function of daily mean temperatures on *doy*95 each year. Then, a series of 365 statistical models was run that used instead daily mean temperatures on *doy*95- 462 1 to dov_{95} -365 as predictors. Slope coefficients and R^2 values for the temperature terms in these models were collated, capturing seasonal variations in the sign and magnitude of the phenology-temperature relationship (i.e. the CSP, Fig. 1). Generalized Linear Models (GLMs) were used.

For two data sets (BTO Nest Record Scheme and PTES National Dormouse Monitoring Scheme, Supplementary Table 1) we modified the above analytical framework. In both of these schemes, the precise location of the biological observations changed among years (*cf* other schemes where monitoring sites are static over time). We extracted matching climatic

data for each specific location in each year, as for all other schemes, but then grouped the 472 bhenological and climatic data at county level (mean area $= 3.440 \text{ km}^2$). Then, for each taxon 473 in each county we used the fixed-effect slope parameters and R^2 values from a series of LME models, instead of GLMs, as a basis for estimating CSPs. In these models, we included fixed effects of temperature on *doy*95 to *doy*95-365 as before, and included a year random effect to account for replicate phenological records for each taxon in each county in each year. For the SAHFOS marine plankton data set, we modified our iterative approach to analyse seasonal timing-temperature relationships at monthly, instead of daily, time steps (the temporal resolution of the sea surface temperature data).

As a final step in estimating the CSP for each series, temporal variation in the sign and magnitude of the seasonal timing-temperature correlation was itself modelled (Extended Data Fig. 8). This was done by fitting Generalized Additive Models (GAMs, Gamma error 484 distribution) to the time series of slope coefficients and R^2 values from the models described above. By smoothing these time series, the GAMs identified periods of the year in which slope coefficients were consistently negative (i.e. warming advances seasonal timing), or consistently positive (i.e. warming delays seasonal timing), and during which the climate-phenology models generating the slope estimates had a their highest goodness-of-fit.

Seasonal "windows" in which the upper and lower limits of temperature sensitivity occurred were identified as periods during which 1) the 95% confidence interval for the GAM fitted to the slope coefficients surpassed the limits of the 2.5 and 97.5 percentiles of the original slope 493 coefficients and 2) the 95% confidence interval for the GAM fitted to the R^2 values surpassed 494 the 97.5 percentile of the original R^2 values. This ensured that seasonal windows were

defined by periods combining the greatest climate effect size and relatively strong predictive 496 power (determined by R^2). Using this framework, we identified the lower limit of CSP_{temp}: the period of the year in which an advancing effect of increasing temperature upon seasonal timing was most likely. This was estimated by determining when the 95% confidence interval of the GAM intersected the *lower* percentile of the seasonal timing-temperature slope coefficients, by "tracking" the most *negative* coefficients (Extended Data Fig. 8). In addition, 501 we identified the upper limit of CSP_{temp} by determining when the 95% confidence interval of the GAM intersected the *upper* percentile of the seasonal timing-temperature slope coefficients, by "tracking" the most *positive* (or least negative) coefficients. Excluding the marine plankton data, the whole modelling process was repeated with precipitation as a predictor instead of air temperature, culminating in the estimation of seasonal periods capturing the limits of phenological responses to changing precipitation.

After this process, temperature and precipitation were each averaged within the two seasonal windows in which the limits of phenological sensitivity occurred. With the exception of the marine plankton data, the final seasonal timing-climate model for each series was then fitted using a GLM with Gamma error distribution including four predictors: inter-annual variations 512 in 1) mean temperature during the period at the lower limit of CSP_{temp} , 2) mean temperature 513 during the period at the upper limit of CSP_{temp} , 3) mean precipitation during the period at the 514 lower limit of CSP_{precup} , 4) mean precipitation during the period at the upper limit of CSP_{precup} . For the marine plankton data, only the first two terms were fitted. For the BTO Nest Records and PTES National Dormouse Monitoring Scheme data sets we implemented these final models in a mixed effects framework with a random effect of year, as before. Therefore, although we modelled changes in statistical parameters (which are not estimated without error) to identify seasonal periods, this step was only used to find the original climatic data to

be used in subsequent modelling. Inferences were not, therefore, directly based upon statistical modelling of uncertain parameter estimates. We categorised the results of all 10,003 CSPs according to three broad response-types (CSP types I–III, Fig. 1), and retained P values for each fitted model term to infer which of the modelled climatic effects were statistically significant. We examined the evidence for trophic-level differences in the mean seasonal timing of climate sensitivity by modelling the relationship between the start date, end date and duration of the seasonal windows capturing the upper and lower limits of phenological sensitivity to temperature and rainfall as a function of trophic level (fixed effect), with random effects of phenological metric, within species, within site. Analyses 529 were conducted using the *base*, *mgcv* and *lme4* packages in R^{38-40} .

Phase 2: "Global" models of phenological climate sensitivity

We estimated the upper and lower limits of phenological climate sensitivity at a multi-species scale by "matching" each phenological series with data on mean temperature and precipitation, during the seasonal windows characterising the CSP for that series (Phase 1, above). We aggregated all 10,003 of these matched phenology-climate data sets. To quantify the average, multi-species, upper and lower limits of climate sensitivity we constructed a linear mixed effects (LME) model, in which phenology (day of year) was modelled as a function of mean temperature and precipitation within the seasonal windows of the amassed CSPs (fixed effects) with random effects of phenological metric, within species, within site. These random effects were necessary since our data could not be considered independent. The timing of events for the same species are more likely to be similar than for different species. Likewise for different sites and the phenological metric-types used to describe the events (e.g. first flight time or seasonal peak abundance). Random slopes and intercepts were

allowed to ensure that each phenological event, for a species at a site, was allowed a different rate of climate response.

For some species, more than one phenological event was recorded in the same year, at the same site. For example, butterflies may have more than one flight period in the same year, and plankton populations may be characterised by more than one seasonal abundance peak. As climate responses are unlikely to be the same for the first event of the year, and subsequent events, we introduced a voltinism factor in the analysis. This allowed us to distinguish between data representing the first/only events of each year (e.g. a spring plankton bloom or butterfly generation) and second events in each year (e.g. the subsequent summer plankton bloom or butterfly generation). This distinction captured all possibilities within our data set.

For site *i*, species *j*, voltmetric *k* (where voltmetric is a unique combination of voltinism class and the metric-type used to identify the event) the corresponding day of year (*DOY*) of a particular seasonal event is modelled as:

561
$$
DOY_{ijk} = \alpha_0 + \beta_1 T L_{ijk} + \beta_2 T U_{ijk} + \beta_3 P L_{ijk} + \beta_4 P U_{ijk} + \varepsilon_{ijk}
$$

562 where $\varepsilon_{ijk} \sim N(0, \sigma^2)$ and the model includes temperature at the upper limit of each CSP (*TU*), temperature at the lower limit of each CSP (*TL*), precipitation at the upper limit of each CSP (*PU*) and precipitation at the lower limit of each CSP (*PL*). Due to the non-independence within the data, we allow the intercepts and coefficients corresponding to all four covariates

to vary by site, species and voltmetric. Preserving the natural nesting of a metric for a species

at a particular site, this gives:

569
$$
\alpha_0 = \gamma_0 + \mu_{0; ijk} + \mu_{0; ijk} + \mu_{0; i, jk}
$$

570 $\beta_l = \gamma_l + \mu_l$; ijk + μ_l ; ijk + μ_l ; ijk

571 $\beta_2 = \gamma_2 + \mu_{2; ijk} + \mu_{2; ijk} + \mu_{2; i,jk}$

572 $\beta_3 = \gamma_3 + \mu_{3; i, j, k} + \mu_{3; i, j, k} + \mu_{3; i, j, k}$

573 $\beta_4 = \gamma_4 + \mu_4$ *; ijk* + μ_4 *; ijk*

575 where each of the μ terms is a random effect following: $\mu \sim N(0,\delta^2)$

This nesting of random effects is most conservative in terms of inference at the global level and is as flexible as possible, allowing each time series to have its own set of model parameters. This permits a high degree of biological realism since each distinct phenological event, for a given species, at a given site, is permitted to have a different slope for the effects of temperature and precipitation i.e. a different climate sensitivity.

In this model framework we are specifically testing the null hypotheses that each of the climate variables show no relation with seasonal timing of biological events. Because of this, and the fact that each parameter is estimated directly, without distributional form assumed *a priori* or as the target distribution, we follow a frequentist approach to analysis. However, because the exact degrees of freedom cannot be evaluated when using restricted maximum likelihood, hence no exact P-value, we present full summaries of all the parameters estimated 589 at species level (as given by: $\gamma + \mu_{i,j,k} + \mu_{i,j,k}$ above). Approximate P-values could be presented based on taking conservative estimates of the degrees of freedom though, given the volume of data available, this will typically lead to the detection of many statistically-significant results that may not be biologically significant. Examining the full range of estimated coefficients across the random effects levels ensures that we present the full range of variation around global parameters and can make more informed inference. In this way we encourage the reader to interpret our results by using biological insight, not by depending upon P-values alone.

To examine high-level differences in climate sensitivity among trophic levels and taxonomic groups we re-fitted the LME model with these attributes as fixed-effect factors, interacting with the fixed-effect climate variables. The fixed-effect slopes from the resulting models allowed us to compare differences in phenological climate sensitivity among these broad organism groups, averaged across all taxa within each group. Supplementary Table 2 shows the number of phenological series, sites and distinct taxa that contributed data to each of these groups. All models were run twice: once to examine correlations with both temperature and precipitation excluding marine plankton data (9,800 time series), and once to examine only temperature-phenology correlations for the whole data set (10,003 time series).

Potential biases

Data availability differed among taxonomic groups. To assess the extent to which mean responses were biased by data inequality we conducted Monte Carlo re-sampling, iteratively selecting 5, 20, 50 and 100 phenological series from each taxonomic group and re-fitting climate-phenology models with these sampled data sets. For taxonomic groups with less data than the larger sample sizes, we retained all available data (Supplementary Discussion). This allowed us to compare taxonomic group and trophic level responses based upon sampled and all data, to fully investigate potential bias.

Another potential bias in our analysis is that phenological time series length is variable, affecting the length of time over which climate-phenology correlations are assessed. In order to assess the extent to which differences in mean trophic level and taxonomic group responses are biased by variable time series length, we also re-fitted our models but based 621 only on pre- and post-1980 data. All models were run in the *lme4* package in $R^{38,40}$.

Estimating future change

To estimate potential future "net" effects of temperature and precipitation change, we compared predictions of seasonal timing under baseline conditions, and under established climate change scenarios. Firstly, estimates of seasonal timing (day of year) were obtained for the same baseline period used in the UKCP09 projections (long term average 1961-1990), using modelled correlations between phenology, temperature and precipitation (from Phase 1). Having obtained these baseline estimates, we applied our models to projected changes in monthly temperature and precipitation for the 2050s (UK Climate Projections, UKCP09, 631 http://ukclimateprojections.metoffice.gov.uk/). We used 10^{th} , 50^{th} and 90^{th} percentile changes

under low, medium and high emissions scenarios (relative to the 1961-90 baseline). The spatial location of each phenological series was matched to climate projection data for the 25 634×25 km grid square in which it occurred, and temporally matched to climatic data from the months-of-year in which its respective climate sensitivity windows occurred. Relative changes in timing, in response to climatic change of the magnitude projected to occur by the 2050s, were summarised by trophic levels and taxonomic groups.

REFERENCES

40. Bates, D., Maechler, M. & Bolker, B. lme4: Linear mixed-effects models using S4 classes. (2011).

FIGURE LEGENDS

Figure 1 | Climate Sensitivity Profiles (CSPs). Climate sensitivity is the change in seasonal 666 timing per unit change in temperature (days ${}^{\circ}C^{-1}$) or precipitation (days mm⁻¹). Irrespective of the date, increasing temperature/precipitation may always correlate with earlier (red curve, CSP type I) or later (blue curve, CSP type III), biological events, but sensitivity to climate variation (correlation magnitude) differs (*cf* w1 and w2, w5 and w6). In contrast, opposing climate-phenology correlations may occur, depending on the date at which climate changes (orange curve, w3 and w4, CSP type II). Panels show hypothetical relationships for seasonal windows w1-w6.

Figure 2 | Climatic change in the UK, 1960-2012. a) Long-term changes in air temperature and precipitation are the differences between 1960 and 2012 monthly means of these variables, derived from a regression fitted through each monthly time series. Error bars indicate the standard deviation of linearly-detrended climatological data, as an indication of inter-annual variation around each trend. b) and c) Examples of spatial variation in the extent of long-term climatic changes are shown for March air temperatures and February precipitation.

Figure 3 | Upper and lower limits of phenological climate sensitivity. Sensitivity is the slope of the relationship between seasonal timing (day of year) and climatic variables. All-684 taxa upper and lower limits in a) temperature (\degree C) and b) precipitation (mm day⁻¹) sensitivity 685 are summarised. Lower (c, d) and upper (e, f) limits of temperature (c, e) and precipitation (d, f) f) sensitivity are shown by trophic level. Inverted triangles indicate average sensitivity. Curves are kernel density plots: estimates of the probability density distribution of species-level climate sensitivity i.e. the relative likelihood of different levels of climate sensitivity 689 within each species group $(n = 370, 725)$.

Figure 4 | Upper and lower limits of phenological climate sensitivity for broad taxonomic groups. Lower (blue) and upper (red) limits of the sensitivity of phenological events to seasonal temperature (a) and precipitation (b) change are shown. Coloured circles 694 indicate the median response, and bars show the $5th$ -to-95th percentile responses for each group. Sensitivity is quantified by summarising the species-level (random effects) responses from a mixed effects model including data for all taxa, and with taxonomic group as a fixed effect (*n* = 370,725).

Figure 5 | Estimated phenological shifts by the 2050s. Modelled responses to projected temperature and precipitation change, assuming contemporary climate sensitivity, for trophic levels (a) and taxonomic groups (b). Projected median shifts in seasonal timing are shown. Change estimates are based on low, medium and high emissions climate scenarios. Bars 703 represent median responses to $50th$ percentile climate change projections under each scenario, 704 while extremes of whiskers represent median responses to $10th$ and $90th$ percentile projected climatic changes under each scenario. Standard deviations indicate variation in projected 706 responses for each group under the $50th$ percentile of the medium emissions scenario.

EXTENDED DATA FIGURE AND TABLE LEGENDS

Extended Data Figure 1 | Limits of phenological temperature sensitivity inclusive of marine plankton data. Upper and lower limits of phenological temperature sensitivity are quantified as the slope of the relationship between seasonal timing (day of year) and temperature (°C) variation within specific seasonal periods. Limits in temperature sensitivity are shown for all taxa (a) and by trophic level (lower limit b, upper limit c). Inverted triangles indicate average sensitivity for all species in each group and curves are probability density plots of species-level variation in sensitivity.

Extended Data Figure 2 | Limits of phenological climate sensitivity for taxonomic groups (top) and trophic levels (bottom), after Monte-Carlo resampling. Lower (blue) and upper (red) limits of the sensitivity of phenological events to seasonal temperature (a) and precipitation (b) change. Coloured circles: responses based upon the full data set. Bars: 723 2.5th-to-97.5th percentile responses for each group, based upon 100 draws from the full data set. Data were sampled so that 5, (dotted bar), 20 (solid bar), 50 (dashed bar) and 100 (dot-dashed bar) phenological time series were drawn from each taxonomic group.

Extended Data Figure 3 | Climate sensitivities, based on different time periods (top: all data, middle: pre-1980 data, bottom: post-1980 data). Sensitivity is the slope of the 729 relationship between seasonal timing (day of year) and temperature $(^{\circ}C)$, or precipitation $(730 \, \text{m})$ (mm day⁻¹). Limits of a) temperature and b) precipitation sensitivity are summarised for all taxa. Lower (c, d) and upper (e, f) limits of temperature (c, e) and precipitation (d, f) sensitivity are shown by trophic level. Inverted triangles: average sensitivity for all species (a,

b) or trophic levels (c-f). Curves: kernel density plots: probability density distributions of species-level climate sensitivity i.e. the relative likelihood of different climate sensitivities within each species group.

Extended Data Figure 4 | Limits of phenological climate sensitivity for broad taxonomic groups (top: all data, bottom: post-1980 data only). Lower (blue) and upper (red) limits of the sensitivity of phenological events to seasonal temperature (a) and precipitation (b) change 740 are shown. Coloured circles indicate the median response, and bars show the $5th$ -to-95th percentile responses for each group. Sensitivity is quantified by summarising the species-level (random effects) responses from a mixed effects model including data for all taxa, and with taxonomic group as a fixed effect.

Extended Data Figure 5 | Seasonal windows for Climate Sensitivity Profiles (CSPs). Estimated climatic sensitivity at the lower (a, c) and upper (b, d) limits of CSPs for 10,003 phenological series. Grey lines are seasonal time periods (x axis) within which climatic variables have their most positive/negative correlations with the seasonal timing of each phenological event. The y-axis indicates the slope coefficient for each of these correlations; a 750 measure of climate sensitivity (days change ${}^{\circ}C^{-1}$, or mm⁻¹). Shown are the lower/upper limits 751 of CSP_{temp} (a, b, respectively) and the lower/upper limits of CSP_{precin} (c, d, respectively). Inset histograms show seasonal time window length (days).

Extended Data Figure 6 | Time lags between phenological events and seasonal windows of climate sensitivity. Frequency histograms showing the time lag (in days) between the mean timing of each phenological event and the end of seasonal windows corresponding to 757 the lower and upper limits of CSP_{temp} (a, b, respectively) and the lower and upper limits of CSP_{treei} (c, d, respectively). Peaks at lags of around 1 year are where windows were identified that ended at the mean seasonal timing of an event, but in the previous year, due to temporal autocorrelation in climate data.

Extended Data Figure 7 | Seasonal windows for Climate Sensitivity Profiles (CSPs) by trophic level. Estimated climatic sensitivity at the lower and upper limits of CSPs for taxa at each of three trophic levels. Formatting is the same as in Extended Data Figure 5. Shown are 765 the lower and upper limits of CSP_{temp} (a, b, respectively) and the lower and upper limits of $CSP_{precision}$ (c, d, respectively).

Extended Data Figure 8 | Example Climate Sensitivity Profile (CSP). Temperature 769 sensitivity (CSP_{temp}) for alderfly (*Sialis lutaria*) emergence from Windermere, UK. Solid black line: sensitivity of first emergence to water temperature on different days of the year 771 (days change $^{\circ}C^{-1}$). Grey horizontal lines: 2.5 and 97.5 percentiles of these sensitivity values. Solid orange curve: GAM smoother fitted through the sensitivity values with associated confidence intervals (dashed orange curves). Horizontal bars indicate where GAM confidence intervals exceed the percentiles of the original sensitivity values, indicating seasonal windows at the limits of the climate sensitivity profile.

Extended Data Table 1 | Modelled relationships between seasonal timing and climate variables for n=10,003 phenological time series. Climate Sensitivity Profiles (CSPs) fall

within three broad response-types; events always advance with increases in the climate variable irrespective of the seasonal timing of climate change (CSP Type I, Fig. 1 - red curve), events are always delayed by increases in the climate variable irrespective of the seasonal timing of climate change (CSP Type III, Fig. 1 - blue curve), and events may be advanced or delayed by increases in the climate variable, depending on the seasonal timing of climate change (CSP Type II, Fig. 1 - orange curve). Shown are the percentage of series that fall in each Type (% series), the percentage of effects that are statistically significant at 786 P<0.05 after multiple testing correction (% effects significant). \Box Based only on freshwater and terrestrial taxa, for which precipitation data were available. † NA indicates effect not evaluated, due to lack of precipitation data for marine taxa

Extended Data Table 2 | Parameter estimates and test statistics from climate-phenology mixed-effects models. Presented are fixed-effect parameter estimates from each model; the intercept and slope for each climatic predictor. Following R convention, absolute parameter estimates are provided for an assigned "baseline" group within each model (*b*), and remaining estimates are given as differences from this baseline (*Δb*). Each estimate has an associated standard error and *t* statistic in parentheses (standard error, *t*). Climatic predictors include mean temperature and precipitation in seasonal windows at the upper and lower limit of the climate sensitivity profile for each phenological series. The number of observations, *n*, is 370,725. □ Models were re-run including the marine plankton data, and excluding 799 precipitation effects (see text). In these models the number of observations, $n = 379,081$

