- **1** Phenological sensitivity to climate across taxa and trophic levels
- 2

Among-species differences in phenological responses to climate change can 3 desynchronise ecological interactions, threatening ecosystem function. To assess these 4 threats, we must quantify the relative impact of climate change on species at different 5 trophic levels. Here, we apply a novel Climate Sensitivity Profile approach to 10,003 6 terrestrial and aquatic phenological data sets, spatially-matched to temperature and 7 precipitation data, quantifying variation in climate sensitivity. The direction, magnitude 8 and timing of climate sensitivity varied markedly among organisms within taxonomic 9 10 and trophic groups. Despite this, we detected systematic variation in the direction and 11 magnitude of phenological climate sensitivity. Secondary consumers showed 12 consistently lower climate sensitivity than other groups. Based upon mid-century 13 climate change projections, we estimate that the timing of phenological events could 14 change more for primary consumers than for other trophic levels (6.2 versus 2.5 - 2.9 days earlier on average), with substantial taxonomic variation (1.1 - 14.8 days earlier on 15 average). 16

17

Numerous long-term ecological changes have been attributed to climate change¹. Shifts in the seasonal timing of recurring biological events such as reproduction and migration (i.e. phenological changes) are especially well documented^{2,3}. Long-term ecosystem studies^{4–7} and global meta-analyses^{2,3,8} have demonstrated that many spring and summer phenological events now occur earlier in the year. Substantial among-species variation in responses has fuelled concerns that key seasonal species interactions may desynchronise over time, with potentially severe consequences for wild populations and, hence, for ecosystem functioning⁹.

Identifying systematic taxonomic and trait-based differences in phenological climate
sensitivity (i.e. change in seasonal timing per unit change in climatic conditions) would have
significant socio-ecological implications. This would afford some predictability to future
ecological outcomes and would identify species that represent effective sentinels of climate
impact, facilitating the development of indicators and estimates of vulnerability for
conservation and national adaptation programmes^{10–12}. Unfortunately, such generalisations
are currently elusive.

33

25

34 Analytical approach and data sets

Among-species differences in phenological change may arise from two aspects of climate 35 36 sensitivity. Firstly, variation may reflect differences in physiological and behavioural 37 responses, microclimate use, and the importance of non-climate related cues, such as photoperiod¹³ or resource availability¹⁴. Therefore, even when species have the same seasonal 38 period (window) for which they are most sensitive to climate change, they show different 39 40 phenological responses to a given climatic change. Secondly, co-occurring species may vary in their seasonal periods of climate sensitivity, each typified by different levels of directional 41 climate change $^{15-17}$. We conceptualise these two aspects of phenological responses as 42 43 species- (or population-) specific Climate Sensitivity Profiles (CSPs, Fig. 1). The CSP 44 approach differs fundamentally from attempts to identify single "critical" seasonal periods within which climatic change most strongly affects seasonal events¹⁷, by quantifying the full 45 range of phenological responses to seasonal climatic change. We ask "How sensitive are 46 phenological events to temperature and precipitation change at different times of year?". By 47 applying this approach to a large, taxonomically-diverse national-scale data set, we discern 48

coherent patterns within a multitude of idiosyncratic biological climate-responses. We assess
whether systematic differences in climate sensitivity underpin differences in phenological
change among taxonomic and trophic groups in the UK⁸.

52

53	We elected against using published climate responses that may be biased in favour of species
54	showing an effect. Instead, we analysed 10,003 long-term (≥20 year) phenological time series
55	for 812 marine, freshwater and terrestrial taxa over the period 1960-2012. Our data set
56	aggregates many of the UK's foremost long-term biological monitoring schemes
57	(Supplementary Table 1), including phenological information on amphibians (spawning),
58	birds (egg laying, migration), planktonic crustaceans (population peaks), fish (spawning,
59	migration), insects (flight periods), mammals (birth dates), phytoplankton (population peaks)
60	and plants (flowering, fruiting, leafing). These taxa represent three broad trophic levels:
61	primary producers (phytoplankton, plants), primary consumers (granivorous birds,
62	herbivorous insects, mammals, planktonic crustaceans) and secondary consumers (predatory
63	amphibians, birds, fish, insects, mammals, planktonic crustaceans). We spatially-matched all
64	10,003 phenological time series with local temperature and precipitation data from a 5×5 km
65	resolution gridded data set, before statistically modelling the relationship between seasonal
66	timing and climatic variables. Between 1960 and 2012 mean UK air temperatures increased
67	in all months, and mean precipitation increased in most months (Fig. 2a).

68

Spatial variability in climatic change (Fig. 2b,c), necessitates local matching of phenological
and climatic datasets rather than the use of regionally-averaged climate data (e.g. Central
England Temperatures) or large-scale climatic indicators (e.g. North Atlantic Oscillation).
We did not make the restrictive assumption that biological events would be related to annual

73 mean climatic conditions, or to conditions within periods based upon calendar months. Our 74 CSP approach identified seasonal periods within which climatic change has its most positive 75 and negative correlations with phenology (hereafter referred to as upper and lower limits of 76 climate sensitivity, respectively). We could identify, for each phenological series, up to two 77 seasonal periods within which climatic variation had a significant correlation with seasonal 78 timing. The method was flexible enough to 1) allow situations in which climatic variation 79 within only a single period had a significant correlation, and 2) identify seasonal windows 80 ranging from a few days to a whole year in length. Our analysis captured the idiosyncrasies 81 of phenological responses, allowed their categorisation into generic types of climate 82 response, and is consistent with current biological understanding of climate-phenology relationships^{15,16}. 83

84

85 Climate response-types in the UK

CSPs fall into three categories. The qualitative type of climate-phenology correlation 86 87 (positive or negative) may remain consistent, irrespective of when in the year climatic change 88 occurs. In this case only the magnitude of the phenological response differs with the time of year at which climatic variables change. The climate-phenology correlation may be 89 90 consistently negative (CSP type I, Fig. 1, red curve) or positive (CSP type III, Fig. 1, blue 91 curve). Alternatively, opposing correlations between seasonal climatic change and the timing 92 of biological events may exist i.e. the direction and magnitude of the phenological response 93 varies (CSP type II, Fig. 1, orange curve). We determined CSPs for responses to temperature 94 (CSP_{temp}) and precipitation (CSP_{precip}).

95

96 Focusing on temperature, CSP type II was most common (Extended Data Table 1, 69.7 % of 97 phenological series): seasonal events were advanced by (i.e. negatively correlated with) 98 warming during one period of the year, and delayed by (i.e. positively correlated with) 99 warming in another period. After multiple testing correction, 44.8% of the observed 100 phenological advances (but only 1.0% of delays) with warming were statistically significant 101 (P < 0.05). CSP type I was the next most common response-type: warming in different 102 seasonal windows was consistently correlated with earlier seasonal events (i.e. negative 103 correlations, 24.7% of series). In this case the lower and upper limits of CSPs represent the 104 "strongest" and "weakest" phenological advances with warming, respectively, and 58.1% of 105 the "strongest" responses were statistically significant (P<0.05, correcting for multiple 106 testing).

107

Phenological events most commonly demonstrated opposing (Fig. 1, CSP type II, 53.0% of
series) or consistently positive (Fig. 1, CSP type III, 28.0% of phenological series)
correlations with increasing seasonal precipitation. Though delayed phenological events may
commonly be associated with higher precipitation (81.0% of events show this type of
response), few of these associations were significant (Extended Data Table 1).

113

114 Climate sensitivity at the UK-scale

We matched each phenological series with four climate variables: mean temperature during the seasonal windows at the upper and lower limits of CSP_{temp}, and similarly-averaged precipitation data for the seasonal windows at the upper and lower limits of CSP_{precip}. We then combined all 10,003 phenological series and their matched climate data, and modelled

the relationships between seasonal timing and climate variables using linear mixed effects
(LME) models. Initially we fitted a "global" model to quantify upper and lower limits of
temperature and precipitation sensitivity, averaged across all phenological events. Marine
plankton data were excluded at this stage, due to a lack of precipitation data.

123

Most phenological events occurred earlier with seasonal warming (average rate -2.6 days °C-124 ¹, Fig. 3a, Extended Data Table 2). Variation in the strength of this correlation was similar 125 126 among sites and species (random-effects variances in site and species level seasonal timing -127 temperature slopes were 2.1 and 1.9, respectively). Some phenological events occurred later with seasonal warming (Fig. 3a) though, in other cases, the upper limit of CSP_{temp} was in fact 128 129 a "weak" advance with warming. The upper limit of temperature sensitivity was more 130 variable among species than sites (random effects variances in species and site level seasonal 131 timing – temperature slopes were 2.3 and 0.4, respectively). Averaged across species and 132 populations, temperature responses were most consistent with CSP type II.

133

134 Most phenological events showed opposing responses to increasing seasonal precipitation 135 (Fig. 1, CSP type II). The tendency for delays with rising precipitation was greatest: the average upper limit of CSP_{precip} exceeded the lower limit (1.4 days mm⁻¹ and -0.4 days mm⁻¹, 136 respectively, Fig. 3b, Extended Data Table 2). The upper limit of CSP_{precip} was more variable 137 138 among species than sites (species and site level random-effects variances in the seasonal 139 timing – precipitation slopes were 1.9 and 1.2, respectively). The fitted climate-phenology 140 model was better supported by the data than a year-only model with the same random effects 141 structure (delta-AIC 293,516). This indicates the presence of real associations between 142 climate and seasonality, rather than purely spurious correlations due to shared temporal

trends. Average sensitivity to temperature was very similar in the model that included marine
plankton data, but excluded precipitation effects (Supplementary Discussion, Extended Data
Fig. 1).

146

147 Taxonomic and trophic group sensitivity

148 We tested the hypothesis that the limits of seasonal climate sensitivity differ coherently 149 among taxonomic groups by including a fixed-effect interaction between taxonomic group 150 and each climatic variable (Fig. 4, Extended Data Table 2). The lower limit of CSP_{temp} was negative for all groups ("earliness" with warming), the strongest responses being found for 151 plants, freshwater phytoplankton, insects and amphibians (4.3, 4.1, 3.7 and 3.4 days earlier 152 $^{\circ}C^{-1}$, respectively). Upper limits of CSP_{temp} indicated that freshwater phytoplankton and 153 154 mammals experienced the greatest phenological delays with seasonal warming (2.9 and 2.0 days later $^{\circ}C^{-1}$, respectively) but that plants showed little evidence of such delays. The 155 156 strongest phenological delays with rising seasonal precipitation were found for freshwater phytoplankton and insects (2.5 and 2.2 days later mm⁻¹, respectively), while freshwater 157 158 phytoplankton also exhibited the strongest phenological advances with rising precipitation during other seasonal windows (1.1 days earlier mm⁻¹). Average temperature and 159 160 precipitation responses were consistent with a CSP type II in most cases. There was 161 considerable within-group variability in sensitivity.

162

We examined trophic-level differences in climate sensitivity by including this in interaction with each climate variable in the global model. The lower limit of CSP_{temp} showed greater systematic variation among trophic levels than the upper limit (Fig. 3c,e). The tendency

166	towards "earliness" with seasonal warming was strongest at lower trophic levels (-4.1, -3.7
167	and -1.9 days °C ⁻¹ for primary producers, primary consumers and secondary consumers,
168	respectively, Extended Data Table 2), consistent with observations of more rapid
169	phenological changes at lower trophic levels, in the UK ⁸ . Conversely, the lower limit of
170	CSP _{precip} varied less among trophic levels than the upper limit (Fig. 3d,f). The tendency for
171	later seasonal events with higher seasonal precipitation was greater for primary producers and
172	primary consumers (1.8 and 2.2 days mm ⁻¹ on average, respectively) than for secondary
173	consumers (1.0 days mm ⁻¹). Variations in climate sensitivity were described more
174	parsimoniously by taxonomic groups than by trophic levels (AICs of taxonomic and trophic-
175	level models 3237611 and 3238061, respectively).
176	
176 177	Results were little-affected when analysing only pre- and post-1980 data, to minimise among-
176 177 178	Results were little-affected when analysing only pre- and post-1980 data, to minimise among- group variation in time series length, and after Monte Carlo re-sampling to assess the
176 177 178 179	Results were little-affected when analysing only pre- and post-1980 data, to minimise among- group variation in time series length, and after Monte Carlo re-sampling to assess the potential effects of taxonomic bias (Supplementary Discussion, Extended Data Figs. 2-4).
176 177 178 179 180	Results were little-affected when analysing only pre- and post-1980 data, to minimise among- group variation in time series length, and after Monte Carlo re-sampling to assess the potential effects of taxonomic bias (Supplementary Discussion, Extended Data Figs. 2-4). The same qualitative trophic-level differences in climate sensitivity were apparent when
176 177 178 179 180 181	Results were little-affected when analysing only pre- and post-1980 data, to minimise among- group variation in time series length, and after Monte Carlo re-sampling to assess the potential effects of taxonomic bias (Supplementary Discussion, Extended Data Figs. 2-4). The same qualitative trophic-level differences in climate sensitivity were apparent when including marine plankton data in a temperature-only LME model (Supplementary
176 177 178 179 180 181 182	Results were little-affected when analysing only pre- and post-1980 data, to minimise among- group variation in time series length, and after Monte Carlo re-sampling to assess the potential effects of taxonomic bias (Supplementary Discussion, Extended Data Figs. 2-4). The same qualitative trophic-level differences in climate sensitivity were apparent when including marine plankton data in a temperature-only LME model (Supplementary Discussion, Extended Data Fig. 1). In contrast to trophic-level differences in the magnitude of
176 177 178 179 180 181 182 183	Results were little-affected when analysing only pre- and post-1980 data, to minimise among- group variation in time series length, and after Monte Carlo re-sampling to assess the potential effects of taxonomic bias (Supplementary Discussion, Extended Data Figs. 2-4). The same qualitative trophic-level differences in climate sensitivity were apparent when including marine plankton data in a temperature-only LME model (Supplementary Discussion, Extended Data Fig. 1). In contrast to trophic-level differences in the magnitude of sensitivity, there was little evidence of similar variation in the seasonal timing of climate
176 177 178 179 180 181 182 183 184	Results were little-affected when analysing only pre- and post-1980 data, to minimise among- group variation in time series length, and after Monte Carlo re-sampling to assess the potential effects of taxonomic bias (Supplementary Discussion, Extended Data Figs. 2-4). The same qualitative trophic-level differences in climate sensitivity were apparent when including marine plankton data in a temperature-only LME model (Supplementary Discussion, Extended Data Fig. 1). In contrast to trophic-level differences in the magnitude of sensitivity, there was little evidence of similar variation in the seasonal timing of climate sensitivity (Supplementary Discussion, Extended Data Figs. 5-7).

185

186 Estimating future change

187 Overall, "net", phenological responses to climatic change combine potentially-opposing
188 responses to conditions in different seasonal periods. We estimated "net" responses by the

189 2050s by applying our fitted models to UKCP09 probabilistic projections (bias-corrected 190 relative to a 1961-90 baseline) of temperature and precipitation change under low, medium 191 and high emissions scenarios. Rather than predicting the absolute timing of future 192 phenological events, we contrasted possible *changes* in seasonal timing among organism 193 groups based upon established climate scenarios and contemporary patterns of climate 194 sensitivity. Estimated average phenological changes for primary producers and secondary 195 consumers were less than those for primary consumers (Fig. 5a). This occurred because, 196 averaged across species, the opposing climate responses of primary producers and secondary 197 consumers are more similar in magnitude than are those for primary consumers (Fig. 3), 198 effectively "cancelling each other out". Our models suggest greater average advances for 199 crustacea, fish and insects than for other groups, such as freshwater phytoplankton, birds and 200 mammals (Fig. 5b). However, response-variation is high for crustacea (Fig. 5b).

201

202 Discussion

203 In the UK, phenological climate sensitivity varies greatly, suggesting effects of locallyvarving non-climatic drivers such as population structure¹⁸, resource availability¹⁹ and 204 adaptation²⁰. This is relevant to the use of phenological change as a tangible climate change 205 indicator^{1,21}. Mediators of phenological climate sensitivity are only known locally for some 206 of the groups in our data set e.g. nutrient availability (freshwater phytoplankton)²². However, 207 208 for others, the climate sensitivity of different biological traits is known to be mediated by alternative drivers^{23,24}. High climate-response variability necessitates wide site and species 209 210 coverage in long-term monitoring schemes aiming to develop robust aggregate indicators of change²¹. Since climatic conditions are more spatially-variable across broader geographic 211 212 domains, site-level replication of phenological monitoring is particularly important when

213 interpreting phenology at continental to global scales. In the UK, average responses for fish 214 and insects appear to provide climate-indicator potential. These groups show consistently 215 strong phenological advances with seasonal warming, and only weak opposing responses, 216 resulting in relatively large (net) changes in seasonal timing. Interpretation of phenological 217 changes for other groups is more complex. For example, freshwater phytoplankton show 218 strong evidence of opposing phenological responses to climatic variation at different times of 219 year and these are near-equivalent in magnitude, such that estimated net changes are 220 negligible. This highlights that long-term observations represent the net effect of potentiallyopposing biological responses²⁵. To fully capitalise on the indicator potential of phenological 221 222 change, we must advance mechanistic understanding of responses to potentially opposing 223 climate and non-climate drivers.

224

225 Despite this variability, we identified coherent patterns in climate sensitivity among the 226 idiosyncratic responses of many wild plant and animal populations. For the first time we 227 show that, on average, trophic levels differ in the magnitude of seasonal climate sensitivity, 228 but not the time-of-year within which climatic change has its most pronounced effects. This 229 may be a key mechanism underpinning observations of trophic level differences in phenological change in the UK⁸. Lower trophic levels demonstrated more pronounced 230 231 variation in their sensitivity to changing temperature and precipitation at different times of 232 year, and stronger phenological responses to climatic change during defined (taxon- and 233 population-specific) seasonal periods.

234

In response to climatic changes projected for the 2050s, relative changes in seasonal timingare likely to be greatest for primary consumers, particularly in the terrestrial environment.

237 The difference in magnitude between opposing climate responses is greatest for primary 238 consumers, resulting in greater "net" change. Our approach makes the simplifying 239 assumption that climatic change has the overriding influence upon seasonality. Nevertheless, 240 this suggests that systematic differences in climate sensitivity could result in widespread 241 phenological desynchronisation. However, factors that shape phenological climate-responses introduce uncertainty into projections of future phenological change. These results should 242 243 catalyse research to improve predictive capacity in the face of multiple environmental and 244 demographic drivers that not only mediate rates of change, but might also confer resilience to desynchronisation e.g. population density-dependence²⁶, compensatory range shifts²⁷, and the 245 formation of novel inter-specific interactions^{28,29}. These findings also underscore the 246 247 importance of developing our capacity to manage ecosystems within a "safe operating space" with respect to the likely impacts of projected climate change 30 . 248

249

250 **Supplementary Information** is linked to the online version of the paper at

251 www.nature.com/nature.

252

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represented by the authorship list hold long-term data that represent a considerable

262 investment in scientific endeavour. Whilst we are committed to sharing these data for

scientific research, users are requested to collaborate before publication of these data to

ensure accurate biological interpretation. We thank four referees for their comments.

265

266 Author Contributions

267 SJT and SW conceived and co-ordinated the study, and led writing of the manuscript. PAH

developed the analysis routine and wrote statistical code to be applied to all data sets. DH

extracted all climatic and sea surface temperature data. IDJ and EBM calculated water

270 temperatures for lakes and streams, respectively. SJT, JRB, MSB, SB, PH, TTH, DJ, DIL,

EBM and DM led analysis of specific data sets using code from PAH. SA, PJB, TMB, LC,

272 THC-B, CD, ME, JME, SJGH, RH, JWP-H, LEBK, JMP, THS, PMT, IW and IJW derived

273 phenological data for analysis, advised on interpretation, and assisted in assigning species

traits. All co-authors commented on the manuscript.

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405 METHODS
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406 Data sets

407 We integrated data from many major UK biological monitoring schemes (Supplementary

408 Table 1), resulting in 10,003 long-term (at least 20-years between 1960 and 2012)

409 phenological series for 812 marine, freshwater and terrestrial taxa. The amassed data sets

410 included records for plants, phytoplankton, zooplankton, insects, amphibians, fish, mammals

and birds (379,081 individual phenological observations). For each study we used a single

412 population-level phenological measure per year (Supplementary Table 1). Since the sampling

resolution for the marine plankton data was monthly, prior to analysis we re-scaled these data

414 into units of days. Trophic level, taxonomic Class and environmental affinity were assigned

to each taxon, to permit analyses of correlations between these attributes and climate

416 sensitivity.

417



425 fitted empirical site-specific relationship between air and water temperature. For the sea trout (Salmo trutta) data, an existing linear relationship³³ was used, while for the Atlantic salmon 426 (Salmo salar) data, a non-linear relationship³⁴ was calculated for a nearby river, the Tarland 427 428 Burn, and applied to air temperatures from the sampling site. For the marine plankton, mean 429 monthly sea surface temperatures were extracted from the Met Office Hadley Centre Sea Ice and Sea Surface Temperature (HadISST) data set³⁵ for each of the Standard Areas³⁶ in which 430 431 phenological data were available. Precipitation data were not available for marine Standard 432 Areas.

433

434 **Statistics**

435 Our analysis was conducted in two distinct phases (Supplementary Notes). Firstly, the CSP 436 for each phenological series was estimated using generalized linear models to quantify 437 associations between the timing of seasonal events and mean temperature and precipitation 438 (within defined seasonal time windows) at the same location. Secondly, the phenological time 439 series were aggregated and a single linear mixed effects (LME) model was run, capturing 440 upper and lower limits of climate sensitivity across many species. CSPs for precipitation were 441 not estimated for marine plankton data (see above), and so the second-phase LME models 442 were run twice: once to examine correlations with temperature and precipitation for all but 443 the marine plankton phenological series (9,800 series), and once to examine only correlations 444 with temperature for the whole data set (10,003 series).

445

446 Phase 1: Estimating Climate Sensitivity Profiles (CSPs) for each time series

447 We used consistent methods to "screen" all phenological events with respect to their climate sensitivity, finding periods of the year in which temperature and precipitation have their most 448 449 positive and negative correlations with seasonal timing (the upper and lower limits of climate 450 sensitivity). This approach was flexible enough to detect when these limits represented 451 opposing correlations between temperature or precipitation and seasonality, depending upon the seasonal timing of climatic change e.g. spring warming may advance budburst, but winter 452 warming may delay it³⁷ (Fig. 1, CSP type II). It could also detect when the direction of the 453 454 correlation between climatic variables and seasonal timing was consistent irrespective of the 455 seasonal timing of climatic change, with only the magnitude of the correlation varying 456 between the limits of the CSP (Fig. 1, CSP types I and III).

457

458 For each phenological time series, we calculated the day of year by which 95% of the 459 recorded seasonal events had occurred (doy₉₅). Inter-annual variations in seasonal timing 460 were statistically modelled as a function of daily mean temperatures on doy_{95} each year Then, 461 a series of 365 statistical models was run that used instead daily mean temperatures on doy_{95} -1 to dov_{95} -365 as predictors. Slope coefficients and R² values for the temperature terms in 462 463 these models were collated, capturing seasonal variations in the sign and magnitude of the 464 phenology-temperature relationship (i.e. the CSP, Fig. 1). Generalized Linear Models (GLMs) were used. 465

466

For two data sets (BTO Nest Record Scheme and PTES National Dormouse Monitoring Scheme, Supplementary Table 1) we modified the above analytical framework. In both of these schemes, the precise location of the biological observations changed among years (*cf* other schemes where monitoring sites are static over time). We extracted matching climatic

data for each specific location in each year, as for all other schemes, but then grouped the 471 phenological and climatic data at county level (mean area = 3.440 km^2). Then, for each taxon 472 in each county we used the fixed-effect slope parameters and R^2 values from a series of LME 473 models, instead of GLMs, as a basis for estimating CSPs. In these models, we included fixed 474 475 effects of temperature on doy_{95} to doy_{95} -365 as before, and included a year random effect to 476 account for replicate phenological records for each taxon in each county in each year. For the 477 SAHFOS marine plankton data set, we modified our iterative approach to analyse seasonal timing-temperature relationships at monthly, instead of daily, time steps (the temporal 478 479 resolution of the sea surface temperature data).

480

481 As a final step in estimating the CSP for each series, temporal variation in the sign and 482 magnitude of the seasonal timing-temperature correlation was itself modelled (Extended Data 483 Fig. 8). This was done by fitting Generalized Additive Models (GAMs, Gamma error distribution) to the time series of slope coefficients and R^2 values from the models described 484 485 above. By smoothing these time series, the GAMs identified periods of the year in which slope coefficients were consistently negative (i.e. warming advances seasonal timing), or 486 consistently positive (i.e. warming delays seasonal timing), and during which the climate-487 488 phenology models generating the slope estimates had a their highest goodness-of-fit.

489

Seasonal "windows" in which the upper and lower limits of temperature sensitivity occurred were identified as periods during which 1) the 95% confidence interval for the GAM fitted to the slope coefficients surpassed the limits of the 2.5 and 97.5 percentiles of the original slope coefficients and 2) the 95% confidence interval for the GAM fitted to the R^2 values surpassed the 97.5 percentile of the original R^2 values. This ensured that seasonal windows were 495 defined by periods combining the greatest climate effect size and relatively strong predictive power (determined by R^2). Using this framework, we identified the lower limit of CSP_{temp} : 496 497 the period of the year in which an advancing effect of increasing temperature upon seasonal 498 timing was most likely. This was estimated by determining when the 95% confidence interval 499 of the GAM intersected the *lower* percentile of the seasonal timing-temperature slope coefficients, by "tracking" the most negative coefficients (Extended Data Fig. 8). In addition, 500 501 we identified the upper limit of CSP_{temp} by determining when the 95% confidence interval of 502 the GAM intersected the *upper* percentile of the seasonal timing-temperature slope 503 coefficients, by "tracking" the most *positive* (or least negative) coefficients. Excluding the 504 marine plankton data, the whole modelling process was repeated with precipitation as a 505 predictor instead of air temperature, culminating in the estimation of seasonal periods 506 capturing the limits of phenological responses to changing precipitation.

507

508 After this process, temperature and precipitation were each averaged within the two seasonal 509 windows in which the limits of phenological sensitivity occurred. With the exception of the 510 marine plankton data, the final seasonal timing-climate model for each series was then fitted 511 using a GLM with Gamma error distribution including four predictors: inter-annual variations 512 in 1) mean temperature during the period at the lower limit of CSP_{temp}, 2) mean temperature 513 during the period at the upper limit of CSP_{temp}, 3) mean precipitation during the period at the 514 lower limit of CSP_{precip}, 4) mean precipitation during the period at the upper limit of CSP_{precip}. For the marine plankton data, only the first two terms were fitted. For the BTO Nest Records 515 and PTES National Dormouse Monitoring Scheme data sets we implemented these final 516 517 models in a mixed effects framework with a random effect of year, as before. Therefore, although we modelled changes in statistical parameters (which are not estimated without 518 519 error) to identify seasonal periods, this step was only used to find the original climatic data to

be used in subsequent modelling. Inferences were not, therefore, directly based upon 520 521 statistical modelling of uncertain parameter estimates. We categorised the results of all 522 10,003 CSPs according to three broad response-types (CSP types I–III, Fig. 1), and retained P 523 values for each fitted model term to infer which of the modelled climatic effects were 524 statistically significant. We examined the evidence for trophic-level differences in the mean 525 seasonal timing of climate sensitivity by modelling the relationship between the start date, 526 end date and duration of the seasonal windows capturing the upper and lower limits of phenological sensitivity to temperature and rainfall as a function of trophic level (fixed 527 528 effect), with random effects of phenological metric, within species, within site. Analyses were conducted using the *base*, *mgcv* and *lme4* packages in \mathbb{R}^{38-40} . 529

530

531 Phase 2: "Global" models of phenological climate sensitivity

532 We estimated the upper and lower limits of phenological climate sensitivity at a multi-species 533 scale by "matching" each phenological series with data on mean temperature and precipitation, during the seasonal windows characterising the CSP for that series (Phase 1, 534 535 above). We aggregated all 10,003 of these matched phenology-climate data sets. To quantify 536 the average, multi-species, upper and lower limits of climate sensitivity we constructed a 537 linear mixed effects (LME) model, in which phenology (day of year) was modelled as a function of mean temperature and precipitation within the seasonal windows of the amassed 538 539 CSPs (fixed effects) with random effects of phenological metric, within species, within site. 540 These random effects were necessary since our data could not be considered independent. 541 The timing of events for the same species are more likely to be similar than for different 542 species. Likewise for different sites and the phenological metric-types used to describe the events (e.g. first flight time or seasonal peak abundance). Random slopes and intercepts were 543

allowed to ensure that each phenological event, for a species at a site, was allowed a differentrate of climate response.

546

For some species, more than one phenological event was recorded in the same year, at the 547 548 same site. For example, butterflies may have more than one flight period in the same year, 549 and plankton populations may be characterised by more than one seasonal abundance peak. 550 As climate responses are unlikely to be the same for the first event of the year, and 551 subsequent events, we introduced a voltinism factor in the analysis. This allowed us to distinguish between data representing the first/only events of each year (e.g. a spring 552 plankton bloom or butterfly generation) and second events in each year (e.g. the subsequent 553 554 summer plankton bloom or butterfly generation). This distinction captured all possibilities 555 within our data set.

556

For site *i*, species *j*, voltmetric *k* (where voltmetric is a unique combination of voltinism class and the metric-type used to identify the event) the corresponding day of year (*DOY*) of a particular seasonal event is modelled as:

560

561
$$DOY_{ijk} = \alpha_0 + \beta_1 T L_{ijk} + \beta_2 T U_{ijk} + \beta_3 P L_{ijk} + \beta_4 P U_{ijk} + \varepsilon_{ijk}$$

where $\varepsilon_{ijk} \sim N(0, \sigma^2)$ and the model includes temperature at the upper limit of each CSP (*TU*), temperature at the lower limit of each CSP (*TL*), precipitation at the upper limit of each CSP (*PU*) and precipitation at the lower limit of each CSP (*PL*). Due to the non-independence within the data, we allow the intercepts and coefficients corresponding to all four covariates to vary by site, species and voltmetric. Preserving the natural nesting of a metric for a species

567 at a particular site, this gives:

568

569
$$\alpha_0 = \gamma_0 + \mu_{0;ijk} + \mu_{0;ij,k} + \mu_{0;i,jk}$$

570 $\beta_l = \gamma_l + \mu_{l;\,ijk} + \mu_{l;\,ij,k} + \mu_{l;\,ij,k}$

571 $\beta_2 = \gamma_2 + \mu_{2; ijk} + \mu_{2;ij,k} + \mu_{2;i,jk}$

572 $\beta_3 = \gamma_3 + \mu_{3; ijk} + \mu_{3; ij,k} + \mu_{3; i,jk}$

573 $\beta_4 = \gamma_4 + \mu_{4; ijk} + \mu_{4;ij,k} + \mu_{4;ij,k}$

574

575 where each of the μ terms is a random effect following: $\mu \sim N(0,\delta^2)$

576

This nesting of random effects is most conservative in terms of inference at the global level and is as flexible as possible, allowing each time series to have its own set of model parameters. This permits a high degree of biological realism since each distinct phenological event, for a given species, at a given site, is permitted to have a different slope for the effects of temperature and precipitation i.e. a different climate sensitivity.

582

In this model framework we are specifically testing the null hypotheses that each of the climate variables show no relation with seasonal timing of biological events. Because of this, and the fact that each parameter is estimated directly, without distributional form assumed a*priori* or as the target distribution, we follow a frequentist approach to analysis. However, 587 because the exact degrees of freedom cannot be evaluated when using restricted maximum 588 likelihood, hence no exact P-value, we present full summaries of all the parameters estimated at species level (as given by: $\gamma + \mu_{ij,k} + \mu_{i,jk}$, above). Approximate P-values could be 589 590 presented based on taking conservative estimates of the degrees of freedom though, given the 591 volume of data available, this will typically lead to the detection of many statisticallysignificant results that may not be biologically significant. Examining the full range of 592 593 estimated coefficients across the random effects levels ensures that we present the full range 594 of variation around global parameters and can make more informed inference. In this way we 595 encourage the reader to interpret our results by using biological insight, not by depending 596 upon P-values alone.

597

598 To examine high-level differences in climate sensitivity among trophic levels and taxonomic 599 groups we re-fitted the LME model with these attributes as fixed-effect factors, interacting 600 with the fixed-effect climate variables. The fixed-effect slopes from the resulting models 601 allowed us to compare differences in phenological climate sensitivity among these broad 602 organism groups, averaged across all taxa within each group. Supplementary Table 2 shows 603 the number of phenological series, sites and distinct taxa that contributed data to each of these 604 groups. All models were run twice: once to examine correlations with both temperature and 605 precipitation excluding marine plankton data (9,800 time series), and once to examine only 606 temperature-phenology correlations for the whole data set (10,003 time series).

607

608 **Potential biases**

Data availability differed among taxonomic groups. To assess the extent to which mean responses were biased by data inequality we conducted Monte Carlo re-sampling, iteratively selecting 5, 20, 50 and 100 phenological series from each taxonomic group and re-fitting climate-phenology models with these sampled data sets. For taxonomic groups with less data than the larger sample sizes, we retained all available data (Supplementary Discussion). This allowed us to compare taxonomic group and trophic level responses based upon sampled and all data, to fully investigate potential bias.

616

Another potential bias in our analysis is that phenological time series length is variable, affecting the length of time over which climate-phenology correlations are assessed. In order to assess the extent to which differences in mean trophic level and taxonomic group responses are biased by variable time series length, we also re-fitted our models but based only on pre- and post-1980 data. All models were run in the *lme4* package in $R^{38,40}$.

622

623 Estimating future change

624 To estimate potential future "net" effects of temperature and precipitation change, we 625 compared predictions of seasonal timing under baseline conditions, and under established 626 climate change scenarios. Firstly, estimates of seasonal timing (day of year) were obtained 627 for the same baseline period used in the UKCP09 projections (long term average 1961-1990), 628 using modelled correlations between phenology, temperature and precipitation (from Phase 629 1). Having obtained these baseline estimates, we applied our models to projected changes in 630 monthly temperature and precipitation for the 2050s (UK Climate Projections, UKCP09, http://ukclimateprojections.metoffice.gov.uk/). We used 10th, 50th and 90th percentile changes 631

under low, medium and high emissions scenarios (relative to the 1961-90 baseline). The spatial location of each phenological series was matched to climate projection data for the 25 \times 25km grid square in which it occurred, and temporally matched to climatic data from the months-of-year in which its respective climate sensitivity windows occurred. Relative changes in timing, in response to climatic change of the magnitude projected to occur by the 2050s, were summarised by trophic levels and taxonomic groups.

638

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664 FIGURE LEGENDS

Figure 1 | Climate Sensitivity Profiles (CSPs). Climate sensitivity is the change in seasonal 665 timing per unit change in temperature (days $^{\circ}C^{-1}$) or precipitation (days mm⁻¹). Irrespective of 666 667 the date, increasing temperature/precipitation may always correlate with earlier (red curve, 668 CSP type I) or later (blue curve, CSP type III), biological events, but sensitivity to climate 669 variation (correlation magnitude) differs (cf w1 and w2, w5 and w6). In contrast, opposing 670 climate-phenology correlations may occur, depending on the date at which climate changes 671 (orange curve, w3 and w4, CSP type II). Panels show hypothetical relationships for seasonal 672 windows w1-w6.

673

Figure 2 | **Climatic change in the UK, 1960-2012.** a) Long-term changes in air temperature and precipitation are the differences between 1960 and 2012 monthly means of these variables, derived from a regression fitted through each monthly time series. Error bars indicate the standard deviation of linearly-detrended climatological data, as an indication of inter-annual variation around each trend. b) and c) Examples of spatial variation in the extent of long-term climatic changes are shown for March air temperatures and February precipitation.

681

Figure 3 | **Upper and lower limits of phenological climate sensitivity**. Sensitivity is the slope of the relationship between seasonal timing (day of year) and climatic variables. Alltaxa upper and lower limits in a) temperature (°C) and b) precipitation (mm day⁻¹) sensitivity are summarised. Lower (c, d) and upper (e, f) limits of temperature (c, e) and precipitation (d, f) sensitivity are shown by trophic level. Inverted triangles indicate average sensitivity. 687 Curves are kernel density plots: estimates of the probability density distribution of species-688 level climate sensitivity i.e. the relative likelihood of different levels of climate sensitivity 689 within each species group (n = 370,725).

690

Figure 4 | Upper and lower limits of phenological climate sensitivity for broad taxonomic groups. Lower (blue) and upper (red) limits of the sensitivity of phenological events to seasonal temperature (a) and precipitation (b) change are shown. Coloured circles indicate the median response, and bars show the 5th-to-95th percentile responses for each group. Sensitivity is quantified by summarising the species-level (random effects) responses from a mixed effects model including data for all taxa, and with taxonomic group as a fixed effect (n = 370,725).

698

699 Figure 5 | Estimated phenological shifts by the 2050s. Modelled responses to projected 700 temperature and precipitation change, assuming contemporary climate sensitivity, for trophic 701 levels (a) and taxonomic groups (b). Projected median shifts in seasonal timing are shown. 702 Change estimates are based on low, medium and high emissions climate scenarios. Bars represent median responses to 50th percentile climate change projections under each scenario, 703 while extremes of whiskers represent median responses to 10th and 90th percentile projected 704 705 climatic changes under each scenario. Standard deviations indicate variation in projected responses for each group under the 50th percentile of the medium emissions scenario. 706

707

708 EXTENDED DATA FIGURE AND TABLE LEGENDS

Extended Data Figure 1 | Limits of phenological temperature sensitivity inclusive of marine plankton data. Upper and lower limits of phenological temperature sensitivity are quantified as the slope of the relationship between seasonal timing (day of year) and temperature (°C) variation within specific seasonal periods. Limits in temperature sensitivity are shown for all taxa (a) and by trophic level (lower limit b, upper limit c). Inverted triangles indicate average sensitivity for all species in each group and curves are probability density plots of species-level variation in sensitivity.

718

Extended Data Figure 2 | Limits of phenological climate sensitivity for taxonomic groups (top) and trophic levels (bottom), after Monte-Carlo resampling. Lower (blue) and upper (red) limits of the sensitivity of phenological events to seasonal temperature (a) and precipitation (b) change. Coloured circles: responses based upon the full data set. Bars: 2.5th-to-97.5th percentile responses for each group, based upon 100 draws from the full data set. Data were sampled so that 5, (dotted bar), 20 (solid bar), 50 (dashed bar) and 100 (dotdashed bar) phenological time series were drawn from each taxonomic group.

726

Extended Data Figure 3 | Climate sensitivities, based on different time periods (top: all data, middle: pre-1980 data, bottom: post-1980 data). Sensitivity is the slope of the relationship between seasonal timing (day of year) and temperature (°C), or precipitation (mm day⁻¹). Limits of a) temperature and b) precipitation sensitivity are summarised for all taxa. Lower (c, d) and upper (e, f) limits of temperature (c, e) and precipitation (d, f) sensitivity are shown by trophic level. Inverted triangles: average sensitivity for all species (a, b) or trophic levels (c-f). Curves: kernel density plots: probability density distributions of
species-level climate sensitivity i.e. the relative likelihood of different climate sensitivities
within each species group.

736

Extended Data Figure 4 | Limits of phenological climate sensitivity for broad taxonomic groups (top: all data, bottom: post-1980 data only). Lower (blue) and upper (red) limits of the sensitivity of phenological events to seasonal temperature (a) and precipitation (b) change are shown. Coloured circles indicate the median response, and bars show the 5th-to-95th percentile responses for each group. Sensitivity is quantified by summarising the specieslevel (random effects) responses from a mixed effects model including data for all taxa, and with taxonomic group as a fixed effect.

744

745 Extended Data Figure 5 | Seasonal windows for Climate Sensitivity Profiles (CSPs). 746 Estimated climatic sensitivity at the lower (a, c) and upper (b, d) limits of CSPs for 10,003 phenological series. Grey lines are seasonal time periods (x axis) within which climatic 747 748 variables have their most positive/negative correlations with the seasonal timing of each phenological event. The y-axis indicates the slope coefficient for each of these correlations; a 749 measure of climate sensitivity (days change $^{\circ}C^{-1}$, or mm⁻¹). Shown are the lower/upper limits 750 of CSP_{temp} (a, b, respectively) and the lower/upper limits of CSP_{precip} (c, d, respectively). Inset 751 752 histograms show seasonal time window length (days).

753

Extended Data Figure 6 | Time lags between phenological events and seasonal windows
 of climate sensitivity. Frequency histograms showing the time lag (in days) between the

mean timing of each phenological event and the end of seasonal windows corresponding to the lower and upper limits of CSP_{temp} (a, b, respectively) and the lower and upper limits of CSP_{precip} (c, d, respectively). Peaks at lags of around 1 year are where windows were identified that ended at the mean seasonal timing of an event, but in the previous year, due to temporal autocorrelation in climate data.

761

Extended Data Figure 7 | Seasonal windows for Climate Sensitivity Profiles (CSPs) by trophic level. Estimated climatic sensitivity at the lower and upper limits of CSPs for taxa at each of three trophic levels. Formatting is the same as in Extended Data Figure 5. Shown are the lower and upper limits of CSP_{temp} (a, b, respectively) and the lower and upper limits of CSP_{precip} (c, d, respectively).

767

768 Extended Data Figure 8 | Example Climate Sensitivity Profile (CSP). Temperature sensitivity (CSP_{temp}) for alderfly (Sialis lutaria) emergence from Windermere, UK. Solid 769 770 black line: sensitivity of first emergence to water temperature on different days of the year (days change °C⁻¹). Grey horizontal lines: 2.5 and 97.5 percentiles of these sensitivity values. 771 Solid orange curve: GAM smoother fitted through the sensitivity values with associated 772 confidence intervals (dashed orange curves). Horizontal bars indicate where GAM 773 confidence intervals exceed the percentiles of the original sensitivity values, indicating 774 775 seasonal windows at the limits of the climate sensitivity profile.

Extended Data Table 1 | Modelled relationships between seasonal timing and climate
 variables for n=10,003 phenological time series. Climate Sensitivity Profiles (CSPs) fall

779 within three broad response-types; events always advance with increases in the climate 780 variable irrespective of the seasonal timing of climate change (CSP Type I, Fig. 1 - red curve), events are always delayed by increases in the climate variable irrespective of the 781 782 seasonal timing of climate change (CSP Type III, Fig. 1 - blue curve), and events may be 783 advanced or delayed by increases in the climate variable, depending on the seasonal timing of climate change (CSP Type II, Fig. 1 - orange curve). Shown are the percentage of series that 784 785 fall in each Type (% series), the percentage of effects that are statistically significant at P<0.05 after multiple testing correction (% effects significant). □ Based only on freshwater 786 787 and terrestrial taxa, for which precipitation data were available. † NA indicates effect not 788 evaluated, due to lack of precipitation data for marine taxa

789

790 Extended Data Table 2 | Parameter estimates and test statistics from climate-phenology 791 **mixed-effects models.** Presented are fixed-effect parameter estimates from each model; the 792 intercept and slope for each climatic predictor. Following R convention, absolute parameter 793 estimates are provided for an assigned "baseline" group within each model (b), and remaining estimates are given as differences from this baseline (Δb). Each estimate has an associated 794 795 standard error and t statistic in parentheses (standard error, t). Climatic predictors include 796 mean temperature and precipitation in seasonal windows at the upper and lower limit of the 797 climate sensitivity profile for each phenological series. The number of observations, n_{i} is 798 370,725.
Models were re-run including the marine plankton data, and excluding precipitation effects (see text). In these models the number of observations, n = 379,081799









