

$\mathcal{K}$	set of all locations
$\mathcal{L}$	set of all pixels
$\mathcal{Z}$	set of all possible depth values
$X_k$	binary occupancy variable for a location $k$
$Z_i$	observed depth value variable at pixel $i$
$z^\infty$	special value for when no depth is observed
$M_{ki}$	segmentation mask at pixel $i$ for location $k$
$\mathcal{S}_k$	object silhouette for $k$ -th location
$ \mathcal{S}_k $	number of pixels in the $k$ -th silhouette
$\langle \cdot \rangle_p$	expectation wrt a distribution $p$
$\sigma(x)$	sigmoid function $(1 + e^{-x})^{-1}$
$\rho_k$	approximate posterior $Q(X_k = 1)$
$\pi^\infty$	probability of observing $z^\infty$
$\pi^\circ$	probability of observing an outlier
$\Delta_{l,i}$	$\log \theta_{l,i}(z)$ , $l \in \mathcal{K} \cup \{bg\}$
$\tau_{ki}$	$\prod_{l < k, i \in \mathcal{M}_l} (1 - \rho_l)$

Table 1: Notations

We know that  $X_k$  is a Bernoulli variable ( $Q(X_k = 1) + Q(X_k = 0) = 1$ ), and keeping in mind Eq. 7 from the paper:

$$Q(X_k = 1) = \frac{\exp(\langle \log P(\mathbf{Z}, \mathbf{X}, \mathbf{M}) | X_k = 1 \rangle)}{\tilde{Z}_{X_K}}$$

$$Q(X_k = 0) = \frac{\exp(\langle \log P(\mathbf{Z}, \mathbf{X}, \mathbf{H}) | X_k = 0 \rangle)}{\tilde{Z}_{X_K}}$$

which allows us to find the normalizing factor  $\tilde{Z}_{X_K}$  and get the following update on  $\rho_k = Q(X_k = 1)$  (Eq. 8 in the paper):

$$\rho_k = \sigma( \langle \log P(\mathbf{Z}, \mathbf{X}, \mathbf{M}) | X_k = 1 \rangle_{Q(\mathbf{x}/X_k)} - \langle \log P(\mathbf{Z}, \mathbf{X}, \mathbf{M}) | X_k = 0 \rangle_{Q(\mathbf{x}/X_k)} )$$

Let's take a closer look at the conditional expectation of the joint  $\langle \log P(\mathbf{Z}, \mathbf{X}, \mathbf{M}) | X_k = \xi \rangle_{Q(\mathbf{x}/X_k)}$ ,  $\xi \in \{0, 1\}$  (omitting  $\mathbf{M}$  for clarity):

$$\begin{aligned}
& \langle \log P(\mathbf{Z}, \mathbf{X} | X_k = \xi) \rangle_{Q(\mathbf{X}/X_k)} = \\
& \langle \log P(\mathbf{Z} | \mathbf{X}) P(\mathbf{X}) | X_k = \xi \rangle_{Q(\mathbf{X}/X_k)} = \\
& \sum_{i \in \mathcal{S}_k} \langle \log P(Z_i | \mathbf{X}) | X_k = \xi \rangle + \sum_{l \in \mathcal{K}} \langle \log P(X_l) | X_k = \xi \rangle
\end{aligned}$$

where we used an assumption of conditional independence between pixels, and also assumption of independence of occupancies a-priori.

We now need to evaluate the expectation of the observation likelihood  $\log P(Z_i | \mathbf{X})$  conditioned on  $X_k = \xi$  for  $\xi \in \{0, 1\}$ . Under our generative model, each pixel is either generated by one of the silhouettes, or by the background. If it was generated by some silhouette  $\mathcal{S}_l$ , then, first, all the silhouettes that are closer to the camera should be absent (which happens with probability  $\tau_{l-1,i}$ ), and, second, the silhouette itself should be present (which happens independently with probability  $\rho_l$ ). Otherwise, all the silhouettes are absent (probability  $\tau_{\mathcal{K}|,i}$ ), and pixel was generated by the background. Now, let's write down the expected log-likelihood for observing some value  $z$  at pixel  $i \in \mathcal{L}$  (omitting segmentation masks for clarity):

$$\langle \log P(Z_i = z | \mathbf{X}) \rangle_{Q(\mathbf{X})} = \sum_{l \in \mathcal{K}} \tau_{l-1,i} \rho_l \log \theta_{li}(z) + \tau_{\mathcal{K}|,i} \log \theta_{bg,i}(z)$$

When conditioned on  $X_k = 1$  the expectation will be:

$$\sum_{l < k} \tau_{l-1,i} \rho_l \log \theta_{li}(z) + \tau_{k-1,i} \log \theta_{ki}(z)$$

And conditioned when on  $X_k = 0$ :

$$\sum_{l < k} \tau_{l-1,i} \rho_l \log \theta_{li}(z) + \frac{\sum_{l > k} \tau_{l-1,i} \rho_l \log \theta_{li}(z) + \tau_{\mathcal{K}|,i} \log \theta_{bg,i}(z)}{(1 - \rho_k)}$$

Now, one can evaluate  $\langle \log P(\mathbf{Z}, \mathbf{X} | X_k = \xi) \rangle_{Q(\mathbf{X}/X_k)}$  (expectations for the priors is trivial, see e.g. Fleuret'09). If one substitutes these into Eq. 8, one will get Eq. 10:

$$\begin{aligned}
\rho_k &= \sigma\left(\log \frac{\epsilon}{1-\epsilon} + \right. \\
& \sum_{i \in \mathbf{M}_k} \tau_{k-1,i} \log \theta_{k,i} - \\
& \left. \sum_{i \in \mathbf{M}_k} \frac{1}{1-\rho_k} \left( \sum_{l > k, i \in \mathbf{M}_l} \tau_{l-1,i} \rho_l \log \theta_{l,i} + \tau_{\mathcal{K}|,i} \log \theta_{bg,i} \right) \right)
\end{aligned}$$