Highly Efficient Structural Learning of Sparse Staged Trees

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Abstract

Several structural learning algorithms for staged tree models, an asymmetric extension of Bayesian networks, have been defined. However, they do not scale efficiently as the number of variables considered increases. Here we introduce the first scalable structural learning algorithm for staged trees, which searches over a space of models where only a small number of dependencies can be imposed. A simulation study as well as a real-world application illustrate our routines and the practical use of such data-learned staged trees.

Keywords: Asymmetric conditional independence; Bayesian networks; Probabilistic graphical models; Staged trees; Structural learning.

1. Introduction

Probabilistic graphical models, and in particular Bayesian networks (BNs), are nowadays widely used in machine learning to conveniently represent the relationships existing between the components of a random vector. The directed acyclic graph (DAG) associated to a BN represents graphically (symmetric) conditional independence statements, which can be assessed using the d-separation criterion (Pearl, 1988). Although the underlying DAG can be expert-elicited, this is often learned from data using algorithms that explore the space of all possible DAGs (see e.g. Scutari et al., 2019).

For quite some time it has been noticed that the strict assumption of symmetric conditional independence may be too restrictive to fully represent the relationship between variables in a dataset (Boutilier et al., 1996; Chickering et al., 1997; Friedman and Goldszmidt, 1996). However, the development and use in practice of probabilistic graphical models embedding asymmetric conditional independence has been limited (see Hyttinen et al., 2018; Nicolussi and Cazzaro, 2021; Talvitie et al., 2019, for some recent proposals). Possible reasons behind the limited use of such models could be: (i) the lack of widely available software; (ii) the complexity of the learning routines; and (iii) the less intuitive visualization of the associated independences which are not explicitly represented by a single graph.

Staged trees (Collazo et al., 2018; Smith and Anderson, 2008) are probabilistic graphical models which, starting from an event tree, represent any type of asymmetric conditional independence by a partitioning/coloring of the vertices of the tree. All the model information and the conditional independences can be read directly from the tree and, in particular, from the coloring of the vertices. The stagedtrees R package (Carli et al., 2022) provides a user-friendly implementation of a wide array of structural learning and inferential routines to fit staged trees to data. Therefore, two of the main limitations to the use of asymmetric

probabilistic graphical models do not apply to staged trees. Furthermore, a wide toolkit of procedures to work with staged trees have been developed, including handling missing data, sensitivity analysis and exploration of equivalence classes, among others (see Collazo et al., 2018, for details).

On the other hand, learning staged trees from data is complex. Although efficient structural learning algorithms for such models have been implemented (e.g. Freeman and Smith, 2011; Leonelli and Varando, 2021; Silander and Leong, 2013), they can only work with a limited number of variables. The main reason behind this is the explosion of the size of the model search space as the number of variables considered increases. As an illustration, the number of DAGs over 6 binary variables is 3781503, whilst there are 1.20019×10^{44} staged trees under the same conditions (Duarte and Solus, 2021). Furthermore, the number of DAGs remains constant if variables have more than two levels, whilst the number of staged trees would further increase dramatically.

Recent proposals for efficient structural learning of staged trees look at sub-classes of staged trees, with the aim of reducing the size of the model space. Carli et al. (2020) defined naive staged trees which have the same number of parameters of a naive BN over the same variables. Leonelli and Varando (2022) considered simple staged trees which have a constrained type of partitioning of the vertices. Duarte and Solus (2021) defined CStrees which only embed symmetric and context-specific types of independence, and not others (Pensar et al., 2016).

One of the first solutions to make structural learning of BNs scalable was to limit the number of parents each variable can have (Friedman et al., 1999; Tsamardinos et al., 2006). This was imposed not only to restrict the model space of possible DAGs, but it also made sense from an applied point of view since most often only a limited number of variables can be expected to have a direct influence on another. The option of setting a maximum number of parents is also available in the standard bnlearn software (Scutari, 2010).

Here, we define a sub-class of staged trees embedding the same idea of limiting the number of variables that can have a direct influence to another. As we formalize below, this means that the BN representation associated to such a staged tree is sparse, meaning that it has a small number of edges. A structural learning algorithm for this class of staged trees is introduced and its features are explored in an extensive simulation study.

2. Bayesian Networks and Staged Trees

Before introducing staged trees, we give a formal definition of BNs. We then describe their relationships with staged tree models.

2.1 Bayesian Networks

Let $G = ([p], E_G)$ be a DAG with vertex set $[p] = \{1, \ldots, p\}$ and edge set E_G . Let $\mathbf{X} = (X_i)_{i \in [p]}$ be categorical random variables with joint mass function P and sample space $\mathbb{X} = \times_{i \in [p]} \mathbb{X}_i$. For $A \subset [p]$, we let $\mathbf{X}_A = (X_i)_{i \in A}$ and $\mathbf{x}_A = (x_i)_{i \in A}$ where $\mathbf{x}_A \in \mathbb{X}_A = \times_{i \in A} \mathbb{X}_i$. We say that P is Markov to G if, for $\mathbf{x} \in \mathbb{X}$,

$$P(\boldsymbol{x}) = \prod_{k \in [p]} P(x_k \mid \boldsymbol{x}_{\Pi_k}),$$

where Π_k is the parent set of k in G and $P(x_k|\mathbf{x}_{\Pi_k})$ is a shorthand for $P(X_k = x_k|\mathbf{X}_{\Pi_k} = \mathbf{x}_{\Pi_k})$. It is customary to label the vertices of a BN so to respect the topological order of G and we henceforth assume that $1, 2, \ldots, p$ is a topological order of G.

The ordered Markov condition implies conditional independences of the form

$$X_i \perp \mathbf{X}_{[i-1]} \mid \mathbf{X}_{\Pi_i}$$
.

Henceforth, P is assumed to be strictly positive. Let G be a DAG and P Markov to G. The Bayesian network model (associated to G) is

$$\mathcal{M}_G = \{ P \in \Delta^{\circ}_{|\mathbb{X}|-1} \mid P \text{ is Markov to } G \}.$$

where $\Delta_{|\mathbb{X}|-1}^{\circ}$ is the $(|\mathbb{X}|-1)$ -dimensional open probability simplex.

2.2 Staged Trees

Differently to BNs, whose graphical representation is a DAG, staged trees visualize conditional independence by means of a colored tree. Let (V, E) be a directed, finite, rooted tree with vertex set V, root node v_0 and edge set E. For each $v \in V$, let $E(v) = \{(v, w) \in E\}$ be the set of edges emanating from v and C be a set of labels.

An **X**-compatible staged tree is a triple $T = (V, E, \theta)$, where (V, E) is a rooted directed tree and:

- 1. $V = v_0 \cup \bigcup_{i \in [p]} X_{[i]};$
- 2. For all $v, w \in V$, $(v, w) \in E$ if and only if $w = \boldsymbol{x}_{[i]} \in \mathbb{X}_{[i]}$ and $v = \boldsymbol{x}_{[i-1]}$, or $v = v_0$ and $w = x_1$ for some $x_1 \in \mathbb{X}_1$;
- 3. $\theta: E \to \mathcal{L} = \mathcal{C} \times \bigcup_{i \in [p]} \mathbb{X}_i$ is a labelling of the edges such that $\theta(v, \boldsymbol{x}_{[i]}) = (\kappa(v), x_i)$ for some function $\kappa: V \to \mathcal{C}$. The function k is called the colouring of the staged tree T.

If $\theta(E(v)) = \theta(E(w))$ then v and w are said to be in the same stage. Therefore, the equivalence classes induced by $\theta(E(v))$ form a partition of the internal vertices of the tree in stages.

Points 1 and 2 above construct a rooted tree where each root-to-leaf path, or equivalently each leaf, is associated to an element of the sample space \mathbb{X} . Then a labeling of the edges of such a tree is defined where labels are pairs with one element from a set \mathcal{C} and the other from the sample space \mathbb{X}_i of the corresponding variable X_i in the tree. By construction, **X**-compatible staged trees are such that two vertices can be in the same stage if and only if they correspond to the same sample space.

Figure 1 reports an (X_1, X_2, X_3, X_4) -compatible staged tree over four binary variables. The *coloring* given by the function κ is shown in the vertices and each edge $(\cdot, (x_1, \ldots, x_i))$ is labeled with $X_i = x_i$. The edge labeling θ can be read from the graph combining the text label and the color of the emanating vertex. The staging of the staged tree in Figure 1 is given by the partition $\{v_0\}$, $\{v_1\}$, $\{v_2\}$, $\{v_3, v_4\}$, $\{v_5, v_6\}$, $\{v_7, v_{11}\}$, $\{v_8, v_{12}\}$, $\{v_9, v_{13}\}$ and $\{v_{10}, v_{14}\}$.

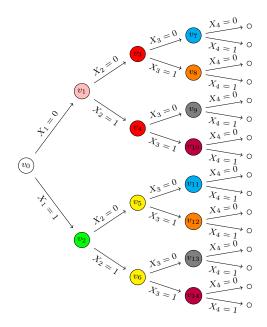


Figure 1: Example of an X-compatible staged tree over four binary random variables.

The parameter space associated to an **X**-compatible staged tree $T = (V, E, \theta)$ with labeling $\theta : E \to \mathcal{L}$ is defined as

$$\Theta_T = \left\{ \boldsymbol{y} \in \mathbb{R}^{|\theta(E)|} \mid \forall \ e \in E, y_{\theta(e)} \in (0, 1) \text{ and } \sum_{e \in E(v)} y_{\theta(e)} = 1 \right\}.$$
 (1)

Equation (1) defines a class of probability mass functions over the edges emanating from any internal vertex coinciding with conditional distributions $P(x_i|\mathbf{x}_{[i-1]})$, $\mathbf{x} \in \mathbb{X}$ and $i \in [p]$. In the staged tree in Figure 1 the staging $\{v_3, v_4\}$ implies that the conditional distribution of X_3 given $X_1 = 0$ and $X_2 = 0$, represented by the edges emanating from v_3 , is equal to the conditional distribution of X_3 given $X_1 = 0$ and $X_2 = 1$. A similar interpretation holds for the staging $\{v_5, v_6\}$. This in turn implies that $X_3 \perp X_2 \mid X_1$, thus illustrating that the staging of a tree is associated to conditional independence statements.

Let l_T denote the leaves of a staged tree T. Given a vertex $v \in V$, there is a unique path in T from the root v_0 to v, denoted as $\lambda(v)$. The depth of a vertex $v \in V$ equals the number of edges in $\lambda(v)$. For any path λ in T, let $E(\lambda) = \{e \in E : e \in \lambda\}$ denote the set of edges in the path λ .

The staged tree model \mathcal{M}_T associated to the **X**-compatible staged tree (V, E, θ) is the image of the map

$$\phi_T : \Theta_T \to \Delta_{|l_T|-1}^{\circ}
 \mathbf{y} \mapsto \left(\prod_{e \in E(\lambda(l))} y_{\theta(e)} \right)_{l \in \mathbf{l}_T}$$
(2)

Therefore, staged trees models are such that atomic probabilities are equal to the product of the edge labels in root-to-leaf paths and coincide with the usual factorization of mass functions via recursive conditioning.

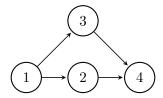


Figure 2: The DAG G such that T_G is the staged tree in Figure 1.

2.3 Staged Trees and Bayesian Networks

Although the relationship between BNs and staged trees was already formalized by Smith and Anderson (2008), a formal procedure to represent a BN as a staged tree has been only recently introduced in Duarte and Solus (2020) and Varando et al. (2021).

Assume X is topologically ordered with respect to a DAG G and consider an X-compatible staged tree with vertex set V, edge set E and labeling θ defined via the coloring $\kappa(\boldsymbol{x}_{[i]}) = \boldsymbol{x}_{\Pi_i}$ of the vertices. The staged tree T_G , with vertex set V, edge set E and labeling θ so constructed, is called the staged tree model of G. Importantly, $\mathcal{M}_G = \mathcal{M}_{T_G}$, i.e. the two models are exactly the same, since they entail exactly the same factorization of the joint probability (Smith and Anderson, 2008). Clearly, the staging of T_G represents the Markov conditions associated to the graph G.

Varando et al. (2021) approached the reverse problem of transforming a staged tree into a BN. Of course, since staged trees represent more general asymmetric conditional independences, given a staged tree T most often there is no BN with DAG G_T such that $\mathcal{M}_T = \mathcal{M}_{G_T}$. However, Varando et al. (2021) introduced an algorithm that, given an X-compatible staged tree T, finds the minimal DAG G_T such that $\mathcal{M}_T \subseteq \mathcal{M}_{G_T}$. Minimal means that such a DAG G_T embeds all symmetric conditional independences that are in \mathcal{M}_T and that there are no DAGs with less edges than G_T embedding the same conditional independences.

As an illustration, the staged tree in Figure 1 can be constructed as the T_G from the BN with DAG in Figure 2, embedding the conditional independences $X_3 \perp \!\!\! \perp X_2 \mid X_1$ and $X_4 \perp \!\!\! \perp X_1 \mid X_2, X_3$. Conversely, consider the staged tree T in Figure 3, which differs from the one in Figure 1 only on a different coloring of the vertices v_6 and v_{14} . Such a staged tree does not embed any symmetric conditional independence, only non-symmetric ones, and therefore there is no DAG G_T such that $\mathcal{M}_{G_T} = \mathcal{M}_T$. Furthermore, the minimal DAG G_T such that $\mathcal{M}_T \subseteq \mathcal{M}_{G_T}$ is the complete one since the staging of the tree implies direct dependence between every pair of variables.

2.4 Structural Learning Algorithms Inducing Sparsity

Since the model search space of staged trees is huge, we consider here structural learning for a subclass of staged trees that we define next.

Definition 1 A staged tree T is in the class of k-parents staged trees if the maximum in-degree in G_T is less or equal to k.

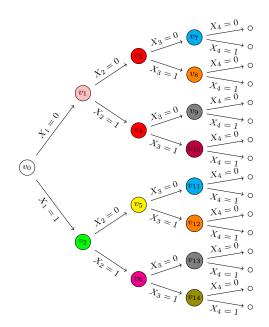


Figure 3: Example of an X-compatible staged tree over four binary random variables.

For instance, the staged tree in Figure 1 is in the class of 2-parents staged trees, whilst the one in Figure 3 is not, since its associated minimal DAG is such that X_4 has three parents.

Of course, the class of k-parents staged trees is much smaller than the one of X-compatible ones, for small values of k, and therefore structural learning is expected to be quicker. Here we define a structural learning algorithm to learn a staged tree in the class of k-parents staged trees which consists of the following steps: (i) learn a BN with DAG G having at most k parents (for instance using bnlearn); (ii) construct the equivalent staged tree T_G ; (iii) run the backward hill-climbing algorithm of Carli et al. (2022) which only joins stages together (no splitting of stages) based on the minimization of the model BIC (Görgen et al., 2022). Call the resulting staged tree T. It can be easily proven that G_T has at most k parents and $\mathcal{M}_{G_T} \subseteq \mathcal{M}_G$.

Although the idea of using the staged tree equivalent to a BN as starting point of a structural learning algorithm (or at least using a partial ordering associated to such a BN) is not new (see e.g. Barclay et al., 2013), here we specifically use such a strategy to limit the complexity of the learned staged tree. This has two major advantages: (i) the speed of the algorithms increases greatly; (ii) the non-symmetric conditional independences can be easily visualized even when a large number of variables are present, as illustrated in Section 4.

3. Experiments

We perform simulation experiments to evaluate the proposed learning strategy for k-parents staged trees. Moreover we it compare to standard learning of staged trees and DAGs.

In all the simulated experiments we generated data from random k-parents staged trees, which are obtained as follow: (1) A random DAG G with fixed topological order X_1, \ldots, X_p

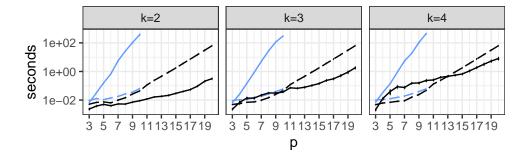


Figure 4: Average computation times for learning the stagedtrees (solid lines) with the standard BHC approach (blue) and the proposed k-parents BHC method (black). Additionally the time spent creating the initial model are shown (dashed lines). Results for simulated data from $p=3,\ldots,20$ binary variables, BHC results are obtained only up to 10 variables. In different columns results for the maximum number of parents 2,3,4.

is obtained by randomly selecting up to k parents uniformly from $\{X_1, \ldots, X_{j-1}\}$ for each X_j ; (2) The equivalent staged tree T_G is obtained; (3) Stages in T_G are randomly merged with probability 0.5. The obtained staged tree T is such that G_T is a sub-graph of G and thus it is a k-parents staged tree; (4) Lastly, we assign random probabilities (uniformly from the simplex) to each stage of the staged tree T. Once we generate the random k-parents staged tree T, we can easily sample observations of X_1, \ldots, X_p from it via sequential sampling (as implemented in the stagedtrees package). For each fixed parameters (k, p) and number of observation sampled) we repeat the experiments 20 times and report averages and standard errors.

3.1 Oracle DAG

We first consider an ideal scenario where we evaluate the performance of the proposed method when the starting DAG is the graph G used to generate the true staged tree T. We thus run a standard backward hill-climbing (BHC) procedure, as implemented in the stagedtrees package, starting from both the full saturated tree model and starting from the model T_G . All heuristic searches optimize the BIC score.

We plot in Figure 4 the average computation time as a function of the system size $p=3,\ldots,20$ and varying the number of maximum parents k=2,3,4. For each one of the two approaches we split the computation time in: build time, the time spent building the starting tree; and search time, the time spent running the search algorithm. We can observe that starting the search from the DAG-equivalent tree allows us to scale the algorithm easily up to 20 variables, while the standard approach starting from the full tree become quickly infeasible after 10 variables.

As a sanity check we also compute the normalized hamming distance (the sum across the depth of the tree of the average number of nodes for which coloring needs to change to obtain the same staged tree) and the context intervention distance (Leonelli and Varando,

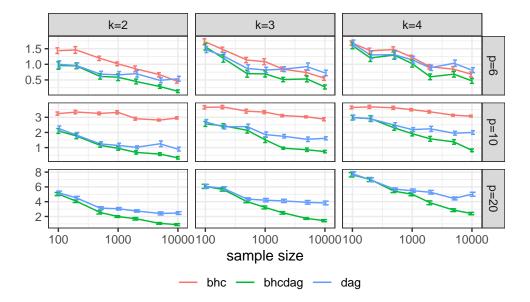


Figure 5: Average normalized hamming distance between true and estimated staged trees as a function of sample size, for different methods, system size (p, rows) and maximum parents (k, columns).

2021) between the true and learned models. As expected the k-parents trees obtain better results, which we do not report here for lack of space.

3.2 Learned DAG

We perform now a simulation study similar to the oracle setting but, as in a more realistic scenario, we do not assume knowledge of the DAG G. We thus, first estimate a DAG \hat{G} from data using the hill-climbing approach in the bnlearn package (Scutari, 2010), and then we apply the BHC learning algorithm starting from the staged tree $T_{\hat{G}}$ (bhcdag). We compare the obtained tree to $T_{\hat{G}}$ (the tree equivalent to the learned DAG, dag) and the output of the BHC algorithm starting from the full saturated model (bhc). We run 20 replications of the experiment for different system sizes (p=6,10,20), for different number of maximum parents in the true DAG (k=2,3,4) and with sample sizes ranging from 100 to 10000.

In Figure 5 we plot the average (across repetitions) normalized hamming distance between the estimated tree and the true one. We can observe that the k-parents staged trees, obtained by the BHC algorithm starting from $T_{\hat{G}}$ (bhcdag), are closer to the true data-generating models, with respect to both the output of BHC starting from the saturated model and the tree $T_{\hat{G}}$ obtained from the estimated DAG.

4. COVID-19 Drivers and Country Risks

We next extend the analysis of Qazi and Simsekler (2022) who developed a BN to investigate how various country risks and risks associated to the COVID-19 epidemics relate to each

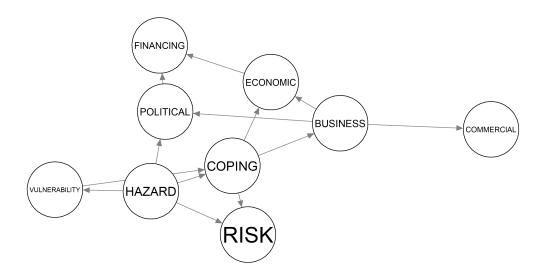


Figure 6: BN learned for the COVID-19 drivers and country risks.

other. In particular here we focus on how various types of risks affect the overall country risk associated to COVID-19.

For this purpose, as in Qazi and Simsekler (2022), the dataset used in the analysis comes from the combination of two sources¹. Country-level exposure to COVID-19 risks are retrieved from INFORM (INFORM, 2022). The data comprises of a score between zero and ten for 191 countries for three drivers of COVID-19 risks, namely hazard and exposure, vulnerability, and lack of coping capacity. An overall COVID-19 risk index, again between zero and ten, is constructed from these three drivers. Country-level exposure to various socioeconomic risk factors are collected from Euler Hermes (Euler Hermes, 2022). The ratings for five drivers of country risk, namely economic, political, financing, commercial and business environment are collected for 188 countries (the indexes are integer-valued between one and four or six). The combined dataset comprises 181 countries. Each variable is discretized into two levels using the clustering method from the arules package (Hahsler et al., 2005).

A BN is learned using the hc function of the bnlearn package with the constraint that the overall COVID-19 risk must be a leaf of the network and is reported in Figure 6. Without specifying it, the learned BN is such that each vertex has at most two parents. The DAG suggests that COVID-19 risk is conditionally independent of all other drivers given the lack of coping capacity and hazard & exposure.

Starting from this BN, a staged tree in the class of 2-parents staged trees is learned using the algorithm of Section 2.4. This staged tree provides a better representation of the data since it has a BIC of 1521.965, compared to the BIC of 1547.634 of the BN. The learned staged tree embeds the same set of symmetric conditional independences as in the BN of Figure 2, but also non-symmetric ones. Of course, the full tree cannot be easily visualized since, for instance, there are $2^9 = 512$ vertices with depth nine. However, since it is known that COVID-19 risk only depends on the lack of coping capacity and hazard & exposure

^{1.} Compared to Qazi and Simsekler (2022), we use more recent versions of the datasets.

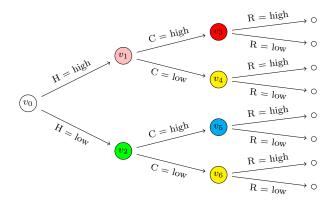


Figure 7: 2-parents staged tree for the COVID-19 risk constructed over hazard & exposure (H), lack of coping capacity (C) and COVID-19 risk (R).

we can construct the staged tree over these three variables only and easily visualize further non-symmetric dependences. This is reported in Figure 7, which shows the presence of the context-specific independence between COVID-19 risk and hazard & exposure for lack of coping capacity equal to low. Similar interpretations could be drawn by constructing the "partial" staged trees associated to other variables.

Of course a generic staged tree would provide a better representation of the data. For instance, one learned with the backward hill-climbing of Carli et al. (2022) starting from the saturated model has a BIC of 1453.587. However, its complete visualization is again unfeasible and plots as the one of Figure 7 are in general not viable since there are no constraints on the number of parents of G_T . Indeed, whilst the DAG G_T for the tree in Figure 7 has 13 edges (as the DAG in Figure 6), the DAG G_T from the generic staged tree is complete and consisting of 36 edges, meaning that all variables are directly related to one another.

5. Discussion

We defined a novel sub-class of staged tree models borrowing the idea of limiting the number of parents in BNs. A structural learning algorithm for such a class has been introduced and its properties illustrated both in simulation experiments and in a real-world application.

The number of parents of a variable in the staged tree is limited by those of the learned BN. For instance, in our application, although no limit on the number of parents was set, there were at most two parents. However, non-symmetric dependences may have been missed by the BN learning algorithm which is specifically designed to account for symmetric ones. One possibility could be to add edges to the learned BN, for instance between variables such that their conditional mutual information is large, and then run a backward hill-climbing algorithm of Carli et al. (2022) over that BN. The feasibility of such an algorithm is the focus of current research.

Acknowledgements

Gherardo Varando's work was funded by the European Research Council (ERC) Synergy Grant "Understanding and Modelling the Earth System with Machine Learning (USMILE)" under Grant Agreement No 855187.

References

- L. M. Barclay, J. L. Hutton, and J. Q. Smith. Refining a Bayesian network using a chain event graph. *International Journal of Approximate Reasoning*, 54:1300–1309, 2013.
- C. Boutilier, N. Friedman, M. Goldszmidt, and D. Koller. Context-specific independence in Bayesian networks. In *Proceedings of the Twelfth Conference on Uncertainty in Artificial Intelligence*, pages 115–123, 1996.
- F. Carli, M. Leonelli, and G. Varando. A new class of generative classifiers based on staged tree models. *arXiv:2012.13798*, 2020.
- F. Carli, M. Leonelli, E. Riccomagno, and G. Varando. The R package stagedtrees for structural learning of stratified staged trees. *Journal of Statistical Software*, 102(6):1–30, 2022.
- D. M. Chickering, D. Heckerman, and C. Meek. A Bayesian approach to learning Bayesian networks with local structure. In *Proceedings of 13th Conference on Uncertainty in Artificial Intelligence*, pages 80–89, 1997.
- R. Collazo, C. Görgen, and J. Smith. Chain event graphs. Chapmann & Hall, 2018.
- E. Duarte and L. Solus. Algebraic geometry of discrete interventional models. $arXiv:2012.03593,\ 2020.$
- E. Duarte and L. Solus. Representation of context-specific causal models with observational and interventional data. arXiv:2101.09271, 2021.
- Euler Hermes. Country risk reports. Retrieved from: https://www.eulerhermes.com/en_global/economic-research/country-reports.html, 2022.
- G. Freeman and J. Q. Smith. Bayesian MAP model selection of chain event graphs. *Journal of Multivariate Analysis*, 102(7):1152–1165, 2011.
- N. Friedman and M. Goldszmidt. Learning Bayesian networks with local structure. In Proceedings of the 12th Conference on Uncertainty in Artificial Intelligence, pages 252–262, 1996.
- N. Friedman, I. Nachman, and D. Pe'er. Learning Bayesian network structure from massive datasets: The "sparse candidate" algorithm. In *Proceedings of the 15th Conference on Uncertainty in Artificial Intelligence*, pages 206–215, 1999.
- C. Görgen, M. Leonelli, and O. Marigliano. The curved exponential family of a staged tree. *Electronic Journal of Statistics*, 16(1):2607–2620, 2022.

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- M. Hahsler, B. Gruen, and K. Hornik. arules A computational environment for mining association rules and frequent item sets. *Journal of Statistical Software*, 14(15):1–25, 2005.
- A. Hyttinen, J. Pensar, J. Kontinen, and J. Corander. Structure learning for Bayesian networks over labeled DAGs. In *International Conference on Probabilistic Graphical Models*, pages 133–144, 2018.
- INFORM. COVID-19 risk index. Retrieved from: https://drmkc.jrc.ec.europa.eu/inform-index/INFORM-Covid-19, 2022.
- M. Leonelli and G. Varando. Context-specific causal discovery for categorical data using staged trees. arXiv:2106.04416, 2021.
- M. Leonelli and G. Varando. Structural learning of simple staged trees. arXiv:2203.04390, 2022.
- F. Nicolussi and M. Cazzaro. Context-specific independencies in stratified chain regression graphical models. *Bernoulli*, 27(3):2091–2116, 2021.
- J. Pearl. Probabilistic reasoning in intelligent systems: networks of plausible inference. Morgan Kaufmann, 1988.
- J. Pensar, H. Nyman, J. Lintusaari, and J. Corander. The role of local partial independence in learning of Bayesian networks. *International Journal of Approximate Reasoning*, 69: 91–105, 2016.
- A. Qazi and M. C. E. Simsekler. Nexus between drivers of COVID-19 and country risks. *Socio-Economic Planning Sciences*, page 101276, 2022.
- M. Scutari. Learning Bayesian networks with the bnlearn R package. *Journal of Statistical Software*, 35(3):1–22, 2010.
- M. Scutari, C. E. Graafland, and J. M. Gutiérrez. Who learns better Bayesian network structures: Accuracy and speed of structure learning algorithms. *International Journal of Approximate Reasoning*, 115:235–253, 2019.
- T. Silander and T. Leong. A dynamic programming algorithm for learning chain event graphs. In *Proceedings of the International Conference on Discovery Science*, pages 201–216, 2013.
- J. Smith and P. Anderson. Conditional independence and chain event graphs. *Artificial Intelligence*, 172(1):42 68, 2008.
- T. Talvitie, R. Eggeling, and M. Koivisto. Learning Bayesian networks with local structure, mixed variables, and exact algorithms. *International Journal of Approximate Reasoning*, 115:69–95, 2019.
- I. Tsamardinos, L. E. Brown, and C. F. Aliferis. The max-min hill-climbing Bayesian network structure learning algorithm. *Machine Learning*, 65(1):31–78, 2006.
- G. Varando, F. Carli, and M. Leonelli. Staged trees and asymmetry-labeled DAGs. arXiv:2108.01994, 2021.