

## Appendices

### A Detailed Proofs of Theorem 1

#### A.1 Proof of Lemma 1

*Proof of Lemma 1.* Define that  $R(s) = \sum_{t=1}^s \max_{k \in \mathcal{K}} X_{k,t} - X_{A_t,t}$ , then we have that  $\mathcal{R}(T) = \mathbb{E}[R(T)]$ .

$$\begin{aligned} \mathcal{R}(T) &= \mathbb{E}[R(T)] \\ &= \mathbb{E}[R(T)\mathbb{I}\{\tau_1 \leq T\}] + \mathbb{E}[R(T)\mathbb{I}\{\tau_1 > T\}] \\ &\leq T \cdot \mathbb{P}(\tau_1 \leq T) + \mathbb{E}[R(T)\mathbb{I}\{\tau_1 > T\}]. \end{aligned} \quad (10)$$

Define  $N_k(t)$  as the number of times arm  $k$  has been selected by the Algorithm 2 in the first  $t$  steps, i.e.,  $N_k(t) = \sum_{i=1}^t \mathbb{I}(A_i = k)$ . No false alarm is raised and we do not restart the UCB algorithm if the event  $\{\tau_1 > T\}$  happens. Therefore, we have the following equation:

$$\mathbb{E}[R(T)\mathbb{I}\{\tau_1 > T\}] = \sum_{\Delta_k^{(1)} > 0} \Delta_k^{(1)} \cdot \mathbb{E}[N_k(T)\mathbb{I}\{\tau_1 > T\}].$$

Thus, it remains to show an upper bound for  $\mathbb{E}[N_k(T)\mathbb{I}\{\tau_1 > T\}]$ . By the definition of Algorithm 2, we have that for any  $k \in \mathcal{K}$

$$\begin{aligned} &N_k(T)\mathbb{I}\{\tau_1 > T\} \\ &= \sum_{t=1}^T \mathbb{I}\{A_t = k, \tau_1 > T, N_k(t) < l\} \\ &\quad + \sum_{t=1}^T \mathbb{I}\{A_t = k, \tau_1 > T, N_k(t) \geq l\} \\ &\leq l + \sum_{t=1}^T \mathbb{I}\{t \bmod \lfloor K/\gamma \rfloor = k, N_k(t) \geq l\} \\ &\quad + \sum_{t=1}^T \mathbb{I}\{k = \underset{\bar{k} \in \mathcal{K}}{\operatorname{argmax}} \operatorname{UCB}_{\bar{k}}, N_k(t) \geq l\} \\ &\leq l + \lceil T\gamma/K \rceil + \sum_{t=1}^T \mathbb{I}\{k = \underset{\bar{k} \in \mathcal{K}}{\operatorname{argmax}} \operatorname{UCB}_{\bar{k}}, N_k(t) \geq l\}, \end{aligned} \quad (11)$$

where the first inequality is due to the fact that if the event  $\{A_t = k, \tau_1 > T\}$  happens, then we do not restart the UCB algorithm before time  $T$  and the selection of the  $k$ th arm is based on either the uniform sampling or the largest UCB index in a stochastic bandit setting. Setting  $l = \lceil 8 \log T / (\Delta_k^{(1)})^2 \rceil$  and following the same argument as in the proof of Theorem 1 of [Auer et al., 2002a], we have that

$$\mathbb{E}[N_k(T)\mathbb{I}\{\tau_1 > T\}] \leq \frac{T\gamma}{K} + \frac{8 \log T}{(\Delta_k^{(1)})^2} + 1 + \frac{\pi^2}{3} + K.$$

Summing over  $k \in \mathcal{K}$  we prove the result. □

#### A.2 Proof of Lemma 2

*Proof of Lemma 2.* Define  $\tau_{k,1}$  as the first detection time of the  $k$ th arm. Then,  $\tau_1 = \min_{k \in \mathcal{K}} \{\tau_{k,1}\}$  since Algorithm 2 is designed to reinitialize the UCB algorithm if a change is detected on any of the  $K$  arms. Using the union bound, we have

that

$$\mathbb{P}(\tau_1 \leq T) \leq \sum_{k=1}^K \mathbb{P}(\tau_{k,1} \leq T).$$

Define that for any  $k \in \mathcal{K}$  and  $t \geq w$

$$S_{k,t} = \left| \sum_{i=t-w/2+1}^{i=t} Z_{k,i} - \sum_{i=t-w+1}^{t-w/2} Z_{k,i} \right|. \quad (12)$$

Then, for any  $k \in \mathcal{K}$ ,  $\tau_{k,1}$  is given by

$$\tau_{k,1} = \inf\{t \geq w : S_{k,t} > b\}$$

Let  $\mathbb{Z}^+$  be the set of all positive integers. Define that for any  $0 \leq j \leq w-1$  the stopping times

$$\tau_{k,1}^{(j)} = \inf\{t = j + nw, n \in \mathbb{Z}^+ : S_{k,t} > b\}.$$

We have that  $\tau_{k,1} = \min\{\tau_1^{(0)}, \dots, \tau_1^{(w-1)}\}$ . Note that under the stationary environment, for any  $0 \leq j \leq w-1$ ,  $\tau_{k,1}^{(j)}$  is a random variable with the geometric distribution

$$\mathbb{P}(\tau_{k,1}^{(j)} = nw + j) = p(1-p)^{n-1},$$

where  $p = \mathbb{P}(S_{k,w} > b)$ . Therefore, considering union bound we have that for any  $k \in \mathcal{K}$

$$\mathbb{P}(\tau_{k,1} \leq T) \leq w \left(1 - (1-p)^{\lfloor T/w \rfloor}\right).$$

The remaining task is to find an upper bound for  $p$ . Note that for any  $k \in \mathcal{K}$ ,  $S_{k,w}$  is a random variable with zero mean. We have by the McDiarmid's inequality and the union bound that

$$p \leq 2 \cdot \exp\left(-\frac{2b^2}{w}\right).$$

Combining the above analysis we conclude the result.  $\square$

### A.3 Proof of Lemma 3

*Proof of Lemma 3.* Assume that  $\delta_k^{(1)} \geq 2b/w + c$  for some  $\tilde{k} \in \mathcal{K}$ . Since the uniformly sampling scheme (step 2-4 of Algorithm 2) guarantees that in any time interval with length larger than  $L/2$  each arm is sampled at least  $w/2$  times, conditioning on  $\{\tau_1 > \nu_1\}$ , we have that

$$\begin{aligned} & \mathbb{P}(\nu_1 < \tau_1 \leq \nu_1 + L/2 \mid \tau_1 > \nu_1) \\ & \geq \mathbb{P}\left(S_{\tilde{k},w} > b\right) \\ & \geq 1 - 2 \exp\left(-\frac{(w|\delta_{\tilde{k}}^{(1)}|/2 - b)^2}{w}\right) \\ & \geq 1 - 2 \exp\left(-\frac{wc^2}{4}\right), \end{aligned} \quad (13)$$

where  $S_{k,t}$  is defined in (12) and we use McDiarmid's inequality in the second inequality.  $\square$

#### A.4 Proof of Lemma 4

*Proof of Lemma 4.* First, define that  $N = \lceil b/\delta_k^{(1)} \rceil \cdot \lceil K/\gamma \rceil$ , we obtain a simple upper bound for the EDD as follows.

$$\begin{aligned} & \mathbb{E}[\tau_1 - \nu_1 \mid \nu_1 < \tau_1 \leq \nu_1 + L/2] \\ &= \sum_{i=1}^{L/2} \mathbb{P}(\tau_1 \geq \nu_1 + i \mid \nu_1 < \tau_1 \leq \nu_1 + L/2) \\ &\leq N + \sum_{i=N}^{L/2} (\mathbb{P}(\tau_1 \geq \nu_1 + i \mid \nu_1 < \tau_1 \leq \nu_1 + L/2)). \end{aligned}$$

Since the uniformly sampling scheme guarantees that we have at least  $i/\lceil K/\gamma \rceil$  samples from each arm within  $i$  time steps, we use McDiarmid's inequality and Lemma 3 to have that

$$\begin{aligned} & \sum_{i=N}^{L/2} (\mathbb{P}(\tau_1 \geq \nu_1 + i \mid \nu_1 < \tau_1 \leq \nu_1 + L/2)) \\ &= \sum_{i=N}^{L/2} \frac{\mathbb{P}(\nu_1 + i \leq \tau_1 \leq \nu_1 + L/2 \mid \tau_1 > \nu_1)}{\mathbb{P}(\nu_1 \leq \tau_1 \leq \nu_1 + L/2 \mid \tau_1 > \nu_1)} \\ &\leq \frac{1}{1 - 2 \exp(-wc^2/4)} \cdot \sum_{i=N}^{L/2} 2 \exp\left(-\frac{(i/\lceil K/\gamma \rceil \delta_k^{(1)} - b)^2}{w}\right) \\ &\leq \frac{\lceil K/\gamma \rceil}{1 - 2 \exp(-wc^2/4)} \cdot \sum_{j=\lceil b/\delta_k^{(1)} \rceil}^{w/2} 2 \exp\left(-\frac{(j\delta_k^{(1)} - b)^2}{w}\right). \end{aligned}$$

Define  $q = \lceil (w/2) \cdot \delta_k^{(1)} \rceil - b$  and we have  $q > 1$  from the assumption that  $\delta_k^{(1)} > 2b/w + c$ . Combining the above analysis, we have that

$$\begin{aligned} & (1 - 2 \exp(-wc^2/4)) \cdot \mathbb{E}[\tau_1 - \nu_1 \mid \nu_1 < \tau_1 \leq \nu_1 + L/2] \\ &\leq N + \lceil K/\gamma \rceil \cdot \sum_{j=\lceil b/\delta_k^{(1)} \rceil}^{w/2} 2 \exp\left(-\frac{(j\delta_k^{(1)} - b)^2}{w}\right) \\ &\leq N + 2\lceil K/\gamma \rceil \cdot \left(1 + \int_1^q \exp\left(-\frac{l^2}{w}\right) dl\right) \\ &\leq N + 2\lceil K/\gamma \rceil \cdot \left[1 + \sqrt{w} \left(1 - \frac{1}{\sqrt{w}} + \int_1^{q/\sqrt{w}} \exp(-u^2) du\right)\right] \\ &\leq N + 2\lceil K/\gamma \rceil \cdot \left(\sqrt{w} + \sqrt{w} \int_1^{q/\sqrt{w}} u \exp(-u^2) du\right) \\ &\leq \left(\lceil b/\delta_k^{(1)} \rceil + 3\sqrt{w}\right) \cdot \lceil K/\gamma \rceil, \end{aligned}$$

where we transform  $l$  into  $u = l/\sqrt{w}$  in the third inequality and we use the fact that  $\exp(-u^2) \leq u \exp(-u^2)$ ,  $u \geq 1$  in the fourth inequality. On the other hand, by the definition of the conditioning event we also have that

$$(1 - 2 \exp(-wc^2/4)) \cdot \mathbb{E}[\tau_1 - \nu_1 \mid \nu_1 < \tau_1 \leq \nu_1 + L/2] \leq L/2.$$

Combining the above analysis we conclude the result.  $\square$

### A.5 Proof of Theorem 1

*Proof of Theorem 1.* Recall that  $L = w \lceil K/\gamma \rceil$ . Algorithm 2 guarantees that in any time interval with length larger than  $L$  each arm is sampled at least  $w$  times. Define events  $F_i = \{\tau_i > \nu_i\}$ ,  $1 \leq i \leq M-1$ . Define events  $D_i = \{\tau_i \leq \nu_i + L/2\}$ ,  $1 \leq i \leq M-2$  and event  $D_{M-1} = \{\tau_{M-1} \leq T\}$ . Therefore, the event  $F_i D_i$  is the good event where the  $i$ th change can be detected correctly and efficiently. Define that  $R(s) = \sum_{t=1}^s \max_{k \in \mathcal{K}} X_{k,t} - X_{A_t,t}$ , then we have that  $\mathcal{R}(T) = \mathbb{E}[R(T)]$ . Equipped with the sequence of good events, we have that

$$\begin{aligned} \mathcal{R}(T) &= \mathbb{E}[R(T)] \leq \mathbb{E}[R(T) \mathbb{I}\{F_1\}] + T \cdot (1 - \mathbb{P}(F_1)) \\ &\leq \mathbb{E}[R(\nu_1) \mathbb{I}\{F_1\}] + \mathbb{E}[R(T) - R(\nu_1)] + 1 \\ &\leq \tilde{C}_1 + \gamma \nu_1 + \mathbb{E}[R(T) - R(\nu_1)] + 1. \end{aligned}$$

Above, the second inequality is due to Lemma 2 that  $\mathbb{P}(F_1) \geq 1 - 1/T$  provided that  $b = [w \log(2KT^2)/2]^{1/2}$  and the third inequality is due to the bound in the end of Lemma 1, which is the bound for the UCB algorithm in a stochastic bandit setting.

The next step is to bound  $\mathbb{E}[R(T) - R(\nu_1)]$ . Using the law of total expectation, we have that

$$\begin{aligned} &\mathbb{E}[R(T) - R(\nu_1)] \\ &\leq \mathbb{E}[R(T) - R(\nu_1) \mid F_1 D_1] + T \cdot (1 - \mathbb{P}(F_1 D_1)) \\ &\leq \mathbb{E}[R(T) - R(\nu_1) \mid F_1 D_1] + 2, \end{aligned}$$

where the last inequality is due to Lemma 3 that we have  $\mathbb{P}(D_1 \mid F_1) \geq 1 - 1/T$  provided that  $c = 2\sqrt{\log(2T)/w}$  and the fact that  $\mathbb{P}(F_1 D_1) = \mathbb{P}(D_1 \mid F_1) \cdot \mathbb{P}(F_1)$  for any probability measure  $\mathbb{P}$ .

Therefore, the remaining task is to bound  $\mathbb{E}[R(T) - R(\nu_1) \mid F_1 D_1]$ . Denote  $\tilde{\mathbb{E}}$  as the expectation according to the piecewise-stationary bandit starts from the second segment. Further splitting the regret, we have that

$$\begin{aligned} &\mathbb{E}[R(T) - R(\nu_1) \mid F_1 D_1] \\ &\leq \mathbb{E}[R(T) - R(\tau_1) \mid F_1 D_1] + \mathbb{E}[R(\tau_1) - R(\nu_1) \mid F_1 D_1] \\ &\leq \tilde{\mathbb{E}}[R(T - \nu_1)] + \mathbb{E}[\tau_1 - \nu_1 \mid F_1 D_1] \\ &\leq \tilde{\mathbb{E}}[R(T - \nu_1)] + \min(L/2, (\lceil b/\delta^{(1)} \rceil + 3\sqrt{w}) \cdot \lceil K/\gamma \rceil) / (1 - 1/T) \end{aligned}$$

where the second inequality is due to the renewal property given that the whole algorithm restarts in the time interval between  $\nu_1$  and  $\nu_1 + L/2$  and the last inequality is due to Lemma 4 by setting  $c = 2\sqrt{\log(2T)/w}$ .

Combining the above analysis, we bound the regret in a recursive manner as follows (assuming  $T \geq 2$ ):

$$\begin{aligned} \mathbb{E}[R(T)] &\leq \tilde{\mathbb{E}}[R(T - \nu_1)] + \tilde{C}_1 \\ &\quad + \gamma \nu_1 + 2 \min(L/2, (\lceil b/\delta^{(1)} \rceil + 3\sqrt{w}) \cdot \lceil K/\gamma \rceil) + 3. \end{aligned}$$

The recursive manner means that we can apply the same method to bound  $\tilde{\mathbb{E}}[R(T - \nu_1)]$ , by conditioning on the event  $D_2 F_2$ . Repeating this procedure  $M-1$  times, we obtain that

$$\begin{aligned} \mathbb{E}[R(T)] &\leq \sum_{i=1}^M \tilde{C}_i + \gamma T \\ &\quad + \sum_{i=1}^{M-1} \frac{2K \cdot \min(\frac{w}{2}, \lceil \frac{b}{\delta^{(i)}} \rceil + 3\sqrt{w})}{\gamma} + 3M. \end{aligned}$$

□