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Approximation by quantization of the filter process and applications to optimal stopping problems under partial observation

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Abstract

We present an approximation method for discrete time nonlinear filtering in view of solving dynamic optimization problems under partial information. The method is based on quantization of the Markov pair process filter-observation (Π, Y) and is such that, at each time step k and for a given size N_k of the quantization grid in period k , this grid is chosen to minimize a suitable quantization error. The algorithm is based on a stochastic gradient descent combined with Monte-Carlo simulations of (Π, Y) . Convergence results are given and applications to optimal stopping under partial observation are discussed. Numerical results are presented for a particular stopping problem : American option pricing with unobservable volatility.

Key words: Nonlinear filtering, Markov chain, quantization, stochastic gradient descent, Monte-Carlo simulations, partial observation, optimal stopping.

MSC Classification (2000): 60G35, 65C20, 65N50, 62L20, 60G40.

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1 Introduction

We consider a discrete time, partially observable process (X, Y) where X represents the state or signal process that may not be observable, while Y is the observation. The signal process $\{X_k, k \in \mathbb{N}\}$ is valued in a measurable space (E, \mathcal{E}) and is a Markov chain with probability transition (P_k) (i.e. the transition from time $k-1$ to time k), and initial law μ . The observation sequence (Y_k) is valued in \mathbb{R}^d and such that the pair (X_k, Y_k) is a Markov chain and

(H) The law of Y_k conditional on (X_{k-1}, Y_{k-1}, X_k) , $k \geq 1$, denoted $q_k(X_{k-1}, Y_{k-1}, X_k, dy')$, admits a bounded density

$$y' \longmapsto g_k(X_{k-1}, Y_{k-1}, X_k, y').$$

For simplicity, we assume that Y_0 is a known deterministic constant equal to y_0 . Notice that the probability transition of the Markov chain $(X_k, Y_k)_{k \in \mathbb{N}}$ is then given by $P_k(x, dx')g_k(x, y, x', y')dy'$ with initial law $\mu(dx)\delta_{y_0}(dy)$.

We denote by (\mathcal{F}_k^Y) the filtration generated by the observation process (Y_k) and by Π_k the filter conditional law of X_k given \mathcal{F}_k^Y :

$$\Pi_k(dx) = \mathbb{P}[X_k \in dx | \mathcal{F}_k^Y], \quad k \in \mathbb{N}.$$

The filter process $(\Pi_k)_k$ allows one to transform problems related to the partially observed process (X, Y) into equivalent problems under complete observation related to the pair (Π, Y) and this latter pair turns out to be Markov with respect to the observation filtration (\mathcal{F}_k^Y) . An important class of problems related to partially observable processes are stochastic control and stopping problems that are dynamic stochastic optimization problems where the partial observation concerns the state/signal process. We shall, in particular, consider optimal stopping under partial information and, in this context, a financial application concerning the pricing of American options in a partially observed stochastic volatility model.

We shall assume that the state space E of the state/signal process (X_k) consists of a finite number m of points. Problems with a more general state space can be approximated by problems with a finite state space (see e.g. [8]). With X_k taking the m values x^i ($i = 1, \dots, m$), the filter process is characterized by an m -vector with components $\Pi_k^i = \mathbb{P}[X_k = x^i | \mathcal{F}_k^Y]$ and takes thus values in the m -simplex K_m in \mathbb{R}^m . While the filter process allows thus to transform a problem under partial observation into one under full observation, it has the drawback that a finite-valued state space becomes infinite. This leads to difficulties when trying to solve dynamic optimization problems also in discrete time. For actual computation one has thus to discretize the process (Π_k) , approximating it with a process that takes a finite number of values in the simplex K_m . Approaches to this effect have appeared in the literature; they are based on discretizing the observations process (Y_k) and then approximating Π_k by a filter of X_k , given the discretized observations path (see e.g. [5], [3], [15], [10] and references therein). Such an approach has the drawback that the number of possible values for the approximating filter grows exponentially fast with the time step.

In this paper, we propose a new approach by exploiting the Markov property of $(\Pi, Y)_k$ with respect to the observation filtration (\mathcal{F}_k^Y) : the conditional law of X_{k+1} given \mathcal{F}_k^Y is summarized by the sufficient statistics (Π_k, Y_k) valued in $K^m \times \mathbb{R}^d$. This suggests to approximate the couple process (Π, Y) with $(\hat{\Pi}, \hat{Y})$ that in the generic time step k takes a number of values N_k that can be assigned arbitrarily. Following standard usage, we shall call this *quantization* and the problem then arises to find an *optimal* quantization, namely such that it minimizes for each time period k the L^2 -approximation error induced by the quantization. This error is related to what e.g. in information theory is called *distorsion*. The implementation of the minimizing algorithm itself is based on a stochastic gradient descent method combined with Monte-Carlo simulations of (Π, Y) . As a byproduct, we obtain an approximation of the probability transition matrices of the Markov chain (Π, Y) . These companion parameters provide in turn an approximation algorithm, via the dynamic programming principle, for computing optimal values associated to dynamic optimization problems under partial observation.

Optimal quantization methods have been developed recently in numerical probability and various problems of optimal stopping, control or nonlinear filtering, see [11], [1], [2], [13], [14], [12]. In our context, we obtain error bounds and rate of convergence for the approximation of the pair filter-observation process (Π, Y) . We then give an application to the approximation of optimal stopping problem under partial observation and study associated convergence results. Finally, we present a numerical illustration for the American option pricing problem with unobservable volatility.

The rest of the paper is organized as follows. In Section 2, we recall some preliminaries on nonlinear filtering. Section 3 is the heart of the paper. We first present the quantization approximation method for the filter-observation Markov chain, and then analyze the induced error. The rate of convergence and the practical implementation of the algorithm for the optimal quantization are discussed. In Section 4, we apply our quantization method to the approximation of an optimal stopping problem under partial observation. Convergence results for the associated optimal value functions are provided. The last Section 5 presents numerical tests for the American option pricing with unobservable volatility.

2 Preliminaries

We denote by $\mathcal{M}(E)$ the set of finite nonnegative measures on (E, \mathcal{E}) and by $\mathcal{P}(E)$ the subset of probability measures on (E, \mathcal{E}) . It is known that $\mathcal{M}(E)$ is a Polish space equipped with the weak topology, hence a measurable space endowed with the Borel σ -field. We recall that from Bayes formula, Π_k is given in inductive form by :

$$\begin{aligned} \Pi_0 &= \mu \\ \Pi_k &= \bar{G}_k(\Pi_{k-1}, Y_{k-1}, Y_k), \quad k \geq 1, \end{aligned} \tag{2.1}$$

where G_k is the continuous function from $\mathcal{M}(E) \times \mathbb{R}^d \times \mathbb{R}^d$ into $\mathcal{M}(E)$ defined by :

$$G_k(\pi, y, y')(dx') = \int_E g_k(x, y, x', y') P_k(x, dx') \pi(dx),$$

and \bar{G}_k is the normalized continuous function valued in $\mathcal{P}(E)$:

$$\bar{G}_k(\pi, y, y') = \frac{G_k(\pi, y, y')}{\int_E G_k(\pi, y, y')(dx')}$$

Denoting by (\mathcal{F}_k) the filtration generated by (X_k, Y_k) , and using the law of iterated conditional expectations, we have for any k and bounded Borelian function φ on $\mathcal{P}(E) \times \mathbb{R}^d$:

$$\begin{aligned} & \mathbb{E} [\varphi(\Pi_{k+1}, Y_{k+1}) | \mathcal{F}_k^Y] \\ &= \mathbb{E} [\mathbb{E}[\varphi(\bar{G}_{k+1}(\Pi_k, Y_k, Y_{k+1}), Y_{k+1}) | \mathcal{F}_k] | \mathcal{F}_k^Y] \\ &= \mathbb{E} \left[\int \varphi(\bar{G}_{k+1}(\Pi_k, Y_k, y'), y') P_{k+1}(X_k, dx') q_{k+1}(X_k, Y_k, x', dy') \middle| \mathcal{F}_k^Y \right] \\ &= \int \varphi(\bar{G}_{k+1}(\Pi_k, Y_k, y'), y') P_{k+1}(x, dx') q_{k+1}(x, Y_k, x', dy') \Pi_k(dx). \end{aligned} \quad (2.2)$$

This proves that the pair $(\Pi_k, Y_k)_k$ is a $(\mathbb{P}, (\mathcal{F}_k^Y)_k)$ Markov chain in $\mathcal{P}(E) \times \mathbb{R}^d$, with initial value (μ, y_0) . Moreover, under **(H)**, this also shows that the (unnormalized) law of Y_k conditional on (Π_{k-1}, Y_{k-1}) , denoted $Q_k(\Pi_{k-1}, Y_{k-1}, dy')$ admits a density given by :

$$y' \longmapsto \int G_k(\Pi_{k-1}, Y_{k-1}, y')(dx') = \int g_k(x, Y_{k-1}, x', y') P_k(x, dx') \Pi_{k-1}(dx). \quad (2.3)$$

Notice that the probability transition R_k (from time $k-1$ to k) of the Markov chain $(\Pi_k, Y_k)_k$ is not explicit in general. Actually, from (2.2), we can write :

$$R_k \varphi(\pi, y) = \int \varphi(\bar{G}_k(\pi, y, y'), y') Q_k(\pi, y, dy'), \quad \forall (\pi, y) \in \mathcal{P}(E) \times \mathbb{R}^d. \quad (2.4)$$

In the sequel, we denote by $|\cdot|_2$ the Euclidian norm and by $|\cdot|_1$ the l^1 norm on \mathbb{R}^l . For any \mathbb{R}^l -valued random variable U , we denote :

$$\|U\|_2 = (\mathbb{E}|U|_2^2)^{\frac{1}{2}} \quad \text{and} \quad \|U\|_1 = \mathbb{E}|U|_1.$$

3 An optimal quantization approach for the approximation of the filter process

3.1 The approximation method

We assume here that the state space E of the signal process (X_k) is finite consisting of m points : $E = \{x^1, \dots, x^m\}$. The initial discrete law $\mu = (\mu^i)$ and the probability transition matrix P_k are defined by :

$$\begin{aligned} \mu^i &= \mathbb{P}[X_0 = x^i], \quad i = 1, \dots, m, \\ P_k^{ij} &= \mathbb{P}[X_k = x^j | X_{k-1} = x^i], \quad i, j = 1, \dots, m. \end{aligned}$$

The random filter Π_k is characterized by its random weights :

$$\Pi_k^i = \mathbb{P}[X_k = x^i | \mathcal{F}_k^Y], \quad i = 1, \dots, m,$$

and may then be identified with a random vector valued in the m -simplex K_m in \mathbb{R}^m of dimension $m - 1$:

$$K_m = \left\{ \pi = (\pi^i) \in \mathbb{R}^m : \pi^i \geq 0, 1 \leq i \leq m, |\pi|_1 = \sum_{i=1}^m \pi^i = 1 \right\}.$$

By (2.1), it is expressed in the recursive form :

$$\begin{aligned} \Pi_0 &= \mu \\ \Pi_k &= \bar{G}_k(\Pi_{k-1}, Y_{k-1}, Y_k) = \frac{GP_k(Y_{k-1}, Y_k)^\top \Pi_{k-1}}{|GP_k(Y_{k-1}, Y_k)^\top \Pi_{k-1}|_1}, \end{aligned} \quad (3.1)$$

where $GP_k(Y_{k-1}, Y_k)$ is a $m \times m$ random matrix given by :

$$GP_k(Y_{k-1}, Y_k)_{ij} = g_k(x_{k-1}^i, Y_{k-1}, x_k^j, Y_k) P_k^{ij}, \quad 1 \leq i, j \leq m. \quad (3.2)$$

Here M^\top denotes the transpose of a matrix M .

A first approach for approximating the filter process (Π_k) consists in discretizing the observation process (Y_k) by replacing it by a process (\hat{Y}_k) taking a finite number N of values, and then approximate for each k the random filter Π_k by a random filter of X_k given $\hat{Y}_1, \dots, \hat{Y}_k$. So at each time step k , given \hat{Y}_k and $\hat{\Pi}_k$, the random variable $\hat{\Pi}_{k+1}$ may take N values. Thus, at time n , the random filter $\hat{\Pi}_n$ is identified with a random vector taking N^n possible values. All these values are precomputed and stored in a “look-up table” but this could be very heavy, typically when n is large. Such an approach was introduced in [5] and [3], and investigated further in [15],[10] and [16].

We propose here a new approach. This starts from the key remark that the pair process (Π_k, Y_k) is Markov with respect to the observation filtration (\mathcal{F}_k^Y) . In other words, the conditional law of X_{k+1} given \mathcal{F}_k^Y may be summarized by the sufficient statistic (Π_k, Y_k) . Therefore, since the Markov chain (Π_k, Y_k) is completely characterized by its probability transitions, the idea is to approximate these probability transitions by suitable probability transition matrices.

In a first step, we discretize for each k the couple (Π_k, Y_k) by approximating it by $(\hat{\Pi}_k, \hat{Y}_k)$ taking a finite number of values. The space discretization (or quantization) of the random vector $Z_k = (\Pi_k, Y_k)$ in $K_m \times \mathbb{R}^d$ is constructed as follows. At initial time $k = 0$, recall that Z_0 is a known deterministic vector equal to $z_0 = (\mu, y_0)$, so we start from the grid with one point in $K_m \times \mathbb{R}^d$:

$$\Gamma_0 = \{z_0 = (\mu, y_0)\}.$$

At time $k \geq 1$, we are given a grid Γ_k of N_k points in $K_m \times \mathbb{R}^d$:

$$\Gamma_k = \left\{ z_k^1 = (\pi_k(1), y_k^1), \dots, z_k^{N_k} = (\pi_k(N_k), y_k^{N_k}) \right\},$$

and we denote by $C_i(\Gamma_k)$, $i = 1, \dots, N_k$, the associated Voronoi tessellations :

$$C_i(\Gamma_k) = \left\{ z \in K_m \times \mathbb{R}^d : \text{Proj}_{\Gamma_k}(z) = z_k^i \right\}, \quad i = 1, \dots, N_k.$$

Here Proj_{Γ_k} is a closest neighbor projection for the Euclidian norm :

$$\|z - \text{Proj}_{\Gamma_k}(z)\|_2 = \min_{i=1, \dots, N_k} \|z - z_k^i\|_2, \quad \forall z \in K_m \times \mathbb{R}^d.$$

Notice that by definition of the Euclidian norm, we clearly have :

$$\text{Proj}_{\Gamma_k} = \left(\text{Proj}_{\Gamma_k^\Pi}, \text{Proj}_{\Gamma_k^Y} \right), \quad C_i(\Gamma_k) = C_i(\Gamma_k^\Pi) \times C_i(\Gamma_k^Y),$$

where :

$$\begin{aligned} \Gamma_k^\Pi &= \{\pi_k(1), \dots, \pi_k(N_k)\} \\ \Gamma_k^Y &= \{y_k^1, \dots, y_k^{N_k}\}. \end{aligned}$$

We then approximate the pair $Z_k = (\Pi_k, Y_k)$ by $\hat{Z}_k = (\hat{\Pi}_k, \hat{Y}_k)$ valued in Γ_k and defined by :

$$\hat{Z}_k = \text{Proj}_{\Gamma_k}(Z_k) = \left(\text{Proj}_{\Gamma_k^\Pi}(\Pi_k), \text{Proj}_{\Gamma_k^Y}(Y_k) \right).$$

In a second step, we approximate the probability transitions of the Markov chain (Z_k) :

$$R_k(z, dz') = \mathbb{P}[Z_k \in dz' | Z_{k-1} = z], \quad k \geq 1, \quad z \in K_m \times \mathbb{R}^d,$$

by the following probability transition matrix :

$$\begin{aligned} \hat{r}_k^{ij} &= \mathbb{P}[\hat{Z}_k = z_k^j | \hat{Z}_{k-1} = z_{k-1}^i] \\ &= \frac{\mathbb{P}[Z_k \in C_j(\Gamma_k), Z_{k-1} \in C_i(\Gamma_{k-1})]}{\mathbb{P}[Z_{k-1} \in C_i(\Gamma_{k-1})]} =: \frac{\hat{\beta}_k^{ij}}{\hat{p}_{k-1}^i}, \end{aligned}$$

for all $k \geq 1$, $i = 1, \dots, N_{k-1}$, $j = 1, \dots, N_k$. We shall see later how the grids Γ_k and the number of points N_k are optimally chosen and implemented, and how the associated probability transition matrix \hat{r}_k can be estimated.

Example : Computation of predictor conditional expectation.

Suppose we are interested in the computation for any $k = 0, \dots, n$, and for arbitrary measurable function φ_{k+1} on $K_m \times \mathbb{R}^d$, of the filter predictor :

$$U_k = \mathbb{E}[\varphi_{k+1}(X_{k+1}, Y_{k+1}) | \mathcal{F}_k^Y].$$

A precise application where such \mathcal{F}_k^Y -measurable random variables appear is presented in Section 4. Then, by introducing the function :

$$\hat{\varphi}_{k+1}(\pi, y) = \sum_{i=1}^m \varphi_{k+1}(x^i, y) \pi^i, \quad \forall \pi = (\pi^i)_i \in K_m, \quad \forall y \in \mathbb{R}^d,$$

and using the law of iterated conditional expectation, we can rewrite U_k as :

$$U_k = \mathbb{E}[\hat{\varphi}_{k+1}(\Pi_{k+1}, Y_{k+1}) | \mathcal{F}_k^Y] = \mathbb{E}[\hat{\varphi}_{k+1}(Z_{k+1}) | \mathcal{F}_k^Y].$$

We thus approximate the sequence of \mathcal{F}_k^Y -measurable random variables U_k by $\hat{U}_k = \hat{v}_k(\hat{Z}_k)$ where the functions \hat{v}_k , $k = 0, \dots, n$, are defined on Γ_k by :

$$\begin{aligned} \hat{v}_k(z_k^i) &= \mathbb{E}[\hat{\varphi}_{k+1}(\hat{Z}_{k+1}) | \hat{Z}_k = z_k^i] \\ &= \sum_{j=1}^{N_{k+1}} \hat{r}_{k+1}^{ij} \hat{\varphi}_{k+1}(z_{k+1}^j), \quad \forall z_k^i \in \Gamma_k. \end{aligned}$$

3.2 The error analysis

The quality of the approximation described in the previous paragraph is measured as follows. We denote for any subset D in \mathbb{R}^l :

$$BL_1(D) = \left\{ \varphi \text{ Borelian from } D \text{ into } \mathbb{R} : \begin{aligned} \|\varphi\|_{sup} &:= \sup_{x \in D} |\varphi(x)| \leq 1, & [\varphi]_{lip} &:= \sup_{x, y \in D, x \neq y} \frac{|\varphi(x) - \varphi(y)|}{|x - y|_1} \leq 1 \end{aligned} \right\}.$$

We make the following assumption.

(H1) There exists a constant L_g such that for all $k \geq 1$:

$$\sum_{i, j=1}^m P_k^{ij} \int |g_k(x^i, y, x^j, y') - g_k(x^i, \hat{y}, x^j, y')| dy' \leq L_g |y - \hat{y}|_1, \quad \forall y, \hat{y} \in \mathbb{R}^d.$$

Proposition 3.1 Under **(H1)**, we have for any n and $\varphi_1, \dots, \varphi_n \in BL_1(K_m \times \mathbb{R}^d)$:

$$\begin{aligned} & \left| \int \varphi_1(z_1) \dots \varphi_n(z_n) (R_1(z_0, dz_1) \dots R_n(z_{n-1}, dz_n) - \hat{r}_1(z_0, dz_1) \dots \hat{r}_n(z_{n-1}, dz_n)) \right| \\ & \leq \sum_{k=1}^n \frac{3\sqrt{m+d}}{2\bar{L}_g - 1} \left[(2\bar{L}_g)^{n-k+1} - 1 \right] \|Z_k - \hat{Z}_k\|_2, \end{aligned} \quad (3.3)$$

where $\bar{L}_g = \max(L_g, 1)$.

We first state a preliminary result on the Lipschitz property of the transition of the Markov chain $(Z_k)_k = (\Pi_k, Y_k)_k$.

Lemma 3.1 Under **(H1)**, we have for all $k \geq 1$ and Borelian function φ on $K_m \times \mathbb{R}^d$:

$$|R_k \varphi(z) - R_k \varphi(\hat{z})| \leq (2[\varphi]_{lip} + \|\varphi\|_{sup}) \bar{L}_g |z - \hat{z}|_1, \quad \forall z, \hat{z} \in K_m \times \mathbb{R}^d.$$

Proof. Recall from (2.3) that the conditional law $Q_k(\pi, y, dy')$ of Y_k given $(\Pi_{k-1}, Y_{k-1}) = (\pi, y)$ admits a density given by :

$$y' \longmapsto f_k(\pi, y, y') = \sum_{i, j=1}^m g_k(x^i, y, x^j, y') P_k^{ij} \pi^i.$$

This conditional density satisfies the Lipschitz property : for all (π, y) and $(\hat{\pi}, \hat{y}) \in K_m \times \mathbb{R}^d$,

$$\begin{aligned} \int |f_k(\pi, y, y') - f_k(\hat{\pi}, \hat{y}, y')| dy' & \leq \sum_{i, j=1}^m P_k^{ij} \int |g_k(x^i, y, x^j, y') - g_k(x^i, \hat{y}, x^j, y')| dy' \\ & \quad + \sum_i^m |\pi^i - \hat{\pi}^i|, \end{aligned} \quad (3.4)$$

where we used the fact that $\sum_j P_k^{ij} \int g_k(x^i, y, x^j, y') dy' = 1$. From (2.4), we then have for any $z = (\pi, y)$ and $\hat{z} = (\hat{\pi}, \hat{y}) \in K_m \times \mathbb{R}^d$:

$$\begin{aligned}
|R_k \varphi(z) - R_k \varphi(\hat{z})| &\leq \int |\varphi(\bar{G}_k(\pi, y, y'), y') - \varphi(\bar{G}_k(\hat{\pi}, \hat{y}, y'), y')| Q_k(\pi, y, dy') \\
&\quad + \left| \int \varphi(\bar{G}_k(\hat{\pi}, \hat{y}, y'), y') (Q_k(\pi, y, dy') - Q_k(\hat{\pi}, \hat{y}, dy')) \right| \\
&\leq [\varphi]_{lip} \int |\bar{G}_k(\pi, y, y') - \bar{G}_k(\hat{\pi}, \hat{y}, y')|_1 f_k(\pi, y, y') dy' \\
&\quad + \|\varphi\|_{sup} \int |f_k(\pi, y, y') - f_k(\hat{\pi}, \hat{y}, y')| dy' \\
&\leq [\varphi]_{lip} \int |\bar{G}_k(\pi, y, y') - \bar{G}_k(\hat{\pi}, \hat{y}, y')|_1 f_k(\pi, y, y') dy' \\
&\quad + \|\varphi\|_{sup} \sum_{i,j=1}^m P_k^{ij} \int |g_k(x^i, y, x^j, y') - g_k(x^i, \hat{y}, x^j, y')| dy' \\
&\quad + \|\varphi\|_{sup} \sum_i^m |\pi^i - \hat{\pi}^i|. \tag{3.5}
\end{aligned}$$

Now, from (3.2), we have :

$$\begin{aligned}
&\int |\bar{G}_k(\pi, y, y') - \bar{G}_k(\hat{\pi}, \hat{y}, y')|_1 f_k(\pi, y, y') dy' \\
&\leq \sum_{j=1}^m \int |\bar{G}_k^j(\pi, y, y') - \bar{G}_k^j(\hat{\pi}, \hat{y}, y')| f_k(\pi, y, y') dy' \\
&= \sum_{i,j=1}^m \int \left| \frac{g_k(x^i, y, x^j, y') P_k^{ij} \pi^i}{f_k(\pi, y, y')} - \frac{g_k(x^i, \hat{y}, x^j, y') P_k^{ij} \hat{\pi}^i}{f_k(\hat{\pi}, \hat{y}, y')} \right| f_k(\pi, y, y') dy' \\
&\leq \sum_{i,j=1}^m P_k^{ij} \hat{\pi}^i \int \frac{|g_k(x^i, y, x^j, y') f_k(\hat{\pi}, \hat{y}, y') - g_k(x^i, \hat{y}, x^j, y') f_k(\pi, y, y')|}{f_k(\hat{\pi}, \hat{y}, y')} dy' \\
&\quad + \sum_{i=1}^m |\pi^i - \hat{\pi}^i| \\
&\leq \sum_{i,j=1}^m P_k^{ij} \int |g_k(x^i, y, x^j, y') - g_k(x^i, \hat{y}, x^j, y')| dy' \\
&\quad + \int |f_k(\pi, y, y') - f_k(\hat{\pi}, \hat{y}, y')| dy' + \sum_{i=1}^m |\pi^i - \hat{\pi}^i| \\
&\leq 2 \sum_{i,j=1}^m P_k^{ij} \int |g_k(x^i, y, x^j, y') - g_k(x^i, \hat{y}, x^j, y')| dy' + 2 \sum_{i=1}^m |\pi^i - \hat{\pi}^i|,
\end{aligned}$$

where we used again (3.4). Plugging into (3.5) and using **(H1)** yield :

$$\begin{aligned}
&|R_k \varphi(z) - R_k \varphi(\hat{z})| \\
&\leq (2[\varphi]_{lip} + \|\varphi\|_{sup}) \left(\sum_{i,j=1}^m P_k^{ij} \int |g_k(x^i, y, x^j, y') - g_k(x^i, \hat{y}, x^j, y')| dy' + \sum_{i=1}^m |\pi^i - \hat{\pi}^i| \right)
\end{aligned}$$

$$\leq (2[\varphi]_{lip} + \|\varphi\|_{sup}) (L_g |y - \hat{y}|_1 + |\pi - \hat{\pi}|_1), \quad (3.6)$$

and then the required result. \square

Remark 3.1 In the case where the law of Y_k conditional on (X_{k-1}, Y_{k-1}, X_k) does not depend on Y_{k-1} , i.e. its density g_k depends only X_{k-1}, X_k , the condition **(H1)** is empty, or in other words is trivially satisfied with $L_g = 0$. Then, inequality (3.6) shows that for all $z = (\pi, y)$, $\hat{z} = (\hat{\pi}, \hat{y}) \in K_m \times \mathbb{R}^d$,

$$|R_k \varphi(z) - R_k \varphi(\hat{z})| \leq (2[\varphi]_{lip} + \|\varphi\|_{sup}) |\pi - \hat{\pi}|_1. \quad (3.7)$$

Proof of Proposition 3.1. For $k = 1, \dots, n$, we define the measurable functions on $K_m \times \mathbb{R}^d$, resp. on Γ_k :

$$\begin{aligned} v_k(z) &= \varphi_k(z) \int \varphi_{k+1}(z_{k+1}) \dots \varphi_n(z_n) R_{k+1}(z, dz_{k+1}) \dots R_n(z_{n-1}, dz_n), \\ \hat{v}_k(z) &= \varphi_k(z) \int \varphi_{k+1}(z_{k+1}) \dots \varphi_n(z_n) \hat{r}_{k+1}(z, dz_{k+1}) \dots \hat{r}_n(z_{n-1}, dz_n), \end{aligned}$$

with the convention that for $k = n$, $v_n = \hat{v}_n = \varphi_n$, we then have the backward induction formulas :

$$v_k(z) = \varphi_k(z) R_{k+1} v_{k+1}(z) := \varphi_k(z) \mathbb{E}[v_{k+1}(Z_{k+1}) | Z_k = z] \quad (3.8)$$

$$\hat{v}_k(z) = \varphi_k(z) \hat{r}_{k+1} \hat{v}_{k+1}(z) := \varphi_k(z) \mathbb{E}[\hat{v}_{k+1}(\hat{Z}_{k+1}) | \hat{Z}_k = z], \quad (3.9)$$

for all $k = 1, \dots, n-1$.

Step 1. We clearly have $\|v_k\|_{sup} \leq 1$. Moreover, from (3.8) and using Lemma 3.1, we have :

$$\begin{aligned} [v_k]_{lip} &\leq [\varphi_k]_{lip} + [R_{k+1} v_{k+1}]_{lip} \\ &\leq 1 + \bar{L}_g + 2\bar{L}_g [v_{k+1}]_{lip}. \end{aligned}$$

Since $[v_n]_{lip} \leq 1$, a standard backward induction yields :

$$[v_k]_{lip} \leq \frac{\frac{3}{2} (2\bar{L}_g)^{n-k+1} - 1 - \bar{L}_g}{2\bar{L}_g - 1}, \quad (3.10)$$

for all $k = 0, \dots, n$.

Step 2. From (3.8)-(3.9), we may write :

$$\begin{aligned} \left\| v_k(Z_k) - \hat{v}_k(\hat{Z}_k) \right\|_1 &\leq \left\| v_k(Z_k) - \mathbb{E}[v_k(Z_k) | \hat{Z}_k] \right\|_1 \\ &\quad + \left\| \mathbb{E} \left[\left(\varphi_k(Z_k) - \varphi_k(\hat{Z}_k) \right) R_{k+1} v_{k+1}(Z_k) \middle| \hat{Z}_k \right] \right\|_1 \\ &\quad + \left\| \mathbb{E} \left[\varphi(\hat{Z}_k) \left(R_{k+1} v_{k+1}(Z_k) - \hat{r}_{k+1} \hat{v}_{k+1}(\hat{Z}_k) \right) \middle| \hat{Z}_k \right] \right\|_1 \\ &= I_1 + I_2 + I_3. \end{aligned} \quad (3.11)$$

By the law of iterated conditional expectation, we have :

$$\begin{aligned} I_1 &\leq \left\| v_k(Z_k) - v_k(\hat{Z}_k) \right\|_1 + \left\| \mathbb{E}[v_k(\hat{Z}_k)|\hat{Z}_k] - \mathbb{E}[v_k(Z_k)|\hat{Z}_k] \right\|_1 \\ &\leq 2 \left\| v_k(Z_k) - v_k(\hat{Z}_k) \right\|_1 \leq 2[v_k]_{lip} \left\| Z_k - \hat{Z}_k \right\|_1. \end{aligned}$$

Since conditional expectation (here with respect to \hat{Z}_k) is a L^1 -contraction, and v_{k+1} is bounded by 1, we have :

$$I_2 \leq \left\| \varphi_k(Z_k) - \varphi_k(\hat{Z}_k) \right\|_1 \leq \left\| Z_k - \hat{Z}_k \right\|_1,$$

recalling that φ_k is in $BL_1(K_m \times \mathbb{R}^d)$. Since \hat{Z}_k is $\sigma(Z_k)$ -measurable, and recalling also that φ_k is bounded by 1, we have :

$$I_3 \leq \left\| v_{k+1}(Z_{k+1}) - \hat{v}_{k+1}(\hat{Z}_{k+1}) \right\|_1.$$

Plugging these estimates of I_1 , I_2 and I_3 into (3.11), we get

$$\left\| v_k(Z_k) - \hat{v}_k(\hat{Z}_k) \right\|_1 \leq (1 + 2[v_k]_{lip}) \left\| Z_k - \hat{Z}_k \right\|_1 + \left\| v_{k+1}(Z_{k+1}) - \hat{v}_{k+1}(\hat{Z}_{k+1}) \right\|_1.$$

Since $\left\| v_n(Z_n) - \hat{v}_n(\hat{Z}_n) \right\|_1 \leq \left\| Z_n - \hat{Z}_n \right\|_1$, a direct backward induction yields :

$$\left\| v_k(Z_k) - \hat{v}_k(\hat{Z}_k) \right\|_1 \leq \sum_{j=k}^n (1 + 2[v_j]_{lip}) \left\| Z_j - \hat{Z}_j \right\|_1.$$

The required result is proved by taking $k = 0$, substituting the estimate (3.10) and using Cauchy-Schwarz inequality : $\left\| Z_j - \hat{Z}_j \right\|_1 \leq \sqrt{m+d} \left\| Z_j - \hat{Z}_j \right\|_2$. \square

3.3 Optimal quantization and rate of convergence

The estimation error (3.3) in Proposition 3.1 shows that to obtain the best approximation of the sequence of probability transitions (R_k) of the Markov chain $(Z_k)_k = (\Pi_k, Y_k)_k$ by this quantization approach, one has to minimize at each time $k \geq 1$ the L^2 quantization error $\|Z_k - \hat{Z}_k\|_2$. By identifying a grid $\Gamma_k = \{z^1, \dots, z^{N_k}\}$ of size $|\Gamma_k| = N_k$ points in $K_m \times \mathbb{R}^d$, with the N_k -tuple $(z^1, \dots, z^{N_k}) \in (K_m \times \mathbb{R}^d)^{N_k}$, the objective is then to minimize the symmetric function :

$$\begin{aligned} D_{N_k}^{Z_k}(z^1, \dots, z^{N_k}) &= \left\| Z_k - \text{Proj}_{\Gamma_k}(Z_k) \right\|_2^2 \\ &= \mathbb{E} \left[\min_{1 \leq i \leq N_k} |Z_k - z^i|_2^2 \right], \forall \Gamma_k = (z^1, \dots, z^{N_k}) \in (K_m \times \mathbb{R}^d)^{N_k} \end{aligned} \quad (3.12)$$

which is the square of the L^2 quantization error and is usually called distorsion. The optimal quantization consists, for each $k \geq 1$ and given a number of points N_k , to find a grid $\hat{\Gamma}_k$ of size N_k that reaches the minimum of the distorsion function $D_{N_k}^{Z_k}$. This question has been tackled for a long time as part of quantization for information theory and signal processing, and more recently in probability for both numerical and theoretical purpose (see [7] or [11]).

We recall these results and apply them in our context in the next paragraph. Now, we focus on the behaviour of the minimum of this function, i.e. the minimal L^2 quantization error, when the number of points N_k goes to infinity. For this, we first recall the so-called Zador theorem, see e.g. [7].

Theorem 3.1 *Let X be a \mathbb{R}^l -valued random variable with distribution \mathbb{P}_X , s.t. $\mathbb{E}|X|_2^{2+\delta} < \infty$ for some $\delta > 0$. Then*

$$\lim_N \left(N^{\frac{2}{l}} \min_{|\Gamma| \leq N} \|X - \text{Proj}_{\Gamma}(X)\|_2^2 \right) = J_{2,l} \left(\int_{\mathbb{R}^l} |f|^{\frac{l}{l+2}}(x) dx \right)^{\frac{l+2}{l}} \quad (3.13)$$

where $\mathbb{P}_X(dx) = f(x) \lambda_l(dx) + \nu(dx)$ is the Lebesgue decomposition of \mathbb{P}_X with respect to the Lebesgue measure λ_l on \mathbb{R}^l . The constant $J_{2,l}$ corresponds to the case where X is the uniform distribution on $[0, 1]^l$.

Remark 3.2 In dimension $l = 1$ and 2 , the values of $J_{2,l}$ are known : $J_{2,1} = 1/6$ and $J_{2,2} = \frac{5}{18\sqrt{3}}$. In higher dimension, the true value of $J_{2,l}$ is unknown but we have an equivalent $J_{2,l} \sim \frac{l}{2\pi e}$ as l goes to infinity.

Here, we cannot apply directly this theorem to Z_k since the distribution \mathbb{P}_{Z_k} of (Π_k, Y_k) is not known in general, and in particular its decomposition with respect to the Lebesgue measure. However, one can prove the following error bound for the minimal distortion error of Z_k .

Proposition 3.2 *For $k \geq 1$, assume that there exists some $\varepsilon > 0$ s.t. :*

$$\int |y_k|_2^{2+\varepsilon} \prod_{l=1}^k g_l(x_{l-1}, y_{l-1}, x_l, y_l) dy_l < \infty, \quad \forall x_0, \dots, x_k \in E. \quad (3.14)$$

Then, we have :

$$\limsup_{N_k \rightarrow \infty} N_k^{\frac{2}{m-1+d}} \min_{|\Gamma_k| \leq N_k} \|Z_k - \text{Proj}_{\Gamma_k}(Z_k)\|_2^2 \leq C(m, d, f_k),$$

where

$$C(m, d, f_k) = \frac{m(d+m-1)}{(md)^{\frac{d}{d+m-1}}} (J_{2,d})^{\frac{d}{d+m-1}} \left(\int_{\mathbb{R}^d} |f_k|^{\frac{d}{d+2}}(y) dy \right)^{\frac{d+2}{d+m-1}},$$

and f_k is the marginal density of Y_k given by :

$$f_k(y) = \int g_k(x_{k-1}, y_{k-1}, x_k, y) P_k(x_{k-1}, dx_k) \prod_{l=1}^{k-1} g_l(x_{l-1}, y_{l-1}, x_l, y_l) P_l(x_{l-1}, dx_l) dy_l \mu(dx_0).$$

Remark 3.3 In the case where the law of Y_k conditional on (X_{k-1}, Y_{k-1}, X_k) does not depend on Y_{k-1} , i.e. the function g_k does not depend on Y_{k-1} , the expression f_k of the marginal density of Y_k simplifies into :

$$f_k(y) = \mathbb{E}[g_k(X_{k-1}, X_k, y)], \quad y \in \mathbb{R}^d, \quad (3.15)$$

and condition (3.14) is written as :

$$\int |y_k|_2^{2+\delta} g_k(x_{k-1}, x_k, y_k) dy_k < \infty, \quad \forall x_{k-1}, x_k \in E,$$

for some $\delta > 0$.

Proof of Proposition 3.2. For any grids

$$\begin{aligned} \Gamma^\Pi &= \{\pi(1), \dots, \pi(M)\} \quad \text{of size } |\Gamma^\Pi| = M \text{ points in } K_m, \\ \Gamma^Y &= \{y^1, \dots, y^L\} \quad \text{of size } |\Gamma^Y| = L \text{ points in } \mathbb{R}^d, \end{aligned}$$

we denote

$$\Gamma^\Pi \otimes \Gamma^Y = \{(\pi(i), y^j) : 1 \leq i \leq M, 1 \leq j \leq L\} \quad \text{of size } ML \text{ points in } K_m \times \mathbb{R}^d.$$

We then have by definition of the norm $\|\cdot\|_2$ and of the projection :

$$\|Z_k - \text{Proj}_{\Gamma^\Pi \otimes \Gamma^Y}(Z_k)\|_2^2 = \|\Pi_k - \text{Proj}_{\Gamma^\Pi}(\Pi_k)\|_2^2 + \|Y_k - \text{Proj}_{\Gamma^Y}(Y_k)\|_2^2,$$

and so

$$\begin{aligned} & \min_{|\Gamma_k| \leq N_k} \|Z_k - \text{Proj}_{\Gamma_k}(Z_k)\|_2^2 \\ & \leq \min_{M, L : ML \leq N_k} \left(\min_{|\Gamma^\Pi| \leq M} \|\Pi_k - \text{Proj}_{\Gamma^\Pi}(\Pi_k)\|_2^2 + \min_{|\Gamma^Y| \leq L} \|Y_k - \text{Proj}_{\Gamma^Y}(Y_k)\|_2^2 \right) \end{aligned} \quad (3.16)$$

For M_0 in $\mathbb{N} \setminus \{0\}$, consider the grid Γ^Π of size $M = M_0^{m-1}$ points in K_m :

$$\Gamma^\Pi = \left\{ \left(\frac{i_1}{M_0}, \dots, \frac{i_{m-1}}{M_0}, 1 - \sum_{l=1}^{m-1} \frac{i_l}{M_0} \right) : i_1, \dots, i_{m-1} = 1, \dots, M_0, \sum_{l=1}^{m-1} i_l \leq M_0 \right\},$$

Denoting for all $a \in \mathbb{R}$, $[a]$ the smallest integer smaller than a , we have for all $\pi = (\pi^i)_{1 \leq i \leq m} \in K_m$:

$$\begin{aligned} |\pi - \text{Proj}_{\Gamma^\Pi}(\pi)|_2^2 &= \min_{\substack{i_1, \dots, i_{m-1} = 1, \dots, M_0 \\ i_1 + \dots + i_{m-1} \leq M_0}} \sum_{l=1}^{m-1} \left| \pi^l - \frac{i_l}{M_0} \right|^2 + \left| \sum_{l=1}^{m-1} \left(\pi^l - \frac{i_l}{M_0} \right) \right|^2 \\ &\leq m \min_{\substack{i_1, \dots, i_{m-1} = 1, \dots, M_0 \\ i_1 + \dots + i_{m-1} \leq M_0}} \sum_{l=1}^{m-1} \left| \pi^l - \frac{i_l}{M_0} \right|^2 \\ &\leq m \sum_{l=1}^{m-1} \left| \pi^l - \frac{[\pi^l M_0]}{M_0} \right|^2 \\ &\leq \frac{m(m-1)}{M_0^2}. \end{aligned}$$

This shows that :

$$\min_{|\Gamma^\Pi| \leq M} \|\Pi_k - \text{Proj}_{\Gamma^\Pi}(\Pi_k)\|_2^2 \leq \frac{m(m-1)}{M^{\frac{2}{m-1}}}. \quad (3.17)$$

On the other hand, notice that condition (3.14) ensures that $\mathbb{E}|Y_k|_2^{2+\varepsilon} < \infty$. Thus, from Theorem 3.1 applied to Y_k , we have for all $\delta > 0$ and L large enough,

$$\min_{|\Gamma^Y| \leq L} \|Y_k - \text{Proj}_{\Gamma^Y}(Y_k)\|_2^2 \leq \frac{\left(J_{2,d} \|f_k\|_{\frac{d}{d+2}} + \delta\right)}{L^{\frac{2}{d}}}, \quad (3.18)$$

where we set :

$$\|f_k\|_{\frac{d}{d+2}} = \left(\int_{\mathbb{R}^d} |f_k|^{\frac{d}{d+2}}(y) dy \right)^{\frac{d+2}{d}}.$$

Substituting (3.17) and (3.18) into (3.16) yields :

$$\begin{aligned} & \min_{|\Gamma_k| \leq N_k} \|Z_k - \text{Proj}_{\Gamma_k}(Z_k)\|_2^2 \\ & \leq \min_{M,L : ML \leq N_k} \left[\frac{m(m-1)}{M^{\frac{2}{m-1}}} + \frac{\left(J_{2,d} \|f_k\|_{\frac{d}{d+2}} + \delta\right)}{L^{\frac{2}{d}}} \right]. \end{aligned}$$

We conclude with the elementary result that for all $a, b > 0$:

$$\min_{M,L : ML \leq N} \left[\frac{a}{M^{\frac{2}{l}}} + \frac{b}{L^{\frac{2}{d}}} \right] = (d+l) \left(\frac{a}{l}\right)^{\frac{l}{d+l}} \left(\frac{b}{d}\right)^{\frac{d}{d+l}} \frac{1}{N^{\frac{2}{d+l}}}.$$

□

3.4 Practical implementation of the optimal approximating filter process

We now come back to the numerical implementation of an algorithm that computes for each k :

- an optimal grid $\hat{\Gamma}_k$ which minimizes the distortion :

$$D_{N_k}^{Z_k}(\Gamma_k) = \|Z_k - \text{Proj}_{\Gamma_k}(Z_k)\|_2^2$$

as well as an estimation of this error,

- the weights of the Voronoi tessellations :

$$\begin{aligned} \hat{p}_k^i &= \mathbb{P} \left[Z_k \in C_i(\hat{\Gamma}_k) \right], \quad i = 1, \dots, N_k, \\ \hat{\beta}_{k+1}^{ij} &= \mathbb{P} \left[Z_{k+1} \in C_j(\hat{\Gamma}_{k+1}), Z_k \in C_i(\hat{\Gamma}_k) \right], \quad i = 1, \dots, N_k, j = 1, \dots, N_{k+1}, \end{aligned}$$

and so the probability transition matrix $\hat{r}_{k+1}^{ij} = \hat{\beta}_{k+1}^{ij} / \hat{p}_k^i$.

This program is based on the following key property of the distortion : The function $D_{N_k}^{Z_k}$ is continuously differentiable at any N_k -tuple $\Gamma_k = (z^1, \dots, z^{N_k}) \in (K_m \times \mathbb{R}^d)^{N_k}$ having pairwise distinct components and its gradient is obtained by formal differentiation inside the expectation operator in (3.12) (see [11]) :

$$\nabla D_{N_k}^{Z_k}(\Gamma_k) = 2 \mathbb{E}[H(\Gamma_k, Z_k)], \quad (3.19)$$

where the $(K_m \times \mathbb{R}^d)^{N_k}$ -vector valued function H is given by :

$$H(\Gamma_k, z) = \left((z^i - z) 1_{z \in C_i(\Gamma_k)} \right)_{1 \leq i \leq N_k}, \quad \Gamma_k = (z^1, \dots, z^{N_k}) \in (K_m \times \mathbb{R}^d)^{N_k}, \quad z \in K_m \times \mathbb{R}^d.$$

This above integral representation for $\nabla D_{N_k}^{Z_k}$ suggests to implement a stochastic gradient descent, whenever one is able to simulate easily independent copies of Z_k . We will come back below on the simulation of Z_k . The stochastic gradient procedure is recursively defined by

$$\Gamma_k^{s+1} = \Gamma_k^s - \delta_{s+1} H(\Gamma_k^s, \xi_k^{s+1}) \quad (3.20)$$

where the initial grid Γ_k^0 has N_k pairwise distinct components, $(\xi_k^s)_{s \geq 1}$ is an i.i.d. sequence of \mathbb{P}_{Z_k} -distributed random vectors, and $(\delta_s)_{s \geq 1}$ is a sequence of step parameters satisfying the usual conditions :

$$\sum_s \delta_s = \infty \quad \text{and} \quad \sum_s \delta_s^2 < \infty.$$

In an abstract framework (see *e.g.* [6] or [9]), under some appropriate assumptions, a stochastic gradient descent associated to the integral representation of a so-called potential function ($D_{N_k}^{Z_k}$ in our problem) converges a.s., when s goes to infinity, toward a local minimum $\hat{\Gamma}_k$ of this potential function :

$$\nabla D_{N_k}^{Z_k}(\hat{\Gamma}_k) = 0.$$

Although these assumptions are not fulfilled by $D_{N_k}^{Z_k}$, the encountered theoretical problems can be partially overcome (see [11]) Practical implementation does provide satisfactory results (a commonly encountered situation with gradient descents). Moreover, computation of the weights of the tessellations and of the distortion can be implemented as by-product of the procedure. We now describe this algorithm, known as the *Competitive Learning Vector Quantization* algorithm.

Simulation of the Markov chain $(Z_k)_k$:

We notice that from (2.4), we are able to simulate the probability transition R_k of the $(\mathbb{P}, \mathcal{F}_k^Y)$ Markov chain $(Z_k)_k = (\Pi_k, Y_k)_k$. For $k = 0$, recall that Z_0 is a known deterministic vector equal to $z_0 = (\mu, y_0)$. For $k \geq 1$, starting from (Π_{k-1}, Y_{k-1}) ,

- we simulate X_{k-1} with probability law Π_{k-1} , and then X_k according to the probability transition P_k .
- we simulate Y_k according to the probability transition $q_k(X_{k-1}, Y_{k-1}, X_k, dy')$.

- we compute Π_k by the formula (3.1) :

$$\Pi_k = \frac{GP(Y_{k-1}, Y_k)^\top \Pi_{k-1}}{|GP(Y_{k-1}, Y_k)^\top \Pi_{k-1}|_1}.$$

Subsequently, we stock S independent copies of the Markov chain (Z_0, \dots, Z_n) , that we denote $\xi^s = (\xi_0^s, \dots, \xi_n^s)$, $s = 1, \dots, S$. The algorithm reads as follows :

Initialisation phase :

- Initialize the n grids $\Gamma_k^0 = (z_k^{0,1}, \dots, z_k^{0,N_k}) \in (K_m \times \mathbb{R}^d)^{N_k}$ for $k = 0, \dots, n$, with $\Gamma_0^0 = z_0$ reduced to $N_0 = 1$ point for $k = 0$.
- Initialize the weights vectors : $p_k^{0,i} = 1/N_k$, $\beta_{k+1}^{0,ij} = 0$, $i = 1, \dots, N_k$, $j = 1, \dots, N_{k+1}$, and the distortion $D_{N_k}^0 = 0$, for $k = 0, \dots, n$.

Updating $s \rightarrow s+1$: At step s , the n grids $\Gamma_k^s = (z_k^{s,1}, \dots, z_k^{s,N_k})$, the weights vectors $p_k^{s,i}$, $\beta_{k+1}^{s,ij}$, $i = 1, \dots, N_k$, $j = 1, \dots, N_{k+1}$, the distortion $D_{N_k}^s$ have been obtained and we use the sample ξ^{s+1} of (Z_0, \dots, Z_n) to update them as follows : for all $k = 0, \dots, n$,

- *Competitive phase* : select $i_k(s+1) \in \{1, \dots, N_k\}$ such that

$$\xi^{s+1} \in C_{i_k(s+1)}(\Gamma_k^s), \quad \text{i.e.} \quad i_k(s+1) \in \operatorname{argmin}_{1 \leq i \leq N_k} |z_k^{s,i} - \xi^{s+1}|_2.$$

- *Learning phase* :

★ Updating of the grid :

$$z_k^{s+1,i} = z_k^{s,i} - \delta_{s+1} \mathbf{1}_{i=i_k(s+1)} \left(z_k^{s,i} - \xi^{s+1} \right), \quad i = 1, \dots, N_k$$

★ Updating of the weights vectors and of the probability transition

$$\begin{aligned} p_k^{s+1,i} &= p_k^{s,i} - \delta_{s+1} \left(p_k^{s,i} - \mathbf{1}_{i=i_k(s+1)} \right), \\ \beta_{k+1}^{s+1,ij} &= \beta_{k+1}^{s,ij} - \delta_{s+1} \left(\beta_{k+1}^{s,ij} - \mathbf{1}_{i=i_k(s+1), j=i_{k+1}(s+1)} \right), \\ r_{k+1}^{s+1,ij} &= \frac{\beta_{k+1}^{s+1,ij}}{p_k^{s+1,i}}, \end{aligned}$$

for all $i = 1, \dots, N_k$, $j = 1, \dots, N_{k+1}$.

★ Updating of the distortion

$$D_{N_k}^{s+1} = D_{N_k}^s - \delta_{s+1} \left(D_{N_k}^s - \left| z_k^{s, i_k(s+1)} - \xi^{s+1} \right|_2^2 \right),$$

It is shown in [11] that on the event $\{\Gamma_k^s \rightarrow \hat{\Gamma}_k\}$, set of trajectories of $(\Gamma_k^s)_s$ that converge to $\hat{\Gamma}_k$ local minimum of the distortion, we have :

$$\begin{aligned} p_k^{s,i} &\longrightarrow \hat{p}_k^i = \mathbb{P} \left[Z_k \in C_i(\hat{\Gamma}_k) \right], \quad a.s. \\ \beta_{k+1}^{s,ij} &\longrightarrow \hat{\beta}_{k+1}^{ij} = \mathbb{P} \left[Z_{k+1} \in C_j(\hat{\Gamma}_{k+1}), Z_k \in C_i(\hat{\Gamma}_k) \right], \quad a.s. \\ D_{N_k}^s &\longrightarrow D_{N_k}^{Z_k}(\hat{\Gamma}_k), \quad a.s. \end{aligned}$$

for all $k = 0, \dots, n$, $i = 1, \dots, N_k$, $j = 1, \dots, N_{k+1}$, as s goes to infinity.

4 Application : optimal stopping under partial observation

We consider the framework of Section 3. We denote by \mathcal{T}_n^Y the set of stopping times adapted with respect to the observation filtration (\mathcal{F}_k^Y) and valued in $\{0, \dots, n\}$. Given two measurable functions f and h on $E \times \mathbb{R}^d$, we consider the following optimal stopping problem under partial observation :

$$u_0 = \sup_{\tau \in \mathcal{T}_n^Y} \mathbb{E} \left[\sum_{k=0}^{\tau} f(X_k, Y_k) + h(X_\tau, Y_\tau) \right]. \quad (4.1)$$

We shall transform this problem into an optimal stopping problem under complete observation. We denote

$$J(\tau) = \mathbb{E} \left[\sum_{k=0}^{\tau} f(X_k, Y_k) + h(X_\tau, Y_\tau) \right], \quad \tau \in \mathcal{T}_n^Y,$$

the expected gain function associated to (4.1). We shall also introduce the functions :

$$\begin{aligned} \hat{f}(\pi, y) &= \sum_{i=1}^m f(x^i, y) \pi^i, \quad \forall \pi = (\pi^i)_i \in K_m, \quad \forall y \in \mathbb{R}^d, \\ \hat{h}(\pi, y) &= \sum_{i=1}^m h(x^i, y) \pi^i, \quad \forall \pi = (\pi^i)_i \in K_m, \quad \forall y \in \mathbb{R}^d. \end{aligned}$$

Then, by using the law of iterated conditional expectations and the definition of the filter (Π_k) , we have for all $\tau \in \mathcal{T}_n^Y$:

$$\begin{aligned} J(\tau) &= \mathbb{E} \left[\sum_{j=0}^n 1_{\tau=j} \left(\sum_{k=0}^j f(X_k, Y_k) + h(X_j, Y_j) \right) \right] \\ &= \mathbb{E} \left[\sum_{j=0}^n 1_{\tau=j} \left(\sum_{k=0}^j \mathbb{E} [f(X_k, Y_k) | \mathcal{F}_k^Y] + \mathbb{E} [h(X_j, Y_j) | \mathcal{F}_j^Y] \right) \right] \\ &= \mathbb{E} \left[\sum_{j=0}^n 1_{\tau=j} \left(\sum_{k=0}^j \hat{f}(\Pi_k, Y_k) + \hat{h}(\Pi_j, Y_j) \right) \right] \\ &= \mathbb{E} \left[\sum_{k=0}^{\tau} \hat{f}(Z_k) + \hat{h}(Z_\tau) \right]. \end{aligned}$$

Therefore, problem (4.1) may be rewritten as :

$$u_0 = \sup_{\tau \in \mathcal{T}_n^Y} \mathbb{E} \left[\sum_{k=0}^{\tau} \hat{f}(Z_k) + \hat{h}(Z_\tau) \right]. \quad (4.2)$$

Since the process $(Z_k)_k = (\Pi_k, Y_k)_k$ is a $(\mathbb{P}, (\mathcal{F}_k^Y))$ Markov chain, problem (4.2) is then an optimal stopping problem under complete observation.

By the dynamic programming principle, we have $u_0 = v_0(z_0)$, where the sequence of measurable functions $v_k : K_m \times \mathbb{R}^d \mapsto \mathbb{R}$, $k = 0, \dots, n$, is given in recursive form by the backward formula :

$$v_n(z) = \hat{h}(z), \quad \forall z \in K_m \times \mathbb{R}^d \quad (4.3)$$

$$v_k(z) = \max \left\{ \hat{h}(z); \mathbb{E} \left[\hat{f}(Z_{k+1}) + v_{k+1}(Z_{k+1}) \middle| Z_k = z \right] \right\}, \\ \forall z \in K_m \times \mathbb{R}^d, \quad k = 0, \dots, n-1. \quad (4.4)$$

Applying the quantization approach described in Section 3 and setting :

$$\hat{Z}_k = \text{Proj}_{\Gamma_k}(Z_k), \quad \Gamma_k = (z_k^1, \dots, z_k^{N_k}) \in (K_m \times \mathbb{R}^d)^{N_k}, \quad k = 0, \dots, n,$$

we then approximate, following [1], the sequence of functions (v_k) by the sequence of functions $\hat{v}_k : \Gamma_k \rightarrow \mathbb{R}$, $k = 0, \dots, n$ defined by :

$$\hat{v}_n(z) = \hat{h}(z), \quad \forall z \in \Gamma_n \\ \hat{v}_k(z) = \max \left\{ \hat{h}(z); \mathbb{E} \left[\hat{f}(\hat{Z}_{k+1}) + \hat{v}_{k+1}(\hat{Z}_{k+1}) \middle| \hat{Z}_k = z \right] \right\}, \\ \forall z \in \Gamma_k, \quad k = 0, \dots, n-1.$$

From an algorithmic viewpoint, this reads as follows :

$$\hat{v}_n(z_n^i) = \hat{h}(z_n^i), \quad i = 1, \dots, N_n \\ \hat{v}_k(z_k^i) = \max \left\{ \hat{h}(z_k^i); \sum_{j=1}^{N_{k+1}} \hat{r}_{k+1}^{ij} \left(\hat{f}(z_{k+1}^j) + \hat{v}_{k+1}(z_{k+1}^j) \right) \right\}, \\ i = 1, \dots, N_k, \quad k = 0, \dots, n-1.$$

The optimal grids Γ_k and the associated probability transition matrix (\hat{r}_{k+1}^{ij}) are estimated following the procedure described in Section 3.

The induced approximation error is provided in the following theorem.

Theorem 4.1 *Assume that f and h are bounded and Lipschitz, uniformly in $x \in E$, i.e.*

$$\|f\|_{sup} := \sup_{x \in E, y \in \mathbb{R}^d} |f(x, y)| < \infty \quad \text{and} \quad [f]_{lip} := \sup_{x \in E, y, \hat{y} \in \mathbb{R}^d, y \neq \hat{y}} \frac{|f(x, y) - f(x, \hat{y})|}{\|y - \hat{y}\|_1} < \infty.$$

Then under **(H1)**, we have for all $k = 0, \dots, n$:

$$\left\| v_k(Z_k) - \hat{v}_k(\hat{Z}_k) \right\|_1 \\ \leq \sqrt{m+d}(\bar{f} + \bar{h}) \sum_{j=k}^n \left[(7 + 2(n-j))\bar{L}_g + (n-j+2) \frac{(2\bar{L}_g)^{n-j+1}}{2\bar{L}_g - 1} \right] \|Z_j - \hat{Z}_j\|_2, \quad (4.5)$$

where $\bar{f} = \max(\|f\|_{sup}, [f]_{lip})$, $\bar{h} = \max(\|h\|_{sup}, [h]_{lip})$.

Remark 4.1 In view of Proposition 3.2, the estimation (4.5) provides a rate of convergence for the approximation of $v_0(Z_0)$ of order

$$\frac{n(2\bar{L}_g)^n}{N^{\frac{1}{m-1+d}}},$$

when $N_k = N$ is the number of points at each grid Γ_k used for the optimal quantization of Z_k , $k = 0, \dots, n$. The term $(2\bar{L}_g)^n$ is important when n is large, but this is consistent with the rate of convergence obtained in approximation of nonlinear filtering by quantization, see [12] or by Monte-Carlo particle methods, see [4].

Lemma 4.1 *Under the assumptions of Theorem 4.1, the functions v_k , $k = 0, \dots, n$, defined in (4.4), are bounded and Lipschitz with :*

$$\|v_k\|_{sup} \leq (n - k + 1)(\bar{f} + \bar{h}), \quad (4.6)$$

$$[v_k]_{lip} \leq \frac{n - k + 3}{2} \frac{(\bar{f} + \bar{h})}{2\bar{L}_g - 1} (2\bar{L}_g)^{n-k+1}. \quad (4.7)$$

Proof. From (4.4), we have $\|v_k\|_{sup} \leq \|\hat{h}\|_{sup} + \|\hat{f}\|_{sup} + \|v_{k+1}\|_{sup}$. Since $\|v_n\|_{sup} = \|\hat{h}\|_{sup}$, $\|\hat{f}\|_{sup} \leq \|f\|_{sup}$, $\|\hat{h}\|_{sup} \leq \|h\|_{sup}$, we obtain immediately the inequality (4.6) by induction.

On the other hand, relation (4.4) and Lemma 3.1 also show

$$\begin{aligned} [v_k]_{lip} &\leq [\hat{h}]_{lip} + [R_{k+1}(\hat{f} + v_{k+1})]_{lip} \\ &\leq [\hat{h}]_{lip} + \bar{L}_g \left(\|\hat{f}\|_{sup} + 2[\hat{f}]_{lip} + \|v_{k+1}\|_{sup} + 2[v_{k+1}]_{lip} \right). \end{aligned} \quad (4.8)$$

By definition of \hat{f} , we clearly have for all $z = (\pi, y)$ and $z' = (\pi', y')$ in $K^m \times \mathbb{R}^d$,

$$\begin{aligned} |\hat{f}(z) - \hat{f}(z')| &\leq \|f\|_{sup} |\pi - \pi'|_1 + [f]_{lip} |y - y'|_1 \\ &\leq \bar{f} |z - z'|_1, \end{aligned}$$

A similar inequality holds for $[h]_{lip}$ i.e. $[\hat{h}]_{lip} \leq \bar{h}$. Plugging into (4.8) and using (4.6) yields

$$\begin{aligned} [v_k]_{lip} &\leq \bar{h} + \bar{L}_g (3\bar{f} + (n - k)(\bar{f} + \bar{h})) \\ &\quad + 2\bar{L}_g [v_{k+1}]_{lip}. \end{aligned} \quad (4.9)$$

Since $[v_n]_{lip} = [\hat{h}]_{lip}$, a straightforward induction gives (4.7). \square

Proof of Theorem 4.1.

We set $\Phi_k(z) = \mathbb{E}[\hat{f}(Z_{k+1}) + v_{k+1}(Z_{k+1}) | Z_k = z]$ and $\hat{\Phi}_k(z) = \mathbb{E}[\hat{f}(\hat{Z}_{k+1}) + \hat{v}_{k+1}(\hat{Z}_{k+1}) | \hat{Z}_k = z]$. Then, for $k = 0, \dots, n - 1$,

$$\begin{aligned} \|v_k(Z_k) - \hat{v}_k(\hat{Z}_k)\|_1 &\leq \|\hat{h}(Z_k) - \hat{h}(\hat{Z}_k)\|_1 + \|\Phi_k(Z_k) - \hat{\Phi}_k(\hat{Z}_k)\|_1 \\ &\leq [\hat{h}]_{lip} \|Z_k - \hat{Z}_k\|_1 + \|\Phi_k(Z_k) - \Phi_k(\hat{Z}_k)\|_1 \end{aligned} \quad (4.10)$$

$$\begin{aligned} &\quad + \left\| \mathbb{E}[\Phi_k(\hat{Z}_k) | \hat{Z}_k] - \mathbb{E}[\Phi_k(Z_k) | \hat{Z}_k] \right\|_1 + \left\| \mathbb{E}[\Phi_k(Z_k) | \hat{Z}_k] - \hat{\Phi}_k(\hat{Z}_k) \right\|_1 \\ &\leq [\hat{h}]_{lip} \|Z_k - \hat{Z}_k\|_1 + 2 \|\Phi_k(Z_k) - \Phi_k(\hat{Z}_k)\|_1 \\ &\quad + \left\| \mathbb{E}[\Phi_k(Z_k) | \hat{Z}_k] - \hat{\Phi}_k(\hat{Z}_k) \right\|_1, \end{aligned} \quad (4.11)$$

by the law of iterated conditional expectation. Since \hat{Z}_k is $\sigma(Z_k)$ -measurable, we have

$$\begin{aligned} & \left\| \mathbb{E}[\Phi_k(Z_k)|\hat{Z}_k] - \hat{\Phi}_k(\hat{Z}_k) \right\|_1 \\ &= \left\| \mathbb{E}[\hat{f}(Z_{k+1}) + v_{k+1}(Z_{k+1})|\hat{Z}_k] - \mathbb{E}[\hat{f}(\hat{Z}_{k+1}) + \hat{v}_{k+1}(\hat{Z}_{k+1})|\hat{Z}_k] \right\|_1 \\ &\leq [\hat{f}]_{lip} \left\| Z_{k+1} - \hat{Z}_{k+1} \right\|_1 + \left\| v_{k+1}(Z_{k+1}) - \hat{v}_{k+1}(\hat{Z}_{k+1}) \right\|_1. \end{aligned}$$

Plugging into (4.11) yields :

$$\begin{aligned} \left\| v_k(Z_k) - \hat{v}_k(\hat{Z}_k) \right\|_1 &\leq \left([\hat{h}]_{lip} + 2[\Phi_k]_{lip} \right) \left\| Z_k - \hat{Z}_k \right\|_1 + [\hat{f}]_{lip} \left\| Z_{k+1} - \hat{Z}_{k+1} \right\|_1 \\ &\quad + \left\| v_{k+1}(Z_{k+1}) - \hat{v}_{k+1}(\hat{Z}_{k+1}) \right\|_1. \end{aligned}$$

Since $\left\| v_n(Z_n) - \hat{v}_n(\hat{Z}_n) \right\|_1 \leq [\hat{h}]_{lip} \left\| Z_n - \hat{Z}_n \right\|_1$, a direct induction gives :

$$\left\| v_k(Z_k) - \hat{v}_k(\hat{Z}_k) \right\|_1 \leq \sum_{j=k}^n a_j \left\| Z_j - \hat{Z}_j \right\|_1 \quad (4.12)$$

where

$$a_j = \begin{cases} [\hat{h}]_{lip} + 2[\Phi_k]_{lip}, & j = k \\ [\hat{h}]_{lip} + 2[\Phi_j]_{lip} + [\hat{f}]_{lip}, & j = k+1, \dots, n-1 \\ [\hat{h}]_{lip} + [\hat{f}]_{lip}, & j = n \end{cases} \quad (4.13)$$

Now, by Lemmata 3.1 and 4.1, we have

$$\begin{aligned} [\Phi_k]_{lip} &= [R_{k+1}(\hat{f} + v_{k+1})]_{lip} \\ &\leq \bar{L}_g \left(\|\hat{f}\|_{sup} + 2[\hat{f}]_{lip} + \|v_{k+1}\|_{sup} + 2[v_{k+1}]_{lip} \right) \\ &\leq \bar{L}_g [3\bar{f} + (n-k)(\bar{f} + \bar{h})] \\ &\quad + (\bar{f} + \bar{h}) \left(\frac{n-k}{2} + 1 \right) \frac{(2\bar{L}_g)^{n-k+1}}{2\bar{L}_g - 1}. \end{aligned}$$

Substituting into (4.13) and (4.12) provides the required result by using also Cauchy-Schwarz inequality.

5 Numerical illustration : Bermudean options in a partially observed stochastic volatility model

We consider an observable risky asset price (S_k) with dynamics given by :

$$S_{k+1} = S_k \exp \left[\left(r - \frac{1}{2} X_k^2 \right) \delta + X_k \sqrt{\delta} \varepsilon_{k+1} \right], \quad S_0 = s_0 > 0,$$

where (ε_k) is a sequence of Gaussian white noise, and (X_k) is the unobservable volatility process. $\delta > 0$ may represent some discretization time step. Equivalently, we observe the

process $(Y_k) = (\ln S_k)$, and we notice that the conditional law of Y_{k+1} given (X_k, Y_k) has a density given by :

$$g(X_k, Y_k, y') = \frac{1}{\sqrt{2\pi X_k^2 \delta}} \exp \left[-\frac{(y' - Y_k - (r - \frac{1}{2} X_k^2) \delta)^2}{2X_k^2 \delta} \right], \quad y' \in \mathbb{R}.$$

We model here the dynamics of (X, S) under some risk neutral martingale measure \mathbb{P} , r representing in this case the riskless interest rate.

We assume that (X_k) is an homogeneous Markov chain taking three possible values $x^b < x^m < x^h$ in $\mathbb{R}_+ \setminus \{0\}$. Its probability transition matrix is given by :

$$P_k = \begin{pmatrix} 1 - (p_{bm} + p_{bh})\delta & p_{bm}\delta & p_{bh}\delta \\ p_{mb}\delta & 1 - (p_{mb} + p_{mh})\delta & p_{mh}\delta \\ p_{hb}\delta & p_{hm}\delta & 1 - (p_{hb} + p_{hm})\delta \end{pmatrix}. \quad (5.1)$$

In this context of a partially observed stochastic volatility model, we consider a Bermudean put option with payoff :

$$h(y) = (\kappa - e^y)_+, \quad y \in \mathbb{R},$$

and we want to compute its price given by :

$$u_0 = \sup_{\tau \in \mathcal{T}_n^Y} \mathbb{E} \left[e^{-r\tau\delta} h(Y_\tau) \right]. \quad (5.2)$$

We consider a model where the volatility (X_k) is a Markov-chain approximation à la Kushner (see [8]) of a mean-reverting process :

$$dX_t = \lambda(x_0 - X_t)dt + \eta dW_t.$$

Denoting by $\Delta > 0$ the spatial step, this corresponds to a probability transition matrix of the form (5.1) with :

$$x^b = x_0 - \Delta, \quad x^m = x_0, \quad x^h = x_0 + \Delta,$$

and

$$\begin{aligned} p_{bm} &= \lambda + \frac{\eta^2}{2\Delta^2}, & p_{bh} &= 0 \\ p_{mb} &= \frac{\eta^2}{2\Delta^2}, & p_{mh} &= \frac{\eta^2}{2\Delta^2} \\ p_{hb} &= 0, & p_{hm} &= \lambda + \frac{\eta^2}{2\Delta^2}. \end{aligned}$$

In order to ensure that P_k is indeed a probability transition matrix, we have the consistency conditions :

$$1 - \left(\lambda + \frac{\eta^2}{2\Delta^2} \right) \delta \geq 0 \quad \text{and} \quad 1 - \frac{\eta^2}{\Delta^2} \delta \geq 0$$

We perform numerical tests with

- Price and put option parameters : $r = 0.05$, $S_0 = 110$, $\kappa = 100$,
- Volatility parameters : $\lambda = 1$, $\eta = 0,1$, $X_0 = 0.15$,
- Spatial step : $\Delta = 0,05$.
- Quantization : Grids are of same size N fixed for each time period with step $\delta = \frac{1}{n}$.

We first compare in Table 1 the filter expectation at the final date computed with a time step size $\delta = 1/5$ and by using the optimal quantization method with increasing grid size N , and with 10^6 Monte Carlo iterations.

	$E[\Pi_n^1]$	$E[\Pi_n^2]$	$E[\Pi_n^3]$	Relative error (%)
Monte Carlo	0.287608	0.422833	0.289558	
Quant. with $N = 300$	0.301651	0.421725	0.276624	0.898
Quant. with $N = 600$	0.301604	0.421458	0.276938	0.886
Quant. with $N = 900$	0.301598	0.421316	0.277086	0.881
Quant. with $N = 1200$	0.301618	0.42122	0.277162	0.879
Quant. with $N = 1500$	0.301605	0.421205	0.27719	0.878

Table 1: Comparison of quantized filter value to its Monte Carlo estimation

We observe that besides the very low error level, the absolute error (plotted in Figure 1) and the relative error are decreasing as the grid size grows.

Secondly, in order to illustrate the effect of the time step, we compute the American option price under partial observation when the time step δ decreases to zero (i.e. n increases) and compare it with the American option price with complete observation of (X_k, Y_k) . Indeed, in the limit for $\delta \rightarrow 0$ we fully observe the volatility, and so the partial observation price should converge to the complete observation price.

Moreover, when we have more and more observations, the difference between the two prices should decrease and converge to zero. This is shown in figure 2, where we performed option pricing over grids of size $N_{\Pi, Y} = 1500$ in case of partial observation. The total observation price is given by the same pricing algorithm carried out on $N_{X, Y} = 45$ points for the product grid of (X_k, Y_k) . We have seen in Remark 4.1 that for fixed n , the rate of convergence for the approximation of the value function under partial observation is of order $N_{\Pi, Y}^{1/(m-1+d)}$ where $N_{\Pi, Y}$ is the number of points used at each time k for the grid of (Π_k, Y_k) valued in $K^m \times \mathbb{R}^d$. From results of [1], we also know that the rate of convergence for the approximation of the value function under full observation is of order $m \times N_Y$ where $N_{X, Y} = m \times N_Y$ is the number of points at each time k , used for the grid of (X_k, Y_k) valued in $E \times \mathbb{R}^d$. This explains why, in order to have comparable results, and with $m = 3$ and $d = 1$, we have chosen $N_Y \sim N_{\Pi, Y}^{1/3}$.

In addition, it is possible to observe the effect of information enrichment as the time step decreases. In fact, if we consider multiples of n as the time step parameter, we notice that the American option price increases for both total and partial observation models (see tables 2 and 3).

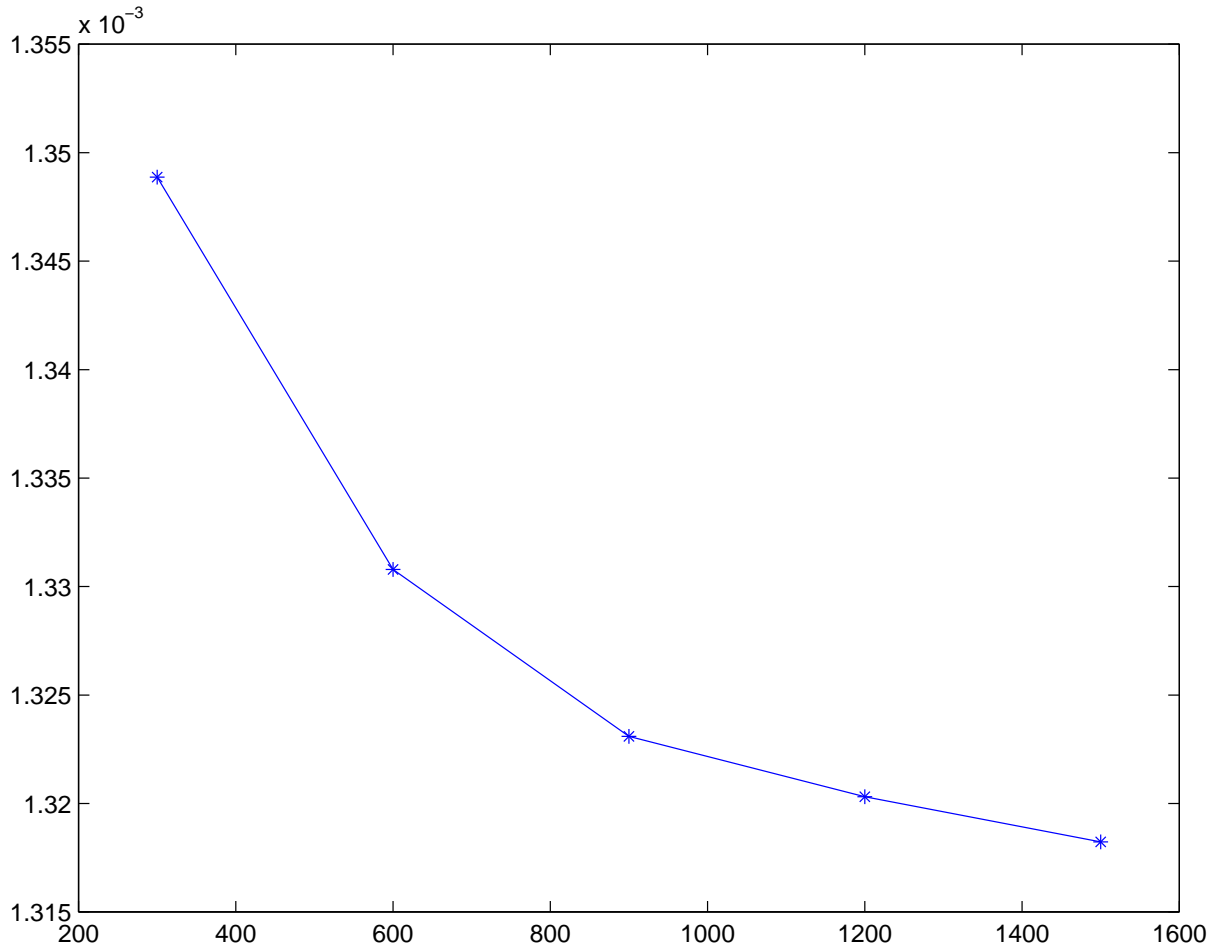


Figure 1: Filter error convergence as N grows

n	4	8	16
Tot. Obs. ($N_{X,Y} = 30$)	1.45863	1.75689	1.77642
Part. Obs. ($N_{II,Y} = 1000$)	0.921729	1.13898	1.47089
Variation	0.53	0.61	0.30

Table 2: American option price for embedded filtrations - First Example

n	5	10	20
Tot. Obs. ($N_{X,Y} = 45$)	1.57506	1.72595	1.91208
Part. Obs. ($N_{II,Y} = 1500$)	0.988531	1.30616	1.59632
Variation	0.58	0.42	0.31

Table 3: American option price for embedded filtrations - Second Example

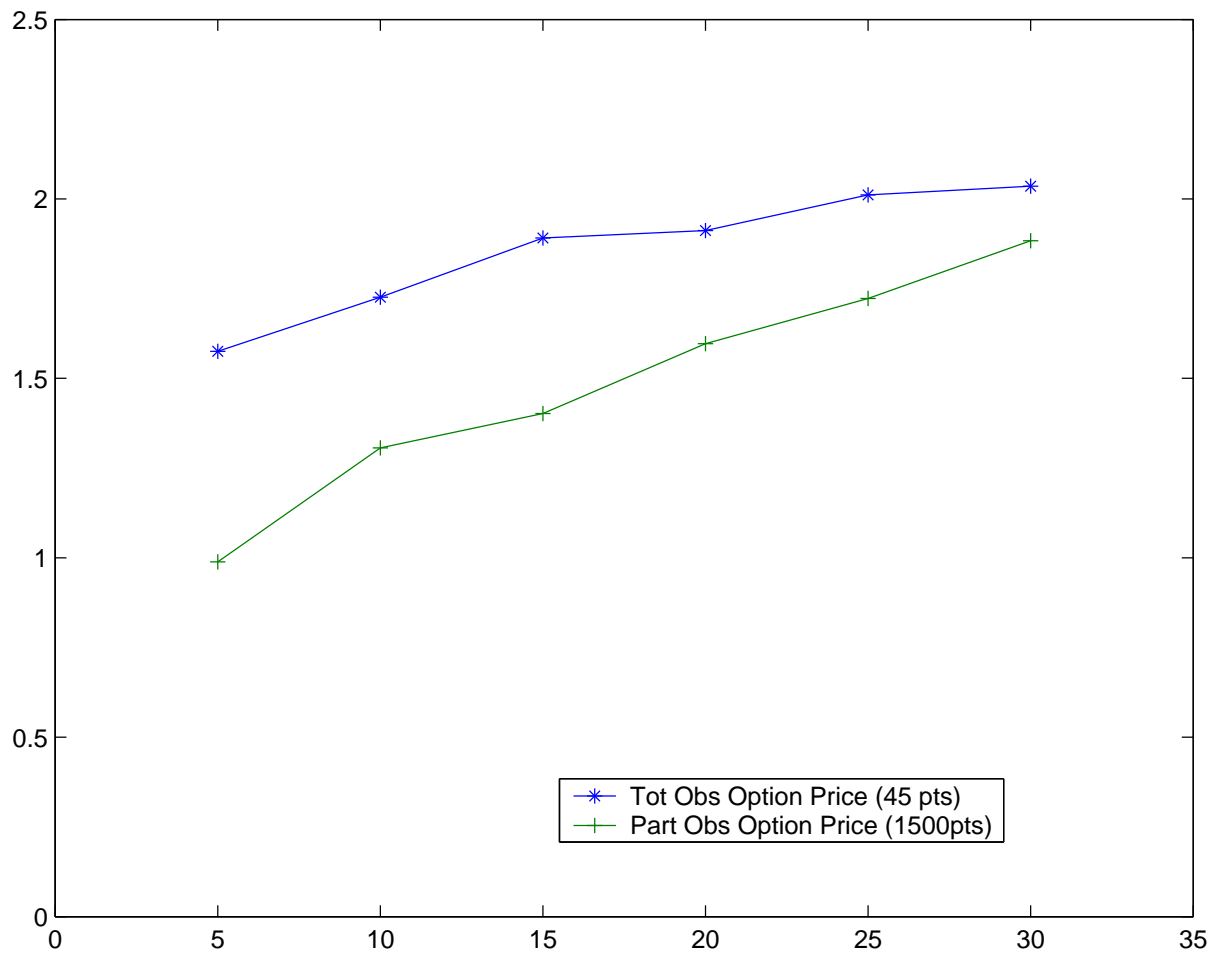


Figure 2: Partial and total observation option prices as $\delta \rightarrow 0$

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